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Descriptive Finding

Changes in seasonality of births in Poland in the years 1900–2009

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Jacek Cypryjański¹

Abstract

BACKGROUND

To date, several countries have documented changes over time in the pattern of seasonal fluctuations of births. Differences in these changes between countries allow us to better understand the social and economic factors determining birth seasonality.

OBJECTIVE

The aim of the present study is to explore birth seasonality in Poland over a period of 110 years.

METHODS

The study was conducted using the birth dates of 48,090,100 Polish citizens born in the years 1900–2009 and registered in the General Electronic System for Registration of the Population. I transformed the data into monthly time series of births, normalized the monthly birth rates for the unequal duration of months and years, and detrended by taking them to the centred 12-month average. Finally, I estimated seasonality patterns for particular decades by means of least squares regression on the harmonic model.

RESULTS

As many as three, relatively stable seasonal birth patterns occurred in Poland in the studied period. In 20th century Europe, the first one (1900–1919), characterised by a minimum birth rate in June, had previously only been observed in Italy. The second pattern with the spring peak and local peak in September remained almost unchanged in Poland for the longest time (1950–1979). Then, within a decade, it transformed into a new, third pattern (1990–2009), with peaks in July and September.

CONTRIBUTION

This article is the first to document in detail the changes in birth seasonality in Poland over a period of 110 years.

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1. Introduction

Research carried out to date clearly shows that seasonal birth fluctuations occur in all studied populations; however, they differ depending on the geographical location and country, and they also change over time (Cowgill 1966a, 1966b; Lam and Miron 1991, 1994; Dorélien 2016). The differences can be the consequence of a number of interrelated factors: biological factors (Smits et al. 1998), various environmental factors such as climate, temperature, and photoperiod (Roenneberg and Aschoff 1990a, 1990b; Lam and Miron 1996; Ruiu and Breschi 2017), and sociocultural and economic factors including seasonality of marriages, birth order, religious calendar, and industrialization (Bobak and Gjonca 2001; Haandrikman 2004; Friger, Shoham-Vardi, and Abu-Saad 2009; Herteliu et al. 2015; Björnsson 2016; Dahlberg and Andersson 2018)

Analysis of changes in seasonal fluctuations over time is difficult due to limited access to sufficiently long time series. Hence, our knowledge of the scale and, most importantly, the dynamics of these changes is fragmented. Nevertheless, we know that in the second half of the 20th century seasonal fluctuations decreased in many countries, leading to a loss of birth seasonality in Spain in the 1990s (Cancho-Candela, Andrés-de Llano, and Ardura-Fernández 2007), and to the appearance of new patterns in, for example, Germany, Scotland, and the Netherlands, (Lerchl, Simoni, and Nieschlag 1993; Russell, Douglas, and Allan 1993; Haandrikman 2004). In Iceland and Sweden significant changes in seasonality took place over 150 and 100 years respectively (Björnsson 2016; Cassel 2002). Comparing these changes in various countries against other indicators can shed more light on the factors determining the seasonality of births.

This paper documents changes in the seasonality of births in Poland in the years 1900–2009. This is the first such study concerning Poland. It shows that over a period of 110 years there were as many as three different patterns of birth seasonality. It also illustrates how historical events interfered with the pattern existing at a given time.

2. Data and methods

The study was conducted on the birth dates of 48,090,100 Polish citizens born in the years 1900–2009 and registered in the General Electronic System for Registration of the Population (PESEL). The PESEL database does not contain all births (Figure 1), but it is the only digital source of individual data for the whole analysed period, including the times when the country was partitioned or at war (Statistics Poland, personal communication).



Figure 1: Number of live births in the years 1900–2009 and its representation in PESEL database

Source: Number of live births obtained from Statistics Poland, Szulc (1936), and Jelonek (1986).

PESEL contains only the birth dates of those who were citizens, lived in the new borders of the Polish state as delimited after World War II, and were alive between 1974 and the early 1990s when PESEL was created. To verify that the PESEL database reflects the actual birth patterns, PESEL data for 1932 was compared with the only extant data that is relatively complete (89.8% of all births): the daily numbers of live births among the wedded in Poland in 1932, excluding the capital city of Warsaw (Zaremba 1936). Both sources generate almost identical monthly birth patterns (the Pearson correlation coefficient is 0.991), showing that the PESEL database is a reliable source for studying seasonality in the first half of the 20th century.

I transformed the data into a monthly series, and then corrected it due to the phenomenon of birth date falsification that occurred in Poland until the beginning of the 1970s, and consisted of shifting the date of birth from December to January. I used an indicator based on the assumption that the January increase in the number of births compared to births in December of the previous year should not exceed the average increase in 1973–1976 (when no phenomena likely to disturb the relationship occurred). Next, the monthly number of births was divided by the number of days a month in order to normalise the unequal length of individual months and the existence of leap years. Finally, the time series was detrended by taking the monthly number of births to the centred 12-month moving average (Figure 2).



Figure 2: Deviations of the corrected and normalised monthly number of births from the 12-month moving average in the years 1900–2009

Note: The probable month of conception is shown in brackets after the month of birth.

The monthly series of deviations from the moving average (DMA) was used to estimate the seasonality patterns for individual decades and for the entire analysed period. For the estimation I used the least squares method and the harmonic model (HM), as well as the seasonal means model (SMM, with dummy variables for each month) for comparison of results (Cornelissen 2014; Cryer and Chan 2008). Harmonic models have the property that the individual harmonics in the model are not correlated, which means that each of them explains a different part of the total variation. Therefore, the percentage to which the individual harmonics explain the variation can be summarised and their sum is equal to the model's coefficient of determination. Moreover, the analysis of individual harmonics helps to identify changes in the seasonality pattern over time.

The HM consisted of six harmonics (sine and cosine functions of equal period) H_i where *i* is not only an index but also shows the number of periods of a given harmonic in a year. After estimating the parameters of the HM, insignificant sine and cosine functions (with a confidence level of 0.01) were removed and the parameters of the adjusted model (AHM) were estimated. All calculations were applied by means of the programming language R, version 3.4.2, and the integrated development environment RStudio, version 1.1.447.

3. Results

Table 1 shows the F-statistics and coefficients of determination R^2 , as well as the percentage to which the individual harmonics of the model explain the total variation.

All the estimated models are statistically significant and differ in the R^2 and in the percentage to which individual harmonics explain the total variation. The number of significant harmonics (in AHM) ranges between 2 and 6, where H_1 and H_2 are present in every model. Moreover, the amplitude and acrophase of the individual harmonics (Figure 3) change over the following decades. The way in which the acrophase of H_1 changes is noteworthy. In the first two decades (1900–1919), H_1 takes the form of a cosine function. In the next three decades its acrophase gradually increases and in the 1950s the H_1 takes the form close to sine, and in 1960–1979 the form of sine function. In the subsequent decades the acrophase of H_1 continues to grow and in the years 2000–2009 H_1 is near the function of –cosine.

Decade	Model	F	R ² (%)	% of the total variation described by the individual harmonics and their groups							
				H1	H ₂	H ₃	H_4	H₅	H ₆	$H_1:H_2$	H ₃ :H ₆
1900–1909	HM	98	91	70.0	16.0	2.6	0.3	0.0	2.0	86.0	4.9
	AHM	214	90	70.0	15.8	2.6			2.0	85.8	4.6
1910–1919	HM	21	68	50.1	15.0	1.4	0.3	0.1	0.8	65.1	2.6
	AHM	98	63	48.3	14.4					62.7	0.0
1920–1929	НМ	58	86	63.4	16.4	2.9	0.4	0.0	2.5	79.7	5.8
	AHM	106	85	63.4	16.4	2.7			2.5	79.7	5.2
1930–1939	НМ	74	88	53.7	20.2	9.4	1.6	0.0	3.4	73.8	14.5
	AHM	105	88	53.7	20.2	9.4	1.6		3.4	73.8	14.5
1940–1949	НМ	46	82	56.3	19.0	4.6	0.9	0.4	1.2	75.2	7.1
	AHM	95	81	56.3	19.0	4.2			1.2	75.2	5.4
1950–1959	НМ	93	90	70.7	14.1	2.6	0.8	0.2	2.1	84.8	5.7
	AHM	142	90	70.7	14.1	2.2	0.8		2.1	84.8	5.1
1960–1969	НМ	249	96	86.0	7.2	0.7	0.9	0.4	1.0	93.1	3.1
	AHM	397	95	85.7	7.0	0.6	0.8	0.3	1.0	92.7	2.7
1970–1979	HM	107	92	82.4	5.2	0.7	0.5	0.4	2.4	87.6	4.0
	AHM	325	89	82.3	4.7				2.4	87.0	2.4
1980–1989	НМ	93	90	73.3	11.9	0.6	0.9	0.6	3.2	85.2	5.3
	AHM	154	89	73.3	11.9		0.7		3.2	85.2	3.9
1990–1999	НМ	88	90	51.5	26.6	1.5	2.8	2.0	5.6	78.0	11.9
	AHM	103	89	51.5	26.6	1.0	2.8	2.0	5.6	78.0	11.3
2000–2009	НМ	90	90	43.4	26.0	5.8	6.0	2.5	6.5	69.3	20.8
	AHM	96	90	43.4	26.0	5.8	6.0	2.2	6.5	69.3	20.5
All decades	HM	115	49	31.3	13.1	1.9	0.5	0.1	2.2	44.4	4.7
	AHM	179	49	31.3	13.1	1.8	0.4		2.2	44.4	4.5

 Table 1:
 Selected parameters of the HM and AHM

Note: F-statistic significant at the <0.0001 confidence level. Any discrepancies in totals are due to the effects of rounding.



Figure 3: Harmonics of the AHM

Note: The probable month of conception is shown in brackets after the month of birth.

Figure 4 presents birth seasonality patterns for individual decades and the whole analysed period. In each case two patterns are presented, one calculated by means of HM and the other with the AHM. Patterns calculated by means of HM are identical to those generated with SMM.



Figure 4: Birth seasonality patterns in Poland in the years 1900–2009

Note: The probable month of conception is shown in brackets after the month of birth.

The course of changes in the birth seasonality pattern in successive decades reflects the changes taking place in the acrophase of H_1 . In the first two decades the birth seasonality patterns are almost the same: from January to March the birth rate is relatively stable at a high level (1.08–1.13); in April it starts falling to reach the minimum in June (0.87). Over the next three months the birth rate grows to exceed 1.00 in September and oscillates around this value for three consecutive months. In the next three decades (1920–1949) the pattern slowly changes: the June minimum disappears and the local maximum shows up in September.

The new pattern appears in the 1950s and lasts, almost unchanged, until the end of the 1970s. In this pattern the birth rate reaches the highest values in March (1.09-1.12). A March maximum means that most conceptions happen in June. From March the birth rate gradually decreases (except for September, in which the birth rate increases for a while) to reach the minimum in October/November (0.88-0.92).

The next change in the pattern occurs in the 1980s, when the March maximum slowly disappears, the number of July births increases, and the period between these months is characterized by a fairly even high birth rate. From the 1990s a new pattern develops, with the maximum in July and September (1.04–1.08), a fairly even birth rate from January to June and August (0.99–1.04), and a lower birth rate in October, November, and December (0.88–0.96). The September maximum translates into most conceptions being in the December holiday season, which is the time when family and social life intensifies.

4. Discussion

Looking at changes in the seasonality pattern through the prism of harmonics (harmonics shifting and appearing and disappearing throughout the decades, Figure 3) one can conclude that if the seasonality of births is determined by biological rhythms, they are strongly disturbed by cultural, social, and economic factors. An example, at least as far as the first birth is concerned, is the seasonality of marriages (Bobak and Gjonca 2001; Haandrikman 2004; Régnier-Loilier 2010), the pattern of which changed significantly in Poland during the analysed period (Szukalski 2015, 2018). Determining what these factors are requires further research; however, some observations can be made by analysing the results in the context of specific incidents in Poland's history, or by comparing them with the results of research from other countries.

Before that, however, I would like to address one methodological issue. The patterns obtained by estimating the seasonality pattern using a harmonic model with only two harmonics (Cancho-Candela, Andrés-de Llano, and Ardura-Fernández 2007; Roenneberg and Aschoff 1990a) are generally characterised by a much smaller fit, and

may omit important details (e.g., the local peak in September). If I had used only two harmonics in my study, the median of the R^2 values presented in Table 1 would decrease by 10 percentage points.

The seasonality pattern that occurs in the Polish territory during the first two decades (at that time there was no independent Poland) stands out because it does not appear in any of the 20th century populations analysed by Cowgill (1966a, 1966b) or Lam and Miron (1994, 1996), whereas a very similar pattern appears in the study by Knodel and Wilson (1981), devoted to German villages in the 18th and 19th centuries. This may indicate a link between birth pattern and the agricultural cycle in a period when 65% of the population in the Polish territory worked in agriculture (Gorzelak 2010). The trough in the seasonality pattern corresponds to reduced conception during the harvest period, e.g., due to labour migration of one of the spouses, and to increased conception in periods of low labour demand in agriculture. However, the problem seems to be more complex. First of all, Knodel and Wilson (1981) present four seasonality patterns for legitimate births of second or higher orders (1750–1774, 1800– 1824, 1825–1849, 1875–1899), of which the first three (which differ only in amplitude) are similar to the Polish model for 1900–1919, while the fourth model resembles the Polish and German models from the 1950s. In turn, Ouetelet (1869), in a pioneering work in which he studied the seasonality of births in the Kingdom of the Netherlands in the years 1815–1826, presents two patterns, separately for rural and urban populations. These patterns are very similar to each other and to the Polish pattern for 1900–1919. only with slight differences in amplitude. Studies on Italy covering the years 1862-2011 observe a similar pattern during the first six decades, with a minimum in June (Ruiu and Breschi 2017). If we also take into account data from before and during the industrial revolution in England (Cowgill 1966c; Wrigley and Schofield, in Lam and Miron 1991) and, for example, 19th century Iceland and Austria (Björnsson 2016; Doblhammer, Rodgers, and Rau 1999), we get a full picture of the puzzle still waiting to be solved.

The second model in Poland, with a spring peak and the local peak in September, remained almost unchanged the longest (from 1950 to 1979). Then within a decade it transformed into a new, third model (1990–2009), with peaks in July and September. Both patterns and the course of their transformation are almost identical to those in Germany (Lerchl, Simoni, and Nieschlag 1993); however, the transformation in Germany took place ten years earlier, in the 1970s. In both countries these transformations are connected to a dramatic drop in the fertility rate, which also occurred earlier in Germany than in Poland. While an increase in the availability of contraception seems to be one of the causes of both phenomena, this problem certainly requires further research.

Finally, it is worth noting that the results presented here also confirm previous observations on how wars disturb the seasonality of births. The seasonality patterns for the decades in which wars took place on Polish territory (World War I 1914–1918; Polish-Bolshevik war 1919–1921; World War II 1939–1945) are characterized by the lowest values of R^2 (Table 1). The invasion of Poland in September 1939, which began the Second World War, resulted nine months later, in June 1940, in a 14-percentage-point decrease in births compared to the June average of 1930–1949. Another example of the impact of the political situation distorting the seasonality pattern is the introduction of martial law in Poland in December 1981. In 1982 the September local peak visible in all patterns in Figure 4, measured by the ratio of the number of births in September to the average birth rate in August and October, was the lowest in the whole analysed period.

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