



# DEMOGRAPHIC RESEARCH

*A peer-reviewed, open-access journal of population sciences*

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## ***DEMOGRAPHIC RESEARCH***

**VOLUME 37, ARTICLE 7, PAGES 147–188**

**PUBLISHED 25 JULY 2017**

<http://www.demographic-research.org/Volumes/Vol37/7/>

DOI: 10.4054/DemRes.2017.37.7

*Research Article*

### **Childbearing trends in Iceland, 1982–2013: Fertility timing, quantum, and gender preferences for children in a Nordic context**

**Ari Klængur Jónsson**

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## **Childbearing trends in Iceland, 1982–2013: Fertility timing, quantum, and gender preferences for children in a Nordic context**

**Ari Klængur Jónsson<sup>1</sup>**

### **Abstract**

#### **BACKGROUND**

Iceland is one of the most gender-equal countries in the world, but one that does not seem to have experienced the same fertility fluctuations as most other countries, following the enhanced role of women in society.

#### **OBJECTIVE**

In this study we examine the childbearing trends in Iceland during 1982–2013 by analysing the progressions to parities one, two, and three. We also investigate whether there is evidence of gender preferences for children among Icelandic parents.

#### **METHODS**

Official individual longitudinal register data is used, covering the total female population born in Iceland between 1941 and 1997. The data is analysed by means of event history analysis.

#### **RESULTS**

We find evidence of tendencies to postpone motherhood during the period, with increases in fertility for women in their 30s and 40s. The propensity to have a second and a third child has not declined; on the contrary, these birth intensities have increased since the mid-1980s. Estimates suggest that Icelandic parents prefer to have daughters.

#### **CONCLUSIONS**

During a period of increased educational attainment and postponed family formation, the resilience of Icelandic fertility is intriguing.

#### **CONTRIBUTION**

The study provides the first comprehensive overview of fertility trends in Iceland.

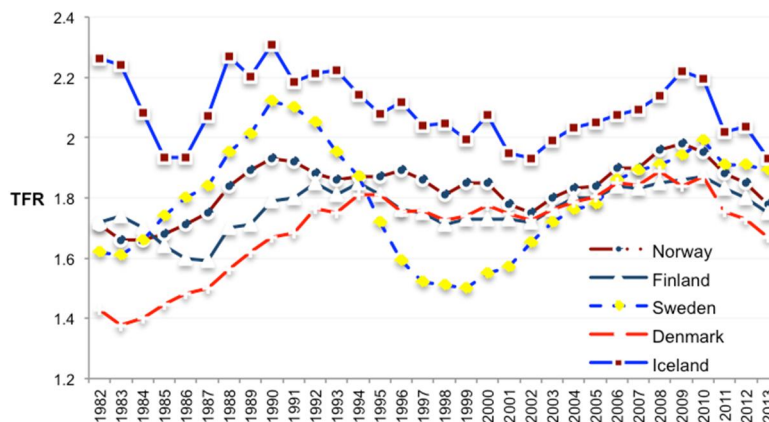
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## 1. Introduction

The aim of the following analyses is to provide the first comprehensive overview of the childbearing trends in Iceland during the last three decades. Iceland is a modern yet insular country with a population of about 330,000. The island is located in the middle of the North Atlantic Ocean, some 1,000 kilometres from the nearest coast of Scandinavia. In many respects it shares the same history and culture, has a similar policy setup, and has experienced many of the same societal changes as in the other Nordic countries: Denmark, Finland, Norway, and Sweden. In some aspects, Iceland can be seen as an ultra-Nordic country. It has fertility levels that exceed those of its Scandinavian neighbours, and is simultaneously a vanguard in terms of childbearing outside of marriage. The latter feature has been a salient attribute of childbearing behaviour for centuries (Björnsón 1969). Around two out of every three births today occur outside marriage, compared to roughly two in five in the early 1980s, and roughly one in four in the 1960s. Nevertheless, the majority of children are born into coresidential unions due to the large proportion of people living in nonmarital cohabitation, another prevailing and long-lasting aspect of family formation in Iceland (Statistical Series 2014; Statistics Iceland 2016; Björnsón 1969).

**Figure 1: Total fertility rate in the Nordic countries, 1982–2013**



Sources: Statistics Norway (2016), Statistics Finland (2016), Statistics Sweden (2016), Statistics Denmark (2016), Statistics Iceland (2016)

Iceland has never experienced periods of the low or lowest-low fertility levels that most of the other European countries have, including the Nordics (Frejka and Sobotka

2008). While all of the other Nordic countries have displayed a total fertility rate (TFR) below 1.7 at some point during the last 30 years, and Denmark and Sweden have even experienced a TFR below 1.5, as of 2013 the Icelandic TFR has never fallen below 1.9. In fact, the country's TFR has flirted with a replacement level of 2.1 children per woman throughout the investigation period, as Figure 1 shows, either just below the mark or just above it. This stability is interesting in light of the many societal and family-related changes that took place during the period and the concurrent postponement of parenthood which is assumed, at least periodically, to be captured by a sizable decrease in the TFR.

In this study, more sophisticated methods than those associated with the TFR will be utilized. As a first in relation to Iceland, we will apply an event history analysis on Icelandic longitudinal individual register data with the purpose of presenting parity-specific childbearing trends. The data covers the total population of women born in Iceland during 1941–1997. This will enable us to produce annual indices of first, second, and third births, presented as the relative risks of giving birth by calendar year, standardized for age and, where appropriate, duration since previous birth. This allows us to investigate the underlying childbearing behaviour – or in other words, to estimate changes in the “force of fertility” over calendar years without many of the shortcomings associated with the TFR (Andersson 1999: 4). Furthermore, on the basis of synthetic cohorts we will estimate whether childlessness has increased over time and investigate whether the proportion of mothers giving birth to a second and a third child has changed. We will also analyse changes in birth spacing patterns over time, which allows us to provide better insight into changes in tempo as well. Finally, we will look for evidence of gender preferences in Iceland. In a fully egalitarian society the gender system should be neutralized, and as a result we should not be able to find daughter or son preferences (Pollard and Morgan 2002). However, the trends observed in Scandinavia suggest that “the role [...] the societal gender system plays in the formation of sex preference for children” is not altogether straightforward (Andersson et al. 2006: 265). Examining gender preferences in Iceland will offer insight into this puzzle.

Together, the different elements give us a comprehensive picture of the childbearing trends over a period of about 30 years in one of the most gender-equal countries in the world, but one that does not seem to have experienced the same (temporary) strain on fertility that most other countries have, following the enhanced role of women in society (Esping-Andersen and Billari 2015). We begin our investigation with a general discussion of the Nordic fertility regime. We provide relevant background information on Iceland before we discuss some of the shortcomings of the TFR and the advantages of our method of analysis. After

deliberation on the data and method, results are presented. Finally, we summarize and reflect upon the main findings.

## 2. Background

### 2.1 The Nordic fertility regime

The Nordic countries, i.e., Iceland, Sweden, Norway, Denmark, and Finland, are commonly grouped together as belonging to the same welfare state regime – the social democratic regime – which, in relative terms, promotes high standards of equality with regard to social rights, state services, and benefits to its citizens (Esping-Andersen 1990). In addition to a shared cultural and historical heritage, all the Nordic countries maintain strong welfare systems promoting social and gender equality. Women’s labour force participation is almost the same as men’s, public support to families is generous, and standards of living are relatively high (Esping-Andersen and Korpi 1987; Esping-Andersen 1990; Korpi and Palme 1998; Esping-Andersen 2002a, 2002b; Eydal and Gíslason 2011). Although each country has its own characteristics, the parental leave schemes, including employment security after childbirth, the education systems, the availability of highly subsidized childcare, and flexible employment opportunities in general, are all important attributes of the Nordic welfare regime (Esping-Andersen 2002b; Eydal and Gíslason 2011).

In light of the relatively high fertility in the region and the high labour force participation among women, an association between the institutions of the welfare regime and fertility outcomes has been suggested on numerous occasions – working both directly and indirectly through other related factors (e.g., Esping-Andersen and Billari 2015; Goldscheider, Bernhardt, and Lappegård 2015; McDonald 2000, 2013; Duvander, Lappegård, and Andersson 2010; Andersson 2008; Garðarsdóttir 2008; Rønsen and Skrede 2008). Several studies have established an association between the social policies in the Nordic countries and childbearing behaviour (e.g., Hoem 1993a; Andersson 1999; Duvander, Lappegård, and Andersson 2010), while others have assessed the impact of the labour market on fertility, often in conjunction with social policies (e.g., Hoem and Hoem 1996; Andersson 2000; Vikat 2002, 2004). To a large extent, this is still uncharted territory in Iceland, but the similarities in timing of family policy reforms<sup>2</sup> and the turning points in the TFR in Figure 1 (e.g., in the late 1980s and at the beginning of the 2000s) suggest that family policies could be associated with childbearing trends – in addition to the covariation with the business cycle during the

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<sup>2</sup> See Eydal and Gíslason 2008 for a detailed discussion about the parental leave reforms in Iceland.

period (Duvander et al. 2016; Jónsson 2016). Moreover, the recent stability in fertility at ‘highest-low’ levels in the Nordics, albeit with some differences between countries over time (Andersson 2004), the stability in completed fertility over cohorts (Andersson et al. 2009), and the evidence of daughter preferences in fertility (Andersson et al. 2006) have led to the notion of a “common Nordic fertility regime” at the turn of the century (Andersson et al. 2009: 339).

A shared finding among three of the four Nordic countries outside Iceland is a development of a preference to have daughters. In the Nordics in general, parents of same-sex children have been shown to have a 20–25% higher propensity to have a third child compared to parents of a boy and a girl, indicating that most parents opt to have at least one child of each sex (Andersson et al. 2006). In the late 1970s in Denmark and the early to mid-1980s in Sweden and Norway, parents of two sons showed a higher propensity to have a third child than did parents of either a son and a daughter or two daughters – which is interpreted as evidence of daughter preferences. In Finland, on the other hand, son preferences have remained stronger; although there were indications of changes in this pattern in the mid- to late 1990s (Andersson et al. 2006; Saarela and Finnäs 2014).

## **2.2 The Icelandic saga: A short introduction**

According to the Economic Forums Global Gender Gap Index, Iceland is one of the most gender-equal countries in the world (Hausmann et al. 2006–2014). Its modern family policies are constructed as to allow parents to combine work and family, and are targeted at increasing gender equality (Eydal 2008). However, over the past decades the Icelandic welfare system has deviated in some important aspects from the welfare systems of the other Nordic countries, and the concept of ‘family policy’ was hardly used in public debate until the 1990s (Eydal and Gíslason 2008). The social expenditure in Iceland, as a percentage of the country’s gross national product, has been lower than the OECD average during most of the period under investigation (OECD 2016a). When it comes to expenditures on benefits and services specifically targeted at families and children, Iceland scores relatively high in an OECD comparison (OECD 2016b); but in a Nordic context, these expenditures were relatively low until the beginning of the 2000s (Nordic Statistics 2016; Ólafsson 1993; Eydal and Gíslason 2008).

Compared to the other Nordic countries, both universal childcare provision and a parental leave scheme were introduced and developed relatively late, and as late as the mid-1980s Iceland (and Norway) offered relatively little support to families with young children. Parents of young children were more dependent on grandparents or other relatives than were parents in other Nordic countries (Eydal 2008; Garðarsdóttir 2008).

Then, during the 1990s and early 2000s, new childcare policies were implemented that allowed Iceland to catch up to its neighbours (Eydal 2008; see Eydal and Gíslason 2011 for an overview of Nordic childcare policies). Nevertheless, even after 2000 the childcare gap – the period from the end of paid parental leave until the child is enrolled in preschool – was still larger in Iceland than in the other Nordics, making it harder for parents to coordinate family life and work (Eydal 2008).

### **2.3 Female labour force participation, educational attainment, and postponement of family formation**

With regard to the lag in implementing institutions that support childbearing and childrearing, the high fertility in Iceland is intriguing (Garðarsdóttir 2008). Female labour force participation increased around the same time here as in the other Nordic countries, with strong increases during the 1960s (Garðarsdóttir 2008). At the beginning of the 1970s, roughly every second woman was active on the labour market. In the 1980s the female employment rate was around 65–70%, and in the 2000s it was around 80% (Statistical Series 2005a; Garðarsdóttir 2008; Statistics Iceland 2016). Furthermore, around 40% of all university students were women in the late 1970s, a figure that had risen to 50% by the mid- to late 1980s and that has remained at roughly 60% since the 2000s. During the period of interest (1982–2013), the absolute number of female university students roughly quadrupled while the number of male students doubled (Statistics Iceland 2016; Statistical Series 2005a).<sup>3</sup> At the same time, aggregate fertility remained high.

Increased educational attainment among women should entail postponement effects on parenthood (Gustafsson 2005; Billari, Liefbroer, and Philipov 2006). In line with expectations, statistics show that people have postponed family formation during the period of interest, when it comes to both entering marriage and becoming a parent. People are marrying later in life, and in the post-2000 decade couples entered marriage in their early to mid-30s, compared to their mid-20s in the 1980s. As with mean age at first marriage, the mean age of women at childbirth increased between 1982 and 2013, the first and last years of our study period. In 1982 the mean age at childbirth, regardless of parity, was 26 years, while in 2013 it was 30. A similar trend is visible when we focus on only primiparas. The mean age of women when becoming a mother for the first time was 22 in 1982 and 27 in 2013. The mean age of fathers has followed a similar pattern during the period, with the men's age two to three years higher than

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<sup>3</sup> For a more detailed account of crude fertility trends, educational attainment, and female labor force participation in a Nordic context, see Garðarsdóttir (2008).



that of the women (Statistical Series 2005b; Garðarsdóttir 2008; Statistics Iceland 2016).

## **2.4 The TFR and different components of fertility**

The postponement of parenthood described above and the stability in the TFR over time do constitute somewhat of an interesting conundrum, as the postponement of first births is expected to result in a decreased TFR (Goldstein, Sobotka, and Jasilioniene 2009; Frejka and Sobotka 2008). As the average age at birth increases, the decrease in the TFR should be proportional to the pace of postponement. An increase by three months per year in the average age at childbirth is estimated to decrease the TFR by 25% from what it would have been if there had been no postponement (Goldstein, Sobotka and Jasilioniene 2009: 673). Between 1987 and 1992 the Icelandic TFR increased by 0.14 points (2.07→2.21), while simultaneously the mean age at first birth increased by almost 15 months (an average of about 3 months per year).<sup>4</sup> The general stability in the TFR over calendar time thus indicates that the fertility recuperation at older ages has mitigated the effects of postponement on the TFR and/or that developments in higher-order birth intensities have counteracted the effects of women's declining propensity to become a mother depending on their age.

The convolution narrated above highlights the lack of transparency associated with the TFR. The TFR does not take into account intervals between births or provide information about parity-specific trends, which ultimately give a more detailed picture of childbearing trends (Ní Bhrolcháin 1992; Hoem 1993b; Andersson 1999; Sobotka and Lutz 2010). The period TFR produces an estimate of how many children a woman is expected to have over her lifetime based on the accumulated age-specific rates in a given time period, usually a calendar year, and is thus based on a synthetic cohort of women (Schoen 2004). As a result, the TFR is sensitive to episodic changes in fertility – i.e., postponements of births and changes in the duration between births – and as a consequence does not reflect or capture the real changes women face over the life course, or take into an account that trends can differ between groups or be caused by compositional changes in the population (Andersson 1999; Hoem 2000; Garðarsdóttir 2008).

Distinguishing between timing and quantum aspects of fertility is pivotal when making assumptions about present trends and future outcomes. Part of the decline in fertility in Europe over the past decades has sometimes been explained as being based on a misinterpretation of childbearing data, stemming from the postponement of births

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<sup>4</sup> The mean age at birth, regardless of parity, increased by roughly 13 months between 1987 and 1992 (Statistics Iceland 2016).

rather than the outright abandonment of having children (Goldstein, Sobotka, and Jasilioniene 2009). Postponement effects can ‘distort’ the TFR, if it is not taken for what it is, and be incorrectly regarded as a cohort indicator instead of a cross-sectional measurement (Ní Bhrolcháin 1992; Sobotka and Lutz 2010). In the same way, recuperations at older ages can have boosting effects on the TFR and lead to a misinterpretation of the efficiency and real effects of, for example, pronatalist policy measures on fertility (Sobotka and Lutz 2010). Furthermore, without adjusting for compositional factors such as age, parity, and duration since previous birth, fertility comparison over time becomes less accurate, the actual group under exposure becomes distorted, and past fertility is allowed to influence future fertility (Ní Bhrolcháin 1987).

### **3. Data and methodology**

#### **3.1 Method**

In the following calculations we will rely on event history analysis and estimate piecewise constant exponential regression models for the progression to first, second, and third births. The model estimates presented are standardized period rates (e.g., Ní Bhrolcháin 1992) that reflect the underlying behaviour of the relevant population at risk. The features of this method allow us to simultaneously account for variations in age as well as other compositional changes and factors that influence fertility, regardless of whether they are fixed or time-varying. Essentially, this method is an improved version of indirect standardization. Rather than looking outside the data for an appropriate age schedule for comparison, this is produced internally by means of maximum likelihood estimation (see Hoem 1993b for a detailed discussion). Furthermore, the method allows us to decompose fertility into parity-specific trends, which again gives us a more detailed picture of the developments, including insight into birth spacing patterns and gender preferences for children (Andersson 1999; Andersson et al. 2006). In order to fulfil the objectives of the study we will thus standardize the Icelandic birth intensities for age and, where appropriate, duration since previous birth, and present them as parity-specific annual indices of relative risks of childbearing. This is based on the same method and figuration Hoem (1993b) and Andersson (1999) apply when displaying childbearing trends in Sweden, and as Andersson and his colleagues (2006) apply when investigating gender preferences for children in the other Nordic countries.

### **3.2 Data**

Icelandic authorities have a long history of gathering information about the country's population and vital events. The first census was conducted in 1703 and as such was the first (preserved) census in the world to cover the entire population of a country. Throughout the 18<sup>th</sup>, 19<sup>th</sup>, and 20<sup>th</sup> centuries several censuses were conducted on an irregular basis, over 20 in total (Statistical Series 2014). Since 1981 no traditional censuses have been conducted, and demographic events are registered as they occur. The current National Registry is largely rooted in the 1981 census.

The Icelandic National Registry records all major demographic events, including all births, deaths, and changes in marital status (including registered cohabitation), as well as any immigration or emigration of each individual ever receiving an identification number in Iceland. Hence, this data source gives access to the complete vital life event history of the Icelandic population with great precision. Demographic events can be combined with other administrative records, such as data from the tax registry, records on educational qualifications, and various other data, which offers great opportunities for demographic research in Iceland.

The data for our calculations is based on longitudinal individual register data derived from the Population Register, Migration Register, and Birth Register of the Icelandic National Registry. The dataset consists of reliable information with monthly accuracy on the demographic events of birth, death (where applicable), and migration history. The data contains information on the sex and birth order of each child. To be included in the study, the subjects had to be women born in Iceland between 1941 and 1997; each woman is followed from when she turns 15 years old until she turns 46. The small population of Iceland makes it possible to include the entire female population in the analyses. This ensures that observed differences are less dependent on statistical power; sampling biases and errors are of no concern.

Minor limitations of the dataset are that cohorts born in 1940 and earlier are missing. The oldest women in the sample were born in 1941, and were thus 41 years old in 1982. Hence, ages 42 and above are missing in 1982, ages 43 and above are missing in 1983, etc. However, this is hardly an issue as we control for age in our analyses. Due to incomplete birth histories for some women or inconsistencies in the birth order records of their children, 2.5% of the original population was left out (approx. 3,000 of 115,000 women).

### 3.3 Models and variables

The birth intensities in the following section are presented as relative risks of giving birth. The risk is dependent on both the number of births and the exposure time at risk. Changes in the relative risk over time can thus reflect changes in either quantum or timing of births, or both. In order to distinguish between the two, we also investigate the developments in birth intervals during our study period, and synthetic cohort measures of the progressions to parities one, two, and three.

In most of the models the birth intensities are given relative to the intensities in 1986 and, where appropriate, to a specific age group. There is no specific reason to choose 1986 over some other year, as the general trends would come out the same regardless of the year of reference. The birth intensities are interchangeably referred to as birth rate or risk in subsequent sections.

Unless otherwise specified, when it comes to first births, the ‘clock’ is set to zero at age 15 and women enter the analysis in 1982 or the month they turn 15, whichever comes last. The main duration variable is the woman’s age, and the time unit is measured in woman-months. We follow childless women until they give birth or are censored – i.e., if they emigrate, die, turn 46, or reach the end of the study period on December 31, 2013, whichever comes first. Women who had their first child before age 15 or before 1982 are excluded from the analyses, as well as everyone who emigrated before 1982 and did not return before their 15<sup>th</sup> birthday.

For second and third births the clock starts at the time of the previous birth. Women enter the analyses at that time or in 1982, whichever comes last, and the main duration variable is measured in months from the previous birth. The duration from the previous birth is categorized into nine groups: The first six years are given in six single-year groups; a duration of 73–96 months is categorized into one group; 97–120 months in another; and 121 months and more in another. The mother’s age is categorized into 12 groups, but the categories differ by parity. Mothers are right-censored on the same grounds as for first births and precensored if they gave birth a second/third time before 1982, had multiple births the first/second time, were older than 45 when they had their previous child, or emigrated before giving birth to their first or second child – depending on parity. Calendar year is a time-varying covariate included in 32 single-year categories in all parity-specific models (1982–2013).

The same criteria apply when gender preferences are investigated as they appear in third-birth rates but an additional variable is included, measuring the sex composition of the two previously born children (i.e., boy + girl; boy + boy; girl + girl). The interpretation of the relative risks and how they relate to gender preferences is based on the difference in risk by the sex composition of the previous children. If the risk were the same regardless of the sex composition, we would conclude that there are no preferences. The approach is based on the assumption that parents who have not

acquired the desired gender composition among their children will be more likely to continue their childbearing (see Hank and Kohler 2000; Pollard and Morgan 2002; Andersson et al. 2006 for discussion). Background statistics on the distribution of live births and exposure times under risk by variable and parity are available in the Appendix.

## **4. Results**

### **4.1 First births**

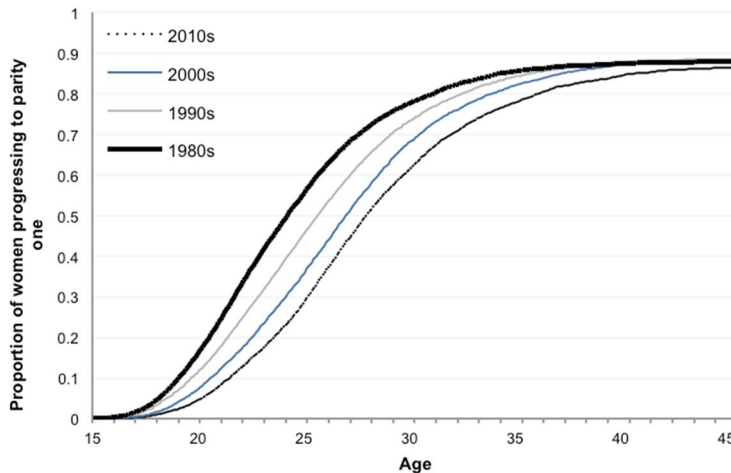
#### **4.1.1 Tempo or quantum: Progression to parity one by decade**

Figure 2 presents a graph of Kaplan–Meier (nonparametric) estimates displaying the cumulative probabilities of synthetic cohorts of childless Icelandic women progressing to parity one before turning 46, by decade: the 1980s, 1990s, 2000s, and 2010s. The estimates indicate that childlessness has not increased in Iceland during the period; at age 45 it ranges between 11 and 13% depending on the decade<sup>5</sup>. However, there are signs of continuous postponement as the estimates suggest that the age at which half of the women have progressed to motherhood has increased from one decade to the next. In the 1980s half of the women had given birth at age 24, while during 2010–2013 half had become mothers at age 27 (and 9 months). Hence, Icelandic women have delayed motherhood by more than a year per decade, but approximately the same number of women gave birth. Nevertheless, as illustrated in Figure 3, the postponement over time results in a decrease in the estimated first-birth rate.

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<sup>5</sup> The cumulative probability estimates should be taken with a grain of salt as they are not standardized for the impact of any compositional changes beyond that of age. They may thus give a somewhat distorted comparison between periods (Ní Bhrolcháin 1992).

**Figure 2: First-birth Kaplan–Meier cumulative probability estimates in Iceland 1982–2013 by age of woman over decades**

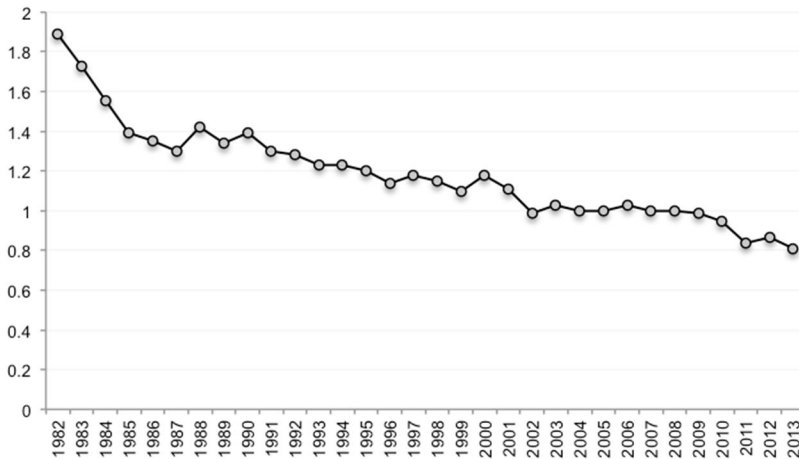


Source: Icelandic register data, author's calculations

#### 4.1.2 The force of fertility: Annual index of first birth

Figure 3, displaying the standardized annual index of first-birth rates relative to the year 2004, with birth intensities standardized for age, demonstrates that the first-birth rate has declined almost continuously throughout the period. Compared to the year 2004, the birth rate was around 90% higher in 1982, and almost 20% higher in 2000. The propensity to become a mother was quite stable between 2002 and 2009, but in 2010 and onward we observe a similar trend of decreasing birth intensities re-emerging as was evident during the 1980s and 1990s. In 2013, the first-birth intensities were about 80% of what they had been in 2004. Women now spend a longer time under risk, which explains the declining intensities over time and, hence, the figure depicts a change in timing but not quantity. This makes it imperative to study the age patterns of first-birth rates.

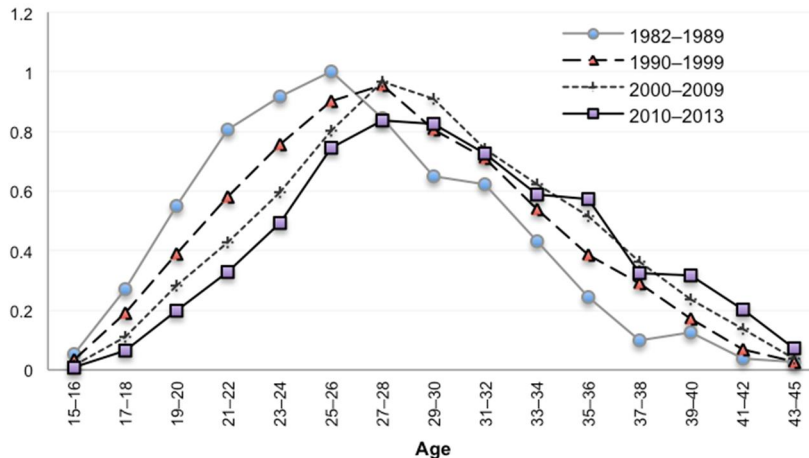
**Figure 3: Relative risk of first birth in Iceland 1982–2013. Standardized for age of woman. Rates are relative to first-birth rates in 2004**



Source: Icelandic register data, author's calculations

Figure 4 demonstrates this more clearly. We can see that the curves, displaying birth intensities by decade and age, have shifted to the right, revealing that women tend to have their first child later than before and, on the aggregate level, more evenly over the childbearing ages than before. These findings are in line with previous work by Garðarsdóttir on age-specific fertility rates in Iceland (2008). The estimates in Figure 4 also tell us that during the study period the propensity to become a mother decreased for women aged 15–26 while it increased for women aged 29–45. The birth intensities at ages 27 and 28, however, were basically the same in the first and last years of observation.

**Figure 4: First-birth rates in Iceland by age of woman 1982–2013. Rates are relative to first-birth rates of 25–26-olds in the 1980s**



Source: Icelandic register data, author's calculations

### 4.1.3 Annual indices of first birth by age groups

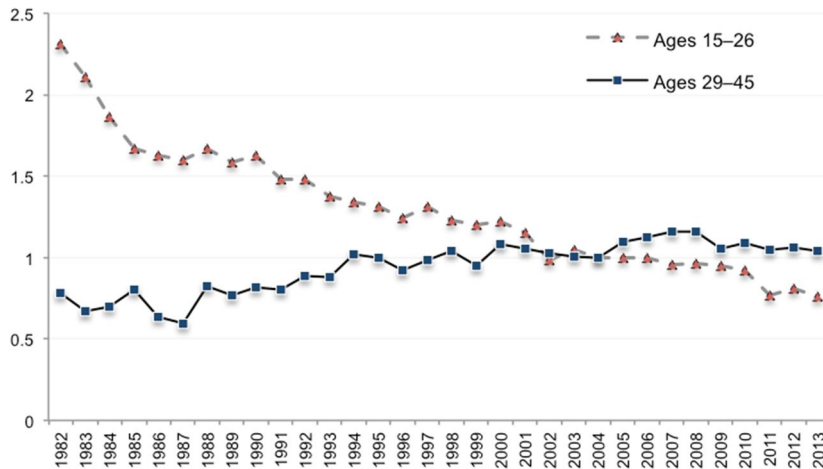
Figure 5 describes this scenario in more detail. Due to the multiplicative nature of the age specific trends, separate models were used for each of the two age groups displayed in the figure: 15–26 years and 29–45 years. The birth intensities at ages 27 and 28 developed differently than the age-specific birth risks in any of these two age groups. As a result they do not belong to either of the model specifications and are excluded from the presentation. As before, the rates are standardized for variation in age (within models) and are relative to the rates in each age group in 2004.<sup>6</sup>

In Figure 5 we can see the same trend of declining birth intensities for ages 15–26 throughout the period like in Figure 3, but with more force. The rate fell sharply between 1982 and 1986. After a relatively stable period between 1987 and 1990 it continued to fall at a somewhat slower rate until 2002, when the declining trend decelerated even further. The first-birth rate was approximately 130% higher in 1982 but around 25% lower in 2013 compared to 2004, indicating a vast decrease in the propensity to become a mother at ages 15 to 26.

<sup>6</sup> The choice of baseline year is entirely arbitrary, but the magnitude of intensities in this year were quite evenly distributed between ages 15–26 and 29–45, respectively.



**Figure 5: Relative risk of first birth in Iceland 1982–2013. Standardized for age of woman, two age groups. Rates are relative to first-birth rates in 2004 for each age group (separate models)**



Source: Icelandic register data, author's calculations

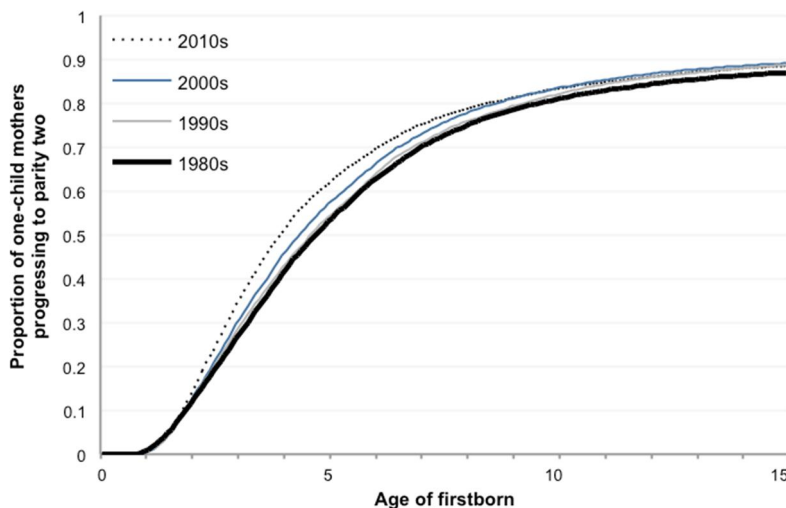
In contrast, the propensity to become a mother for women aged 29 to 45 increased (Figure 5). There are more fluctuations in the birth rate for the older women from one year to the next, but the trend nonetheless describes a general increase in fertility for women in their 30s and 40s. From 1986–1987 onward, birth intensities increased gradually until they reached a peak in 2007–2008 – when the propensity to become a mother was around 15% higher than in 2004. The birth intensities at the end of the study period, in 2013, were about 60% higher than in 1986 and still almost 5% higher than in 2004. The development in first births during the period is thus best described by forceful postponement at younger ages in tandem with recuperation at older ages. The postponement in first-birth fertility still produces a larger population of women under risk to become a mother, and thus the decline in aggregate first-birth intensities observed in Figure 3.

## 4.2 Second and third births

### 4.2.1 Tempo or quantum: Progression to parities two and three by decade

As with first births, the second-birth Kaplan–Meier cumulative probability estimates, derived from synthetic cohorts of Icelandic women, do not suggest much change in the final progression rate over the period. According to the estimates, between 87 and 89% of one-child mothers will have a second child by the time their firstborn turns 15, depending on decade (Figure 6). However, the estimates also suggest that the interval between first and second births was smaller in 2010–2013 than in previous decades. Half of the women had given birth to a second child by the time their first child was four years old during 2010–2013, compared to almost five years old in the 1980s.

**Figure 6: Second-birth Kaplan–Meier cumulative probability estimates in Iceland 1982–2013 by age of firstborn over decades**

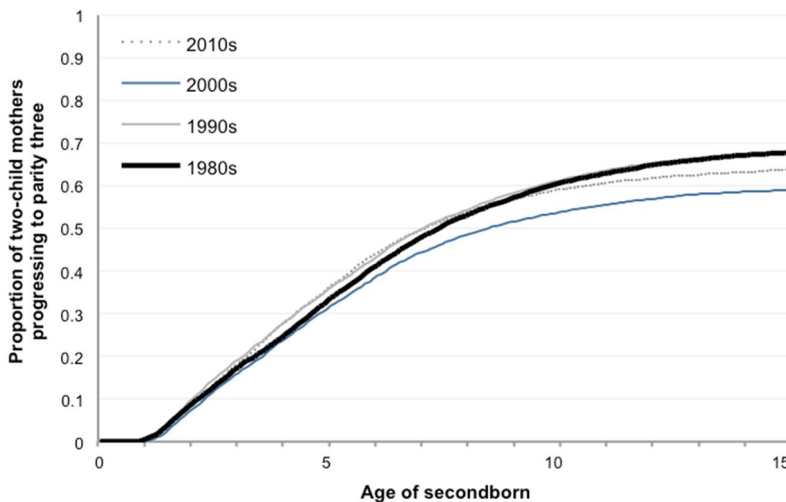


Source: Icelandic register data, author's calculations

According to corresponding third-birth Kaplan–Meier cumulative probability estimates, by the time a secondborn child turns 15, around 67% of two-child mothers have given birth to a third child (Figure 7). By the time the secondborn is seven and a half years old, 50% of two-child mothers have already progressed to parity three. The estimates suggest minor differences in the timing between births by decade, and

relatively small differences in the final levels of women giving birth for the third time – another indication of stability in fertility over time, as measured based on synthetic cohorts. Nevertheless, to compare similar groups over time we will need to apply standardization in order to depict changes in the force of fertility during the period.

**Figure 7: Third-birth Kaplan–Meier cumulative probability estimates in Iceland 1982–2013 by age of secondborn over decades**



Source: Icelandic register data, author's calculations

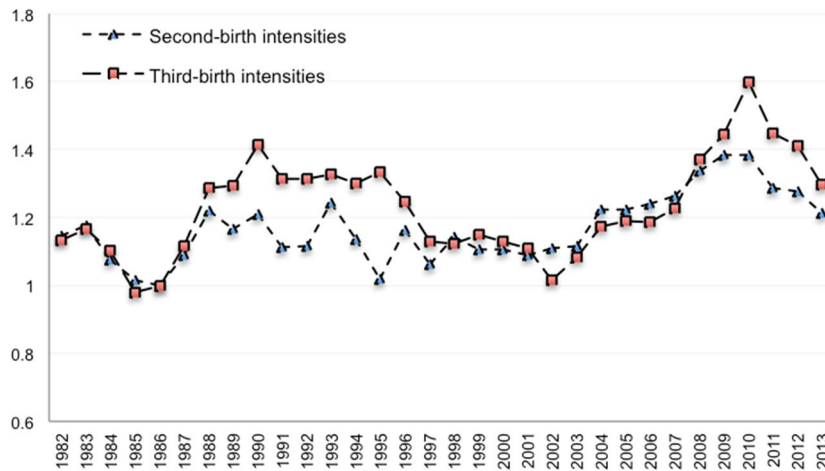
#### 4.2.2 The force of fertility: Annual indices of second and third births rates

Figure 8 features the standardized annual indices of second- and third-birth rates between 1982 and 2013 derived from two separate models. The rates are presented as birth intensities relative to the year 1986 for each parity, and are standardized for the mother's age and the duration since the previous birth over calendar years. Hence, both are annual indices only comparable to the specific birth order, but do not tell us the absolute differences in the propensity to give birth to a second and a third child.

Although there are some fluctuations in both second- and third-birth rates from one year to the next, there are no indications that the propensity to have a second or a third child has declined during the period (Figure 8). The second-birth rate reached a peak in 2009 and 2010, when one-child mothers had almost 40% higher birth intensities

than those at comparable ages in 1986. During the period from 2004 to 2013 the birth intensities were on average more than 20% higher than in 1986, with a gradual increase from one year to the next between 2003 and 2010. In 2011 the second-birth rate started to fall, but the propensity to have a second child was still 20% higher in 2013 than in 1986.

**Figure 8: Relative risk of second and third births in Iceland 1982–2013. Standardized for age of mother and age of youngest child. Rates are relative to the rates in 1986 for each birth order (separate models)**



Source: Icelandic register data, author's calculations

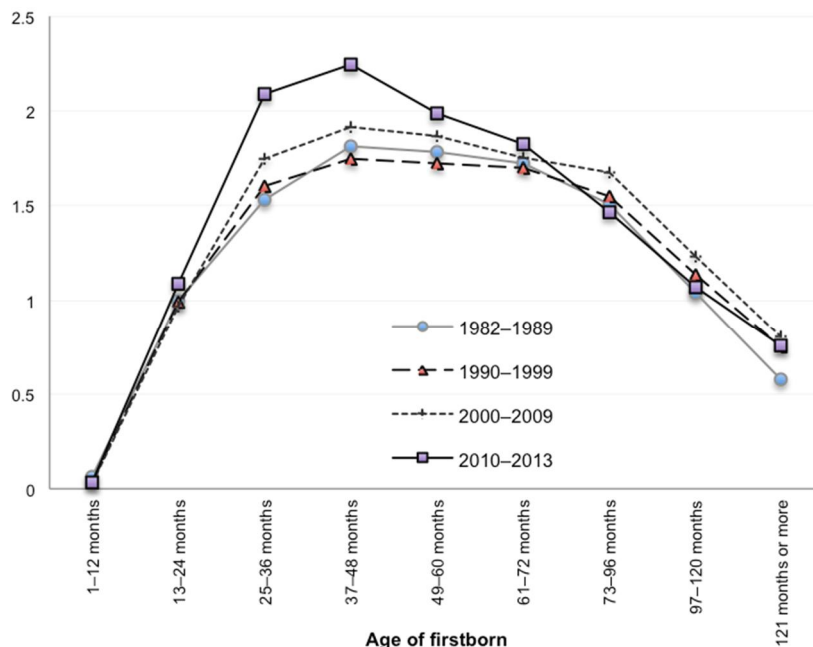
In relative terms, there are slightly ‘deeper waves’ in the third-birth rate compared to the second-birth rate (Figure 8). There are two points of upsurge during the period: between 1987 and 1990 and between 2003 and 2010. The most striking decline in the rates occurs between 1996 and 2002, and again between 2011 and 2013. During the former upswing the rate reaches its high point in 1990, when two-child mothers had some 40% higher propensity to have a third child than in 1986. During the latter upswing the age-standardized third-birth intensities reached a peak in 2010, when the propensity to have a third child was around 60% higher than in 1986. In 2011–2013 the rate fell from one year to the next, but the birth intensities were still about 30% higher in 2013 than in 1986.

In order to get a better understanding of these developments, we will look at changes in birth spacing over time. This gives us more insight into whether the increased birth rates are related to changes in the timing of births or changes in levels.

### 4.2.3 Changes in birth spacing over time

When we examine the birth spacing patterns and their development by decade between 1982 and 2013 (standardized for the age of mother, Figure 9a), we observe an earlier arrival of the second child in the last years of our study period.

**Figure 9a: Second-birth rates in Iceland 1982–2013 by time since previous birth. Standardized for age of mother. Rates are relative to duration 13–24 months in the 1980s**

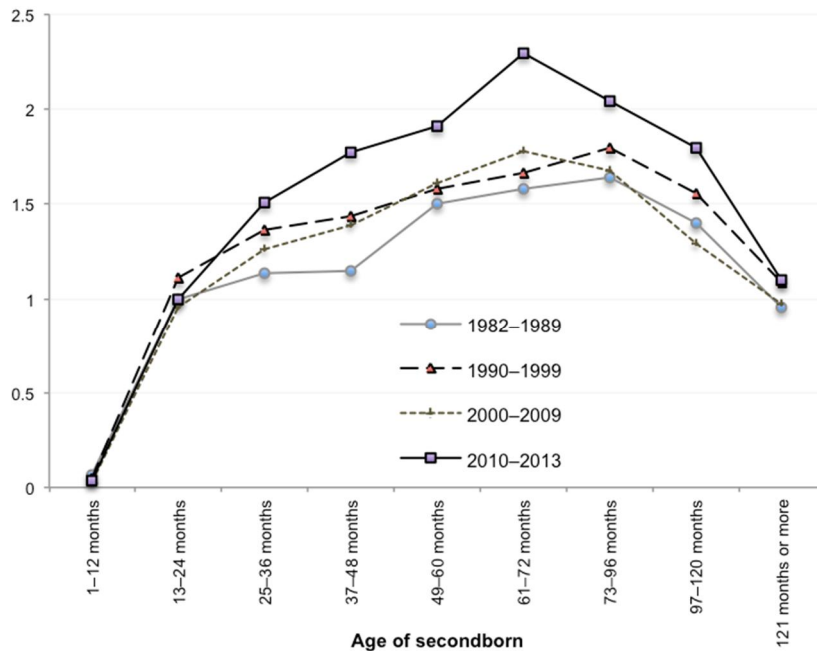


Source: Icelandic register data, author's calculations

The propensity to have a second child was higher in the first six years from previous birth during the post-2000 period compared to the pre-2000s, and especially so

in the last years of observation. In other words, the interval between first and second births was smaller in 2010–2013 than in previous decades.

**Figure 9b: Third-birth rates in Iceland 1982–2013 by time since previous birth. Standardized for age of mother. Rates are relative to duration 13–24 months in the 1980s**



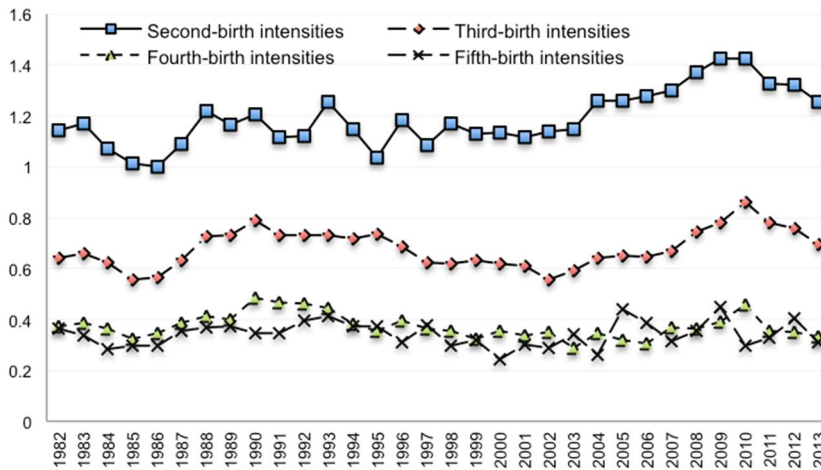
Source: Icelandic register data, author's calculations

Similarly, after we have standardized the third-birth rates for mother's age, an examination of the birth spacing between the second and the third child reveals changes in the third-birth interval during the last years of observation (Figure 9b). The patterns during the 1980s, 1990s, and 2000s are very similar. The intensity is somewhat lower in the first five years from previous birth during the 1980s compared to later decades, but overall the patterns are relatively homogeneous. However, during 2010–2013 we observe a new pattern. Compared to earlier decades and at comparable ages, two-child mothers had a higher propensity to have a third child at practically all durations since

the second birth during this period. To a large extent, this increase in third-birth rates corresponds to a shift in the age distribution of two-child mothers to higher ages.<sup>7</sup>

#### 4.2.4 Second- and third-birth rates relative to higher-order births

**Figure 10: Relative risk of second, third, fourth, and fifth births in Iceland 1982–2013. Standardized for age of mother and age of youngest child. Rates are relative to second-birth rates in 1986**



Source: Icelandic register data, author's calculations

An overview of the combined annual indices of second, third, fourth, and fifth births, standardized for the mother's age and the duration since the previous birth over calendar year, reveals striking similarities in trends among the different parities of mothers. All rates are relative to the second-birth intensities in 1986, and in Figure 10 we can see how the third- and fourth-birth rates covary with the second-birth rate over time. There is more random variation in the fifth-birth rate, but it more or less follows suit. This observation is not unique to Iceland, as a similar trend has been observed in

<sup>7</sup> The estimations derived from the third-birth multivariate model in the graph in Figure 9b have been standardized for the mother's age, which explains what seems to be a controversy between the cumulative probabilities given in Figure 7 by decade and the hazards by decade in Figure 9b. When the mother's age is not included in the model, the estimations of the hazard are not visibly higher in the 2010s than in the 1980s (estimates not shown but available upon request).

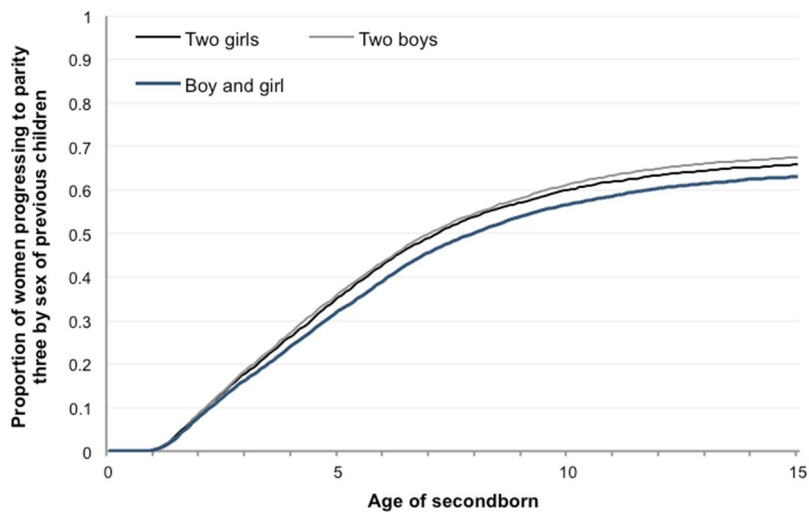
the case of Sweden, where the fluctuations in the third- and fourth-birth rates correspond closely to those in the second-birth rate (Andersson and Kolk 2015).

On average during the period, in relation to the second-birth rate, the third-birth rate was 43% lower, the fourth-birth rate was 68% lower, and the fifth-birth rate was 71% lower. The third-birth intensities are thus between those for second and fourth births, while the fourth and fifth birth intensities are at the same levels. Naturally, the populations under risk become increasingly small with increasing parity.

### 4.3 Gender preferences

#### 4.3.1 Progression to parity three by sex of previously born children

**Figure 11: Third-birth Kaplan–Meier cumulative probability estimates in Iceland 1982–2013 by age of secondborn over sex composition of previously born children**



Source: Icelandic register data, author's calculations

In our analysis of gender preferences for children in Iceland, we start out with a nonparametric description as before. The synthetic cohorts' Kaplan–Meier cumulative probability estimates (Figure 11) indicate a slight difference in the progression to parity three depending on the sex of the previously born children. Already three years after



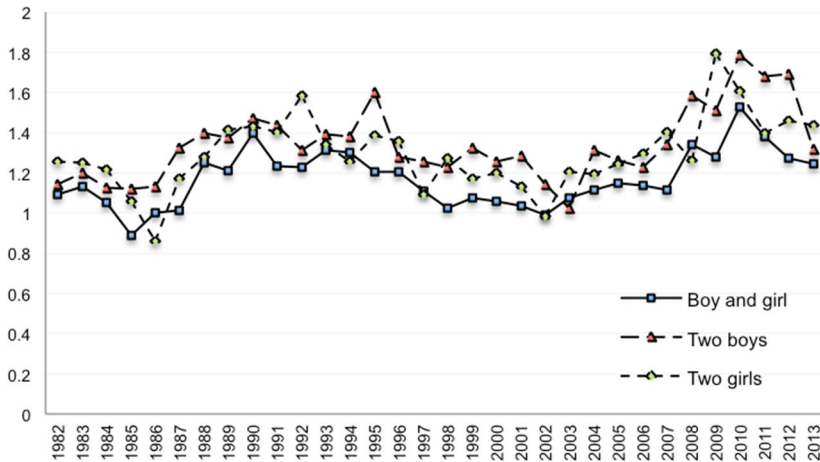
previous birth, there are indications that mothers of a boy and a girl progress more slowly to parity three than do mothers of same-sex children. Furthermore, at the end of the follow-up, fewer women with one child of each sex have had a third child. The lines do not converge, indicating that there is a difference not only in timing but in ultimate fertility as well.

When it comes to mothers of two daughters on the one hand and those of two sons on the other, a higher percentage of mothers of two sons have a third child. In the first six to seven years from the birth of the second child, both groups of women progress at a similar rate to parity three, but after this the two lines diverge. By the time the secondborn turns 15, 63% of mothers with a son and a daughter have had a third child, compared to 66% of mothers of two girls and 67.5% of mothers of two sons.

#### 4.3.2 Annual indices of third births by sex of previously born children

Multivariate analyses support these indications of mixed-sex and daughter preferences. Figure 12 captures the trend over time with rates relative to those of a mother with a boy and a girl in 1986.

**Figure 12: Relative risk of third births in Iceland 1982–2013 by sex composition of previous children. Standardized for age of mother and age of secondborn. Rates are relative to mothers of a boy and a girl in 1986**



Source: Icelandic register data, author's calculations

Figure 12 describes an underlying pattern of mixed-sex and daughter preferences throughout the period, but with a great deal of random variation. On average during 1982 and 2013, mothers of two girls had a 10% higher risk of having a third child compared to mothers of a boy and a girl, while mothers of two boys had a 14% higher risk. In other words, if we look at the propensity to have a third child among mothers of two boys on the one hand and those of two girls on the other, relative to one another, mothers of two boys showed an almost 4% higher risk than did mothers of two girls. Hence, we can assert that there is evidence of mixed-sex as well as slight daughter preferences for children in Iceland at the turn of the century.

## 5. Summary and discussion

In European demography, Iceland is an outlier with a fertility rate that far exceeds that of most countries on the continent. In times of low and declining fertility levels in many regions of the world, Iceland's stability in fertility patterns across three decades makes the country an interesting case to study.

After changes in age schedules and past fertility have been accounted for, we do not observe any dramatic changes in the underlying childbearing behaviour across calendar time. What we do observe is a change in timing, especially with regard to first births. The propensity to become a mother has declined for women in their early and mid-20s – depicting a trend of gradual postponement of motherhood over time. During a period of increased educational attainment and postponement of family formation in general this is expected. The decline in first-birth intensities at the early ages is paralleled with increases for women in their 30s and 40s: an indication that fertility recuperation is taking place. We find no indications of increased childlessness. The cumulative probability estimates for synthetic cohorts of Icelandic women indicate that nine out of ten women become a mother, regardless of decade.

At the same time, the standardized second- and third-birth rates have fluctuated around remarkably stable averages of progression to parities two and three, and there are no obvious indications that the force of fertility has declined over the period when we compare similar groups over time. On the contrary, the second- and third-birth intensities have not only maintained their levels, but when we compare the early years of the study period to the last, we find that birth intensities were somewhat stronger in the latter.

There have also been some changes in birth spacing patterns over time, but mainly during the last years of observation. During 2010–2013 there was a change in tempo with regard to having a second child, and two-child mothers had a higher propensity to

have a third child at almost all durations since the second birth. Overall, the stability in second- and third-birth rates, also when estimated based on synthetic cohorts, is compelling: On average during the period, the estimates show that almost nine out of ten one-child mothers have a second child, and two out of three two-child mothers have a third child.

Although we have refrained from assessing any relationships, it is not unlikely that the timing of policy reforms has helped stabilize fertility trends and even helped increase fertility. We find strong indications of period effects in the late 1980s to the early 1990s, and also during most of the 2000s. The increases in the standardized second- and third-birth rates between 1987–1990, and again between 2004–2010, correspond to the timing of reforms made to the parental leave scheme. We observe a similar trend for fourth- and fifth-birth as well, and even a slight increase in the first-birth rate during the former period. The extent to which certain social and family policies have influenced the parity-specific birth intensities are topics for future research (cf. working paper by Duvander et al. 2016). The same holds for the role of any influence of the business cycle on the childbearing trends (cf. working paper by Jónsson 2016). Indeed, the economic shock that hit Iceland in 2008 is likely to have affected the behaviour in the very last years of observation – depicted by the declining birth intensities after 2010 (Figures 3 and 8).

We speculated earlier that possible explanations for the conundrum surrounding the stability of the TFR, given the increase in the average age at first birth, were that higher-order birth intensities had compensated for the postponement effects on the TFR, and that a process of recuperation in first-birth fertility was also underway. The analyses offer some support for these claims, although we should be careful when comparing different fertility measures, as the two methods are not fully comparable. Our period measures, stemming from regression models, are standardized for compositional changes in the actual population under risk.

Why we find a preference for having daughters is open to debate. It might be related to a combination of the ‘caring role’ of women and daughters and the strong gender egalitarianism in the Nordic context. According to such an argument, as formulated by Brockmann (2001), women and daughters could be seen as both earners and more reliable caregivers. The latter would (albeit arguably, as the welfare system should simultaneously minimize women’s caregiver role) give them more value over men. A daughter preference might also be associated with increased autonomy and opportunities for women and mothers, and “their relative bargaining power within the household” being stronger than before, as Saarela and Finnäs have suggested (2014: 52). However – and this is an important caveat – due to the time frame of the analysis we cannot establish that the daughter preferences in Iceland are a new or recent phenomenon. Hence, we cannot establish a clear association between the modern

Nordic welfare model and the emergence of daughter preferences, as we do not know the onset of these preferences.

Nevertheless, with all things considered – the postponement of motherhood in tandem with increased fertility at older ages, the persistently strong second- and third-birth intensities, the overall stability in the childbearing trends over time, and the evidence of daughter preferences – we believe we are safe in locating Iceland at the higher end of a common Nordic fertility regime. Iceland deviates from its Nordic neighbours in terms of its higher fertility, but most of the basic parameters of its childbearing dynamics appear to be similar to those in the other Nordic countries. However, it remains for the future to tell whether Iceland will remain a high-fertility country in the decades to come.

## **6. Acknowledgements**

Many of the approaches applied in this manuscript were first proposed and presented by the late Jan M. Hoem. I am indebted to him for his inspiration. Special gratitude goes to Gunnar Andersson for his valuable help, comments, and guidance throughout the work. I would also like to thank Gerda Neyer for her helpful input, and Ólöf Garðarsdóttir for her assistance. Juho Härkönen and Sofi Ohlsson-Wijk provided me with useful comments along the way, as well as my anonymous reviewers. I am grateful for the financial support received from the Swedish Research Council (Vetenskapsrådet) via the Swedish Initiative for Research on Microdata in the Social and Medical Sciences (SIMSAM), grant 340–2013–5164, and the Data Protection Board (Trúnaðargagnanefnd) at Statistics Iceland for approving this study (Application 7/2014). Finally, I am grateful for language editing by Judith Rinker Öhman.

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## Appendix

### Background statistics

**Table A-1: Background statistics: Distribution of live births and exposure times under risk by variables and parity**

Variable		First births		Variable			
Age	Live births	First-birth model		Calendar year	Live births	Woman-months count	%
		Woman-months count	%				
15	100	685,163	8.5	1982	1,631	214,935	2.7
16	449	754,679	9.4	1983	1,530	217,729	2.7
17	1,170	743,233	9.2	1984	1,419	220,256	2.7
18	2,100	719,152	8.9	1985	1,330	223,822	2.8
19	2,825	681,899	8.5	1986	1,320	225,177	2.8
20	3,521	628,016	7.8	1987	1,310	229,349	2.8
21	3,810	561,551	7.0	1988	1,462	233,842	2.9
22	3,773	494,818	6.1	1989	1,411	236,095	2.9
23	3,741	429,704	5.3	1990	1,484	237,707	2.9
24	3,563	366,526	4.5	1991	1,402	240,140	3.0
25	3,387	307,652	3.8	1992	1,422	242,159	3.0
26	2,996	255,421	3.2	1993	1,387	243,470	3.0
27	2,541	210,917	2.6	1994	1,407	247,164	3.1
28	2,124	173,545	2.2	1995	1,388	250,725	3.1
29	1,563	143,663	1.8	1996	1,328	252,412	3.1
30	1,278	119,624	1.5	1997	1,383	253,941	3.1
31	969	100,246	1.2	1998	1,378	255,764	3.2
32	774	86,246	1.1	1999	1,317	257,508	3.2
33	588	74,531	0.9	2000	1,444	257,779	3.2
34	443	65,612	0.8	2001	1,361	257,055	3.2
35	352	58,715	0.7	2002	1,230	256,979	3.2
36	293	52,975	0.7	2003	1,291	258,988	3.2
37	193	48,405	0.6	2004	1,249	261,842	3.2
38	161	45,043	0.6	2005	1,265	266,770	3.3
39	144	42,276	0.5	2006	1,307	271,007	3.4
40	82	40,550	0.5	2007	1,284	273,480	3.4
41	69	38,661	0.5	2008	1,305	277,411	3.4
42	37	36,465	0.5	2009	1,298	279,874	3.5
43	19	34,698	0.4	2010	1,268	280,709	3.5
44	12	33,157	0.4	2011	1,130	281,445	3.5
45	19	31,646	0.4	2012	1,206	283,997	3.5
				2013	1,149	275,258	3.4
<b>No. of subjects: 79,119</b>		<b>No. of woman-months: 8,064,789</b>		<b>No. of live births: 43,096</b>			

Table A-1: (Continued)

Variables	Second and third births					
	Second-birth model			Third-birth model**		
Age groups	Live births	Woman-months count	%	Live births	Woman-months count	%
<=18/<=20*	126	50,592	1.6	40	12,918	0.3
19-21/21-23*	2,035	299,977	9.3	589	113,636	2.9
22-24/24-26*	6,300	574,791	17.9	2,321	333,943	8.4
25-26	5,898	443,073	13.8	2,892	358,891	9.1
27-28	6,506	413,365	12.9	3,699	429,449	10.8
29-30	5,690	337,240	10.5	4,033	451,494	11.4
31-32	4,162	253,112	7.9	3,628	429,960	10.9
33-34	2,647	191,357	6.0	2,748	389,715	9.8
35-36	1,516	152,770	4.8	1,522	359,554	9.1
37-38	915	130,009	4.0	794	338,893	8.6
39-40	434	117,511	3.7	275	315,673	8.0
41-45	278	251,114	7.8	119	427,422	10.8
<b>Duration from birth</b>						
1-12 months	192	506,763	15.8	140	437,638	11.0
13-24 months	4,799	478,846	14.9	2,946	418,946	10.6
25-36 months	6,957	401,281	12.5	3,212	377,130	9.5
37-48 months	6,256	318,426	9.9	2,972	338,285	8.5
49-60 months	4,793	248,304	7.7	2,899	299,234	7.6
61-72 months	3,688	196,077	6.1	2,590	263,179	6.6
73-96 months	4,754	278,617	8.7	3,796	434,189	11.0
97-120 months	2,315	193,021	6.0	2,074	349,176	8.8
+120 months	2,753	593,576	18.5	2,031	1,043,771	26.3
<b>Sex composition</b>						
Boy and girl				10,990	2,040,536	51.5
Two boys				6,215	1,007,939	25.5
Two girls				5,430	911,013	23.0
<b>Calendar year</b>						
1982	1,247	106,937	3.3	786	114,934	2.9
1983	1,283	109,437	3.4	825	119,546	3.0
1984	1,199	110,880	3.4	784	123,594	3.1
1985	1,150	111,325	3.5	706	127,621	3.2
1986	1,131	111,062	3.5	719	130,951	3.3
1987	1,210	109,958	3.4	795	133,155	3.4
1988	1,313	108,662	3.4	892	133,281	3.4
1989	1,221	107,351	3.3	878	132,977	3.4
1990	1,241	106,045	3.3	925	131,254	3.3
1991	1,135	106,187	3.3	837	129,791	3.3
1992	1,153	106,823	3.3	824	128,765	3.3
1993	1,289	106,600	3.3	816	129,367	3.3

**Table A-1: (Continued)**

Variables	Second and third births			Third-birth model**		
	Second-birth model			Third-birth model**		
Calendar year	Live births	Woman-months count	%	Live births	Woman-months count	%
1994	1,168	105,253	3.3	792	129,550	3.3
1995	1,051	105,388	3.3	800	127,517	3.2
1996	1,199	104,709	3.3	724	125,919	3.2
1997	1,081	102,986	3.2	643	125,287	3.2
1998	1,150	102,703	3.2	634	125,087	3.2
1999	1,097	101,731	3.2	638	125,029	3.2
2000	1,097	101,682	3.2	617	124,696	3.1
2001	1,078	102,197	3.2	596	123,805	3.1
2002	1,100	100,700	3.1	537	123,277	3.1
2003	1,090	98,448	3.1	568	123,015	3.1
2004	1,165	96,869	3.0	606	122,995	3.1
2005	1,132	94,469	2.9	605	122,497	3.1
2006	1,113	92,904	2.9	596	121,828	3.1
2007	1,100	91,424	2.8	605	121,285	3.1
2008	1,126	89,356	2.8	662	120,579	3.0
2009	1,138	88,063	2.7	682	118,593	3.0
2010	1,104	86,150	2.7	732	115,908	2.9
2011	1,015	84,711	2.6	648	112,376	2.8
2012	998	82,404	2.6	625	109,852	2.8
2013	933	81,497	2.5	563	107,217	2.7
<b>Second births, no. of subjects: 51,103</b>			<b>Third births, no. of subjects: 45,824</b>			
<b>No. of woman-months: 3,214,911</b>			<b>No. of woman-months: 3,961,548</b>			
<b>No. of live births: 36,507</b>			<b>No. of live births: 22,660</b>			
Variables	Fourth and fifth births			Fifth-birth model		
Age groups	Fourth-birth model			Fifth-birth model		
	Live births	Woman-months count	%	Live births	Woman-months count	%
19–21	5	1,388	0.0	–	–	–
22–24	92	19,317	0.6	7	1,358	0.1
25–26	215	47,563	1.6	26	4,595	0.5
27–28	559	99,701	3.3	49	14,205	1.5
29–30	788	173,101	5.8	139	30,519	3.1
31–32	1,184	253,102	8.4	225	53,836	5.5
33–34	1,384	322,246	0.8	278	84,038	8.6
35–36	1,297	371,762	2.4	326	114,496	11.8
37–38	943	394,205	1.2	284	139,642	14.4
39–40	503	398,507	1.3	223	153,551	15.8
41–45	310	915,892	3.6	166	375,764	38.7

**Table A-1: (Continued)**

Variables	Fourth and fifth births					
	Fourth-birth model			Fifth-birth model		
Duration from birth	Live births	Woman-months count	%	Live births	Woman-months count	%
1–12 months	89	274,744	9.2	28	91,200	9.4
13–24 months	1,214	267,906	8.9	357	89,866	9.2
25–36 months	1,190	252,319	8.4	302	86,120	8.9
37–48 months	933	237,933	7.9	220	82,779	8.5
49–60 months	791	223,484	7.5	182	79,245	8.2
61–72 months	704	210,736	7.0	149	75,839	7.8
73–96 months	1,058	385,907	12.9	225	138,839	14.3
97–120 months	627	337,884	11.3	126	116,923	12.0
+120 months	674	805,880	26.9	134	211,233	21.7
<b>Calendar year</b>						
1982	283	79,093	2.6	78	25,941	2.7
1983	298	85,095	2.8	73	28,352	2.9
1984	283	91,061	3.0	62	31,163	3.2
1985	254	96,326	3.2	65	33,511	3.4
1986	267	101,219	3.4	64	35,676	3.7
1987	292	104,146	3.5	75	37,127	3.8
1988	301	106,062	3.5	76	37,392	3.8
1989	287	107,231	3.6	75	37,511	3.9
1990	344	108,451	3.6	68	37,911	3.9
1991	330	108,681	3.6	68	38,304	3.9
1992	326	108,461	3.6	79	38,446	4.0
1993	312	107,598	3.6	83	38,434	4.0
1994	265	106,711	3.6	75	37,927	3.9
1995	242	105,947	3.5	73	37,164	3.8
1996	267	104,693	3.5	58	35,901	3.7
1997	237	102,170	3.4	67	35,079	3.6
1998	224	100,050	3.3	51	33,879	3.5
1999	196	98,126	3.3	52	33,078	3.4
2000	206	96,250	3.2	37	31,980	3.3
2001	186	94,356	3.1	43	30,843	3.2
2002	184	92,120	3.1	39	29,999	3.1
2003	145	89,623	3.0	43	28,632	2.9
2004	165	87,505	2.9	30	27,221	2.8
2005	147	85,928	2.9	46	25,809	2.7
2006	138	84,405	2.8	37	24,146	2.5
2007	163	82,424	2.7	28	22,996	2.4
2008	157	80,973	2.7	30	21,811	2.2
2009	166	79,644	2.7	37	20,657	2.1
2010	194	78,616	2.6	24	19,776	2.0

**Table A-1: (Continued)**

Variables	Fourth and fifth births			Fifth-birth model		
	Fourth-birth model			Fifth-birth model		
Calendar year	Live births	Woman-months count	%	Live births	Woman-months count	%
2011	150	77,923	26.0	27	19,187	2.0
2012	151	76,543	2.5	34	18,451	1.9
2013	142	75,711	2.5	26	17,740	1.8
<b>Fourth births, no. of subjects: 9,216</b>				<b>Fifth births, no. of subjects: 9,605</b>		
<b>No. of live births: 7,302</b>				<b>No. of live births: 1,723</b>		
<b>No. of woman-months: 3,003,142</b>				<b>No. of woman-months: 972,044</b>		

Source: Icelandic register data, author's calculations.

Notes: \* Third-birth age grouping. ++There are marginal differences between the analyses of third-birth and gender preferences due to 31 missing values in the variable 'Sex composition' (first or second).

## Results from models presented in paper

**Table A-2: Results from main effects models**

Relative risk of giving birth to a first child by calendar year standardized for age of mother. Rates relative to 1986			Relative risk of giving birth to a first child by calendar year standardized for age of mother. Rates relative to 1986. Ages 15–26		
Calendar year	Haz. ratio	P>z	Calendar year	Haz. ratio	P>z
1982	1.91	0.000	1982	2.31	0.000
1983	1.74	0.000	1983	2.11	0.000
1984	1.56	0.000	1984	1.87	0.000
1985	1.41	0.000	1985	1.67	0.000
1986	1.36	0.000	1986	1.62	0.000
1987	1.31	0.000	1987	1.60	0.000
1988	1.42	0.000	1988	1.67	0.000
1989	1.35	0.000	1989	1.58	0.000
1990	1.39	0.000	1990	1.62	0.000
1991	1.30	0.000	1991	1.48	0.000
1992	1.29	0.000	1992	1.48	0.000
1993	1.24	0.000	1993	1.38	0.000
1994	1.24	0.000	1994	1.34	0.000
1995	1.20	0.000	1995	1.31	0.000
1996	1.14	0.001	1996	1.25	0.000
1997	1.18	0.000	1997	1.31	0.000
1998	1.16	0.000	1998	1.23	0.000
1999	1.09	0.023	1999	1.20	0.000
2000	1.18	0.000	2000	1.22	0.000

**Table A-2: (Continued)**

<b>Relative risk of giving birth to a first child by calendar year standardized for age of mother. Rates relative to 1986</b>			<b>Relative risk of giving birth to a first child by calendar year standardized for age of mother. Rates relative to 1986. Ages 15–26</b>		
<b>Calendar year</b>	<b>Haz. ratio</b>	<b>P&gt;z</b>	<b>Calendar year</b>	<b>Haz. ratio</b>	<b>P&gt;z</b>
2001	1.11	0.008	2001	1.15	0.003
2002	1.00	0.903	2002	0.98	0.760
2003	1.04	0.361	2003	1.04	0.370
2004	1.00	–	2004	1.00	–
2005	1.01	0.896	2005	1.00	0.000
2006	1.03	0.489	2006	1.00	0.000
2007	1.00	0.960	2007	0.95	0.000
2008	1.00	0.903	2008	0.97	0.000
2009	0.99	0.796	2009	0.95	0.000
2010	0.96	0.268	2010	0.92	0.000
2011	0.84	0.000	2011	0.77	0.000
2012	0.87	0.001	2012	0.81	0.000
2013	0.82	0.000	2013	0.76	0.000
<b>Age</b>			<b>Age</b>		
15	0.03	0.000	15	0.03	0.000
16	0.11	0.000	16	0.10	0.000
17	0.28	0.000	17	0.28	0.000
18	0.52	0.000	18	0.52	0.000
19	0.74	0.000	19	0.73	0.000
20	1.00	–	20	1.00	–
21	1.22	0.000	21	1.22	0.000
22	1.37	0.000	22	1.38	0.000
23	1.58	0.000	23	1.59	0.000
24	1.77	0.000	24	1.78	0.000
25	2.01	0.000	25	2.03	0.000
26	2.15	0.000	26	2.33	0.000
27	2.21	0.000			
28	2.26	0.000			
29	2.01	0.000			
30	1.98	0.000			
31	1.79	0.000			
32	1.66	0.000			
33	1.46	0.000			
34	1.25	0.000			
35	1.10	0.076			
36	1.02	0.764			
37	0.73	0.000			
38	0.66	0.000			
39	0.62	0.000			

**Table A-2: (Continued)**

Relative risk of giving birth to a first child by calendar year standardized for age of mother. Rates relative to 1986			Relative risk of giving birth to a first child by calendar year standardized for age of mother. Rates relative to 1986. Ages 15–26		
Age	Haz. ratio	P>z	Age	Haz. ratio	P>z
40	0.37	0.000			
41	0.32	0.000			
42	0.19	0.000			
43	0.10	0.000			
44	0.07	0.000			
45	0.11	0.000			

Relative risk of giving birth to a first child by calendar year standardized for age of mother. Rates relative to 1986. Ages 29–45			Relative risk of having a second child by calendar year standardized for age of mother and age of firstborn. Rates relative to 1986		
Calendar year	Haz. ratio	P>z	Calendar year	Haz. ratio	P>z
1982	0.78	0.038	1982	1.16	0.000
1983	0.67	0.001	1983	1.18	0.000
1984	0.70	0.003	1984	1.08	0.079
1985	0.80	0.052	1985	1.02	0.722
1986	0.64	0.000	1986	1	–
1987	0.60	0.000	1987	1.09	0.040
1988	0.82	0.063	1988	1.22	0.000
1989	0.77	0.012	1989	1.17	0.000
1990	0.82	0.053	1990	1.21	0.000
1991	0.80	0.029	1991	1.11	0.011
1992	0.89	0.227	1992	1.12	0.009
1993	0.88	0.185	1993	1.24	0.000
1994	1.02	0.845	1994	1.13	0.002
1995	1.00	0.986	1995	1.02	0.665
1996	0.92	0.389	1996	1.16	0.000
1997	0.98	0.848	1997	1.06	0.149
1998	1.04	0.667	1998	1.14	0.001
1999	0.95	0.565	1999	1.11	0.017
2000	1.08	0.390	2000	1.11	0.017
2001	1.05	0.553	2001	1.09	0.043
2002	1.03	0.768	2002	1.11	0.013
2003	1.01	0.927	2003	1.11	0.011
2004	1.00	–	2004	1.22	0.000
2005	1.10	0.282	2005	1.22	0.000
2006	1.13	0.160	2006	1.24	0.000
2007	1.16	0.085	2007	1.26	0.000
2008	1.16	0.078	2008	1.34	0.000
2009	1.06	0.532	2009	1.38	0.000



Table A-2: (Continued)

Relative risk of giving birth to a first child by calendar year standardized for age of mother. Rates relative to 1986. Ages 29–45			Relative risk of having a second child by calendar year standardized for age of mother and age of firstborn. Rates relative to 1986		
Calendar year	Haz. ratio	P>z	Calendar year	Haz. ratio	P>z
2010	1.09	0.300	2010	1.38	0.000
2011	1.05	0.579	2011	1.29	0.000
2012	1.06	0.494	2012	1.28	0.000
2013	1.04	0.621	2013	1.21	0.000
<b>Age</b>			<b>Age group</b>		
29	1.80	0.000	<=18	0.47	0.000
30	1.78	0.000	19–21	0.64	0.000
31	1.61	0.000	22–24	0.77	0.000
32	1.50	0.000	25–26	0.86	0.000
33	1.32	0.000	27–28	1.00	–
34	1.13	0.097	29–30	1.08	0.000
35	1.00	–	31–32	1.10	0.000
36	0.92	0.324	33–34	0.99	0.691
37	0.67	0.000	35–36	0.76	0.000
38	0.60	0.000	37–38	0.58	0.000
39	0.57	0.000	39–40	0.32	0.000
40	0.34	0.000	41–45	0.11	0.000
41	0.30	0.000	<b>Duration from last birth</b>		
42	0.17	0.000	1–12 months	0.04	0.000
43	0.09	0.000	13–24 months	1.00	–
44	0.06	0.000	25–36 months	1.69	0.000
45	0.10	0.000	37–48 months	1.89	0.000
			49–60 months	1.83	0.000
			61–72 months	1.76	0.000
			73–96 months	1.58	0.000
			97–120 months	1.14	0.000
			121 months or more	0.73	0.000
<b>Relative risk of having a third child by calendar year standardized for age of mother and age of secondborn. Rates relative to 1986</b>			<b>Relative risk of having a second, third, fourth, and fifth child by calendar year standardized for age of mother and age of youngest born. Rates relative to second-birth rates in 1986</b>		
Calendar year	Haz. ratio	P>z	Calendar year	Haz. ratio	P>z
1982	1.13	0.018	1982	1.14	0.000
1983	1.17	0.003	1983	1.16	0.000
1984	1.10	0.061	1984	1.08	0.012
1985	0.98	0.706	1985	0.99	0.833
1986	1.00	–	1986	1.00	–
1987	1.12	0.031	1987	1.11	0.001

**Table A-2: (Continued)**

Relative risk of having a third child by calendar year standardized for age of mother and age of secondborn. Rates relative to 1986			Relative risk of having a second, third, fourth, and fifth child by calendar year standardized for age of mother and age of youngest born. Rates relative to second-birth rates in 1986		
Calendar year	Haz. ratio	P>z	Calendar year	Haz. ratio	P>z
1988	1.29	0.000	1988	1.24	0.000
1989	1.29	0.000	1989	1.21	0.000
1990	1.41	0.000	1990	1.29	0.000
1991	1.33	0.000	1991	1.21	0.000
1992	1.30	0.000	1992	1.21	0.000
1993	1.33	0.000	1993	1.28	0.000
1994	1.30	0.000	1994	1.19	0.000
1995	1.33	0.000	1995	1.13	0.000
1996	1.25	0.000	1996	1.19	0.000
1997	1.13	0.025	1997	1.09	0.005
1998	1.12	0.034	1998	1.12	0.000
1999	1.15	0.011	1999	1.10	0.002
2000	1.13	0.028	2000	1.10	0.002
2001	1.11	0.061	2001	1.09	0.009
2002	1.02	0.787	2002	1.07	0.027
2003	1.08	0.153	2003	1.08	0.012
2004	1.17	0.004	2004	1.18	0.000
2005	1.19	0.002	2005	1.19	0.000
2006	1.19	0.002	2006	1.19	0.000
2007	1.23	0.000	2007	1.23	0.000
2008	1.37	0.000	2008	1.32	0.000
2009	1.44	0.000	2009	1.38	0.000
2010	1.60	0.000	2010	1.44	0.000
2011	1.45	0.000	2011	1.31	0.000
2012	1.41	0.000	2012	1.29	0.000
2013	1.30	0.000	2013	1.21	0.000

Relative risk of having a third child by calendar year standardized for age of mother and age of secondborn. Rates relative to 1986			Relative risk of having a second, third, fourth, and fifth child by calendar year standardized for age of mother and age of youngest born. Rates relative to second-birth rates in 1986		
Age group	Haz. ratio	P>z	Age group	Haz. ratio	P>z
<=20	1.12	0.485	<=18	0.44	0.000
21–23	1.07	0.121	19–21	0.66	0.000
24–26	1.01	0.597	22–24	0.81	0.000
27–28	1.00	–	25–26	0.89	0.000
29–30	0.97	0.212	27–28	1.00	–
31–32	0.94	0.012	29–30	1.02	0.227

**Table A-2: (Continued)**

Relative risk of having a third child by calendar year standardized for age of mother and age of second-born. Rates relative to 1986			Relative risk of having a second, third, fourth, and fifth child by calendar year standardized for age of mother and age of youngest born. Rates relative to second-birth rates in 1986		
Age group	Haz. ratio	P>z	Age group	Haz. ratio	P>z
33–34	0.85	0.000	31–32	0.99	0.716
35–36	0.71	0.000	33–34	0.88	0.000
37–38	0.44	0.000	35–36	0.71	0.000
39–40	0.25	0.000	37–38	0.47	0.000
41–42	0.10	0.000	39–40	0.27	0.000
43–45	0.03	0.000	41–45	0.08	0.000
<b>Duration from last birth</b>			<b>Duration from last birth</b>		
1–12 months	0.04	0.000	1–12 months	0.05	0.000
13–24 months	1.00	–	13–24 months	1.00	–
25–36 months	1.26	0.000	25–36 months	1.44	0.000
37–48 months	1.35	0.000	37–48 months	1.54	0.000
49–60 months	1.57	0.000	49–60 months	1.58	0.000
61–72 months	1.71	0.000	61–72 months	1.60	0.000
73–96 months	1.71	0.000	73–96 months	1.53	0.000
97–120 months	1.43	0.000	97–120 months	1.21	0.000
121 months or more	1.00	0.887	121 months or more	0.83	0.000
<b>Sex composition*</b>			<b>Birth order</b>		
Boy and girl	1.00	–	Second	1.00	–
Two boys	1.14	0.000	Third	0.57	0.000
Two girls	1.10	0.000	Fourth	0.32	0.000
			Fifth	0.29	0.000

Source: Icelandic register data, author's calculations.

Notes: \* The variable 'Sex composition' is not included in the main third-birth model but rather in a later model for appropriate purposes.

