



**Demographic Research** a free, expedited, online journal  
of peer-reviewed research and commentary  
in the population sciences published by the  
Max Planck Institute for Demographic Research  
Konrad-Zuse Str. 1, D-18057 Rostock · GERMANY  
[www.demographic-research.org](http://www.demographic-research.org)

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## ***DEMOGRAPHIC RESEARCH***

**VOLUME 24, ARTICLE 34, PAGES 831-854  
PUBLISHED 28 JUNE 2011**

<http://www.demographic-research.org/Volumes/Vol24/34/>  
DOI: 10.4054/DemRes.2011.24.34

*Research Article*

### **A dynamic extension of the period life table**

**Frank T. Denton**

**Byron G. Spencer**

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## A dynamic extension of the period life table

**Frank T. Denton<sup>1</sup>**

**Byron G. Spencer<sup>2</sup>**

### Abstract

The standard period life table is based entirely on the death probabilities of the given period. Popular (not expert) usage of life expectancies from a period table typically ignores the fact that the expectancies make no allowance for future declines in mortality rates. But the historical record provides overwhelming evidence to suggest that declines will continue, and the period expectancies can therefore be misleading, in a practical context. We propose a “dynamic” extension of the period table that draws out the implications, for survivorship and life expectancy, of observed *rates of change* of death probabilities.

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<sup>1</sup> Frank T. Denton, M.A., LL.D. Research Institute for Quantitative Studies in Economics and Population and Department of Economics. McMaster University. Hamilton, Ontario L8S 4M4.  
Telephone: 905-525-9140-24595. Fax: 905-521-8232. E-mail: dentonf@mcmaster.ca.

<sup>2</sup> Byron G. Spencer, PhD. Research Institute for Quantitative Studies in Economics and Population and Department of Economics. McMaster University. Hamilton, Ontario L8S 4M4.  
Telephone: 905-525-9140-24594. Fax: 905-521-8232. E-mail: spencer@mcmaster.ca.

## 1. Introduction

The life table is a time-honoured and useful device for studying the implications of observed mortality rates in a population (Standard references include Chiang 1984; Keyfitz 1968; Keyfitz and Caswell 2005; Kintner 2004; among others.) The *period* life table, by far the most common form, is based on the rates for a particular year, or averages over a few consecutive years. What the table does is to provide the probabilities of dying and their implications for survivorship, including the expected years of life remaining at different ages. What it does *not* do, and is not intended to do, is to allow for changes in mortality probabilities. All calculations are based on the probabilities of the given period.

The mean expectations of remaining years of life – life expectancies, as they are commonly termed – are frequently cited and used in practical contexts, and only seldom (in popular usage) is it pointed out that they imply no further reduction of mortality. But the history of mortality rates asserts a likelihood to the contrary. The historical record in developed countries reveals continuous declines for as far back as modern statistical record-keeping allows.<sup>3</sup> We propose an extension of the period table that would allow explicitly for the possibility of further declines. We refer to it as a “dynamic” extension, and the life expectancies associated with it as “dynamic life expectancies.” The kernel idea is captured by comparing the following statements:

The period life table draws out the implications for survivorship and life expectancy of the observed age-specific mortality probabilities of a given period, under the assumption that the probabilities remain constant.

The *dynamic extension* of the period life table draws out the implications for survivorship and life expectancy of the observed mortality probabilities of a given period *and the observed rates of change of those probabilities*, under the assumption that the *rates of change* remain constant.

The extension is not proposed as a *replacement* for the period life table, nor should it be interpreted as providing an explicit forecast. It is intended rather as an optional

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<sup>3</sup> Oeppen and Vaupel (2002) examine the world historical time path of female life expectancy at birth since 1840 based on leading-country life expectancies in each period. They show the path to be virtually linear over 160 years, with an average increment of a quarter of a year per year and no sign of slowing. For two other examples see the study of mortality rates in the G7 countries by Tuljapurkar, Li, and Boe (2000) and the study of the Canadian record in Denton, Feaver, and Spencer (2005); the downward paths of mortality based on logarithmic summary measures in these two studies are virtually linear too, again with no sign of moderation.

supplement to the standard table. Calculations of dynamic life expectancies are "what if" calculations: "what if" observed rates of change were to continue? In that sense they can be compared with the "what if" implications of the standard period life expectancies: "what if" mortality rates were to remain unchanged?

We note briefly the mathematics of the standard period life table and then develop the dynamic framework. We offer some general observations on the interpretation of the framework in the context of previous literature on the modeling of mortality rates and provide a more specific comparison with earlier work on age-independent changes in mortality. We show how the framework can be represented by an augmented Lexis diagram and illustrate dynamic life expectancy by an application to Canadian data, including comparisons of alternative dynamic expectancies with published 2001 period expectancies. We emphasize that the calculated death rates underlying the dynamic expectancies should be interpreted as "what if" calculations, not as forecasts.

## 2. The period life table

Consider a cohort of newborn children observed, as it grows older, at exact ages  $x = 0, 1, \dots, m$ , where  $m$  is the oldest age at which there are any survivors. Let  $l_0$  be the initial size of the cohort (an arbitrary number) and let  $l_x$  be the number surviving to age  $x$ . The probability that a survivor to age  $x$  will die in the interval  $x$  to  $x+1$  is constant at  $q_x$  and an entry in the sequence  $l_1, l_2, \dots, l_m$  can be calculated recursively from  $l_{x+1} = (1 - q_x)l_x$ , or directly from  $l_{x+1} = l_0 \prod_{t=0}^x (1 - q_t)$ .

Let  $L_x$  be the number of person-years lived in the interval  $x$  to  $x+1$  and write  $L_x = f(l_x, l_{x+1})$ , where  $f$  denotes an interpolation function, the form of which depends on the pattern of deaths in the interval. If deaths occur uniformly then  $f$  is linear and  $L_x = (l_x + l_{x+1})/2$ . (The assumption of uniformity is often considered a reliable approximation for all but the youngest ages.) The total number of person-years yet to be lived at age  $x$  is then  $T_x = \sum_{t=x}^m L_t$ , and life expectancy, the mean number of years remaining at age  $x$ , is  $e_x = T_x/l_x$ .

### 3. Dynamic extension

Let  $x$  stand again for exact age (with bounds 0 and  $m$ ) but refer to it now as *initial age*. Let  $y$  stand also for age and refer to it as *subsequent age*, with  $y \geq x$ . Assume there are period life tables for two periods (years)  $n$  years apart, and refer to the more recent one as the *reference period*. Let  $q_x$  be the death probability at age  $x$  in the reference period table and let  $\bar{q}_x$  be the corresponding probability in the earlier period table. The annual rate of change for any age  $x$  is then  $r_x = (q_x/\bar{q}_x)^{1/n} - 1$ .<sup>4</sup> Now let  $l_{xy}$  denote the population of initial age  $x$  that survives to age  $y$ . (Note that the initial population in the reference period is then  $l_{xx}$ , the first entry in the cohort sequence  $l_{xx}, l_{x,x+1}, \dots, l_{xy}$ . We refer to the cohort as the  $l_{xx}$  cohort.) Assuming constancy of the death probability rate of change at each age, the probability that a member of the  $l_{xx}$  cohort who has survived to age  $y$  will die in the interval  $y$  to  $y+1$  is then  $q_{xy} = q_y (1 + r_y)^{y-x}$ , where  $q_y$  is the death probability at age  $y$  in the reference period,  $r_y$  is the annual rate of change of that probability, and  $y-x$  is the number of years elapsed in time (as well as age) since the cohort was of age  $x$ , and hence the number of years over which the age  $y$  probability has changed. The cohort sequence of  $l_{xy}$  values ( $y \geq x$ ) can be calculated recursively for each cohort from  $l_{x,y+1} = l_{xy} (1 - q_{xy})$ , or directly from  $l_{x,y+1} = l_{xx} \prod_{t=x}^y (1 - q_{xt})$ .

Now let  $L_{xy}$  (the dynamic analogue of  $L_x$ ) stand for the number of person-years lived by survivors of the  $l_{xx}$  cohort in the interval  $y$  to  $y+1$ . Adapting the notation used earlier,  $L_{xy} = f(l_{xy}, l_{x,y+1})$ . As before, if the pattern of deaths is uniform in the interval, then the interpolation function  $f$  is linear, and  $L_{xy} = (l_{xy} + l_{x,y+1})/2$ . The total

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<sup>4</sup> This is our preferred representation of the rate of change of  $q_x$  for present purposes. An alternative would be the average annual change in  $\ln q_x$  or in logit  $q_x$ . (The logit form has merit in a more general context because it forces a rate to lie in the interval (0, 1). However given the characteristics of death rates that is not an issue here.) A linear representation (the arithmetic change in  $q_x$ ) would run the risk of generating negative values when extrapolated.

number of person-years yet to be lived by the  $l_{xx}$  cohort is  $T_{xx} = \sum_{t=x}^m L_{xt}$  and the dynamic life expectancy is  $e_{xx} = T_{xx}/l_{xx}$ .

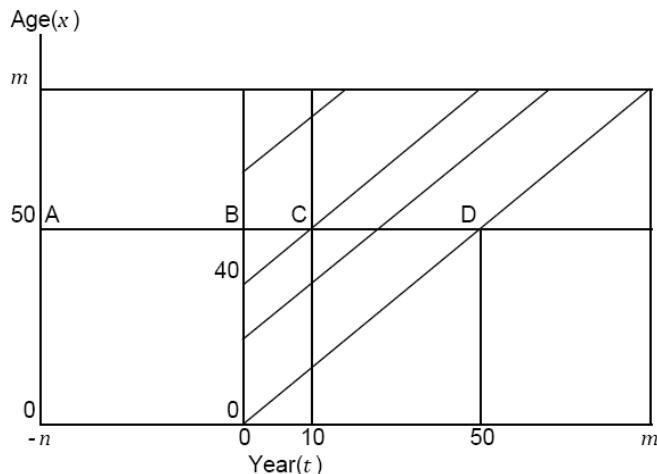
There are thus two innovations on which the dynamic extension is based. The first is that each  $l_x$  entry in the period table is redefined as  $l_{xx}$ , the initial value of a separate cohort, with its own age sequence of death probabilities. The second is that the no-change assumption about death probabilities in the period table is replaced with the “what if” assumption that the observed rates of change since a previous period table will continue for as long as there are any survivors of a cohort.

#### 4. Diagrammatic depiction

The framework for the dynamic extension can be represented in the form of an augmented Lexis diagram, as in Figure 1. Age is measured in the vertical dimension, time in the horizontal dimension. For convenience we choose 0 as the date (year) of the reference life table, and thus  $-n$  as the date of the earlier table. With the scaling of age ( $x$ ) and time ( $t$ ) equal, the upward-sloping 45 degree lines represent the remaining age/time life paths of the  $m+1$  cohorts ( $l_{xx}$ ) defined by the reference table and the horizontal lines represent the time paths of death probabilities ( $q_x$ ). For example the line AB represents the  $n$ -year interval over which the annual rate of change of  $q_{50}$  is calculated. BC, the extension of AB, then represents the 10-year interval over which that annual rate of change is applied (by compounding over the ten years) to arrive at the  $q_{50}$  value for the survivors of the cohort of initial age 40 when (10 years later) that cohort is of age 50. (The application of the death probability to the cohort survivors occurs at the intersection point C. In the “dynamic” notation, the number of cohort survivors at that point is  $l_{xy} = l_{40,50}$  and the death probability is  $q_{xy} = q_{40,50}$ .) Similarly the line BD represents the interval over which the same annual rate of change is applied (by compounding over 50 years) to arrive at the  $q_{50}$  value for the survivors of the cohort of initial age 0 when (50 years later) that cohort is of age 50. Note that if  $n=0$  (so that the first two vertical lines coincide) the annual rate of change calculation is nullified and the diagram reduces to a simple depiction of the standard period life table: the set of death probabilities in the reference table then remains constant through time for each cohort and the probabilities lie on the vertical line passing through the point B. Note further, for another comparison, that if the rates of change of death probabilities over the  $n$ -year interval between the original life tables are simply used as an extrapolative device to move the time 0 probabilities forward another 10 years, say,

then a new life table can be generated at time 10. The new death probabilities would then lie on the vertical line passing through the point C in the diagram, rather than on the cohort-specific diagonal lines. The difference between the dynamic calculations and this simple extrapolative procedure is that the former assumes the rates of change will continue to compound, for each age cohort, for as long as there are any survivors of that cohort, while the latter makes no such assumption – it treats the changes as representing a one-time shift in the schedule of death probabilities. In all of this it is important to keep in mind one basic idea: the dynamic calculations are intended to provide an answer to the simple “what if” question: what if observed rates of change in death probabilities were to be maintained?

**Figure 1: Augmented Lexis diagram**



## 5. A note on interpretation

There is a long history in demography of attempts to model and project age-specific mortality rates. The history includes (among much else) the “law of mortality” model proposed by Gompertz (1825), the modification of that model by Makeham (1860), various adaptations of the logistic model see Bongaarts (2005), and the stochastic modeling and projection method developed by Lee and Carter (1992). (See also Keyfitz (1982) for an interesting discussion and perspective on the modeling of mortality patterns.) One theme that runs through much of the literature on parametric modeling is

the distinction between “senescent mortality” (associated with aging) and “background mortality” (independent of age); see Bongaarts (2009) for discussion. (In the Gompertz or similar sigmoid-type representation of mortality, with discrete age,  $\ln q_x = \ln \alpha + f(x)$ , where  $\alpha$  is the background or baseline component and  $f(x)$  is the senescence component; historically, continuing declines in  $\alpha$  have been associated with consistently rising human life expectancy.) Model life table systems take advantage of the regularities in mortality patterns by reducing their representation to a small number of parameters that make it possible to generate complete tables from limited empirical information.<sup>5</sup> All of this work is interesting in itself but different in nature and purpose from the contribution of the present paper. We mention it simply for comparison with our proposed extension of the standard life table and the drawing out of implications of observed changes in mortality rates for life expectancy, were they to be maintained.

We note also the following: Our procedure applies to cohorts and their death probabilities. In a population projection context one might think in terms of cohorts also but there one would need to specify death rates for all cohorts, including those born throughout the projection period. In our case the concern is with only those cohorts defined (implicitly) by the reference life table, and with death probabilities only for as long as there are any survivors of those cohorts. The death probabilities of as yet unborn cohorts are irrelevant.

## 6. Age-independent changes in mortality

The effects of age-independent changes in mortality rates have been studied in the context of the stable population model by Coale (1963), Coale and Demeny (1967), Keyfitz (1977), and Keyfitz and Caswell (2005). Two assumptions are considered by Keyfitz and Caswell: (a) uniform *arithmetic* changes in rates and (b) uniform *proportional* changes. The second assumption is the more realistic of the two (as the authors point out) and we focus on it. The Keyfitz and Caswell analysis is in a continuous time framework but we adapt it to the discrete time framework that we use here, for comparison with our proposed “dynamic” calculations.

Consider then an age schedule of population mortality rates, or more consistently for our purposes, death probabilities,  $q_x$  ( $x = 0, 1, \dots, m$ ). Now (following Keyfitz and Caswell) suppose that the death probability at each age changes by the same fraction  $\delta$

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<sup>5</sup> Well known applications of this idea are the model systems of Coale and Demeny (1966), Ledermann (1969), Brass (1971), and United Nations (1982). See Suchindran (2004) and Murray et al. (2001) for further references and discussion.

each year, yielding after  $k$  years a new schedule  $q_x^* = (1+k\delta)q_x$  ( $x=0,1,\dots,m$ ). ( $k\delta$  must exceed  $-1$  for the calculation to make sense.) Assuming the same initial input  $l_0$  in both cases, the corresponding number of survivors at any age  $x+1$  is then  $l_{x+1}^* = l_0 \prod_{t=0}^x (1-q_t^*) = l_0 \prod_{t=0}^x (1-(1+k\delta)q_x)$ , which can be expressed as a polynomial in  $\delta$  of degree  $x+1$ . Also adapting our earlier notation, we can write  $L_x^* = f(l_x^*, l_{x+1}^*)$  and  $e_x^* = T_x^*/l_x^*$ . If  $f$  denotes a linear function it is easily shown that both  $e_x^*$  and  $e_x^* - e_x$  can be expressed as polynomials in  $\delta$  of degree  $m$  for every  $x$  (assuming the terminal probability  $q_m$  is fixed at 1). A new period life table can be derived from the  $q_x^*$  probability schedule, representing a new stationary population.

Now consider for comparison, the assumption of age-independent rates of change in death probabilities in the context of our dynamic extension of the period life table. Using our earlier notation, that assumption implies the special case  $r_x = r$  for all  $x$ , so that  $q_{xy} = q_y (1+r)^{y-x}$ . (Note that for  $r$  and  $\delta$  to have similar effects over the interval  $y-x$  requires  $(y-x)\delta = (1-r)^{y-x} - 1$ , or  $\delta = ((1+r)^{y-x} - 1)/(y-x)$ .) Unlike the model used by Keyfitz and Caswell, the rate of change is thus compounded, and compounded separately for each of the  $m$  cohorts implicit in the life table (excluding the age  $m$  cohort with  $q_m$  fixed at 1). Carrying through the substitution to the calculation of  $l_{x,y+1}$  yields a polynomial in  $r$  of degree  $y-x$ , and carrying the substitution further, to  $L_{xy}$  and  $T_{xx}$ , leads to expressions for  $e_{xx}$  and  $e_{xx} - e_x$  that can be interpreted as polynomials in  $r$  of degree  $m-x$ . (The polynomials are of degree  $m-x$  rather than  $m$  since, unlike the period life table,  $l_{xx}$  is treated as given at each age  $x$  and only the death probabilities for the remaining  $m-x$  years are functions of  $r$ .) To sum up the comparison, the Keyfitz and Caswell analysis of a uniform proportional change in death probabilities converts one stationary population (and associated life table) into another. Our proposed “dynamic” extension is different (by its nature); it does not yield a stationary population, even if the rate of change is uniform and proportional. We offer the comparison to help in understanding where the differences lie.

## 7. Application

Examples are provided in Tables 1 and 2. The period life expectancies at five-year age intervals from 0 to 100 are taken from the Statistics Canada life tables for 2001 (based on average data for the three-year period 2000-2002) and compared with three dynamic life expectancies. Comparisons are shown for males in Table 1, females in Table 2. Complete single-age comparisons are provided in the Appendix; they are shown in graphical form as differences between dynamic and period life expectancies, in Figure 2. The dynamic expectancies are calculated using average rates of change of  $q_x$  values over the previous 10, 25, and 50 years.<sup>6</sup> The younger the age, the longer the period over which changes in death probabilities take place. The oldest age at which there are any survivors in the calculations exceeds 100: thus at age 0 the dynamic expectancies incorporate over a century of changes, at age 50 over half a century, and at age 75 over a quarter of a century.

**Table 1: Comparison of period life expectancies with alternative dynamic life expectancies at 5-year age intervals: Canada, 2001, males**

Age <i>x</i>	Period life expectancy <i>e<sub>x</sub></i>	Dynamic life expectancy ( $e_{xx}$ ) based on $q_x$ rates of change in the previous --					
		10 years		25 years		50 years	
		$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$
0	76.9	85.9	9.0	87.3	10.4	84.1	7.2
5	72.4	81.0	8.5	82.2	9.8	79.1	6.7
10	67.5	75.5	8.0	76.6	9.1	73.7	6.2
15	62.5	69.9	7.4	70.9	8.4	68.2	5.6
20	57.7	64.5	6.8	65.4	7.6	62.8	5.1
25	53.0	59.1	6.2	59.9	6.9	57.5	4.6
30	48.2	53.7	5.5	54.3	6.2	52.2	4.1
35	43.4	48.3	4.9	48.8	5.4	46.9	3.5
40	38.6	42.9	4.2	43.3	4.6	41.6	3.0
45	34.0	37.5	3.6	37.8	3.8	36.5	2.5
50	29.4	32.3	2.9	32.5	3.1	31.4	2.0

<sup>6</sup> The dynamic calculations are as described in a previous section. To allow for possible nonuniformity in the distribution of deaths within a one-year age interval we did the following. For each period table we first calculated  $L_{xy}$  by simple averaging  $L_{xy} = (l_{xy} + l_{x,y+1})/2$  and then multiplied it by a correction factor equal to the ratio of the actual value of  $L_x$  in the period table (which may have incorporated nonlinear interpolation) to the value that would have been shown in that table had simple averaging been used. We then assumed, for each pair of tables used in the calculations (1991 and 2001, 1976 and 2001, 1951 and 2001), that the correction factor would continue to change at the same proportionate annual rate as it had over the interval between the tables. The correction factors were all less than or equal to 1, and different from 1 only for ages 0 to 4. The most extreme correction factor was 0.98267, for age 0 in the 1951 life table.

**Table 1:** (Continued)

Age	Period life expectancy	Dynamic life expectancy ( $e_{xx}$ ) based on $q_x$ rates of change in the previous --					
		10 years	25 years	50 years	$e_{xx} - e_x$	$e_{xx} - e_x$	$e_{xx} - e_x$
$x$	$e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$
55	25.0	27.2	2.2	27.4	2.4	26.6	1.5
60	20.8	22.4	1.6	22.6	1.7	22.0	1.2
65	17.0	18.0	1.1	18.2	1.2	17.8	0.8
70	13.5	14.1	0.6	14.2	0.8	14.0	0.5
75	10.3	10.6	0.3	10.8	0.4	10.7	0.3
80	7.7	7.8	0.1	7.9	0.2	7.9	0.2
85	5.5	5.5	0.0	5.6	0.1	5.6	0.1
90	3.9	3.9	0.0	3.9	0.1	3.9	0.0
95	2.8	2.8	0.0	2.9	0.0	2.9	0.0
100	2.0	2.0	0.0	2.0	0.0	2.0	0.0

Notes: Period life expectancies are from life tables in Statistics Canada (2006). Dynamic life expectancies are based on rates of change of  $q_x$  values calculated from life tables in that publication and in Dominion Bureau of Statistics (1960), Statistics Canada (1979), and Statistics Canada (1995).

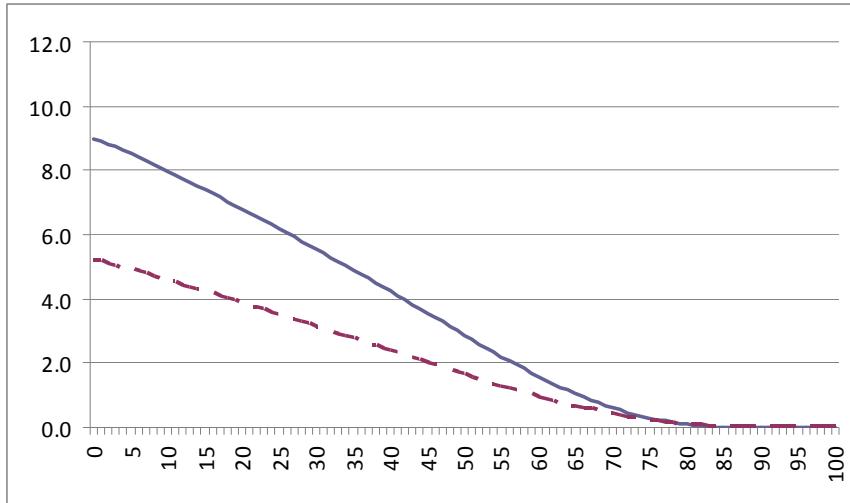
**Table 2:** Comparison of period life expectancies with alternative dynamic life expectancies at 5-year age intervals: Canada, 2001, females

Age	Period life expectancy	Dynamic life expectancy ( $e_{xx}$ ) based on $q_x$ rates of change in the previous --					
		10 years	25 years	50 years	$e_{xx} - e_x$	$e_{xx} - e_x$	$e_{xx} - e_x$
$x$	$e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$
0	82.0	87.3	5.2	91.8	9.8	91.9	9.8
5	77.5	82.4	4.9	86.7	9.2	86.7	9.3
10	72.5	77.1	4.6	81.0	8.5	81.1	8.6
15	67.6	71.8	4.3	75.4	7.9	75.5	8.0
20	62.7	66.6	3.9	69.8	7.2	70.0	7.3
25	57.8	61.3	3.5	64.3	6.5	64.4	6.7
30	52.9	56.0	3.2	58.7	5.8	58.8	6.0
35	48.0	50.8	2.8	53.1	5.2	53.3	5.3
40	43.1	45.5	2.4	47.6	4.5	47.8	4.6
45	38.4	40.4	2.0	42.2	3.8	42.3	4.0
50	33.7	35.3	1.7	36.9	3.2	37.0	3.3
55	29.1	30.4	1.3	31.7	2.6	31.8	2.7
60	24.7	25.7	1.0	26.8	2.0	26.9	2.2
65	20.5	21.2	0.7	22.1	1.5	22.1	1.6
70	16.6	17.0	0.4	17.7	1.1	17.7	1.2
75	12.9	13.2	0.3	13.7	0.7	13.7	0.8
80	9.7	9.8	0.1	10.1	0.5	10.1	0.5
85	7.0	7.0	0.1	7.2	0.3	7.2	0.3
90	4.9	5.0	0.0	5.1	0.2	5.1	0.1
95	3.5	3.5	0.0	3.6	0.1	3.5	0.1
100	2.4	2.5	0.1	2.5	0.1	2.4	0.0

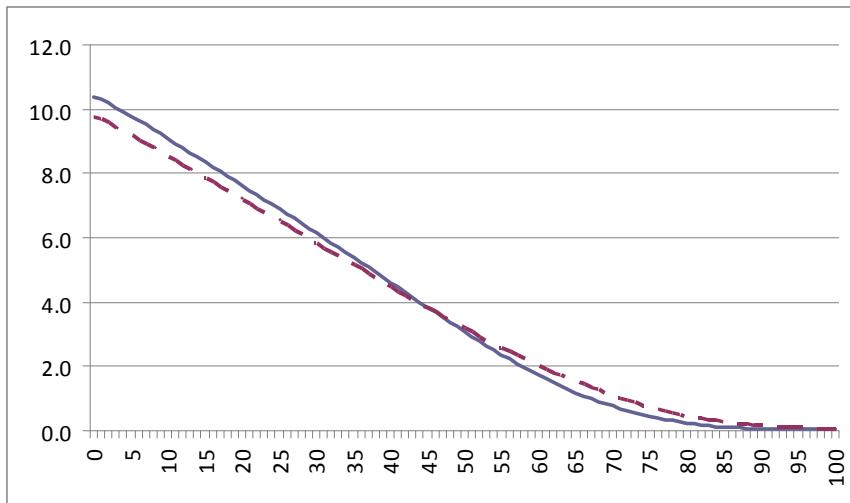
Note: See note to Table 1.

**Figure 2:** Number of years by which dynamic life expectancies exceed period life expectancies, Canada 2001, based on rates of change of death probabilities over the previous 10, 25, and 50 years

*Based on previous 10 years*

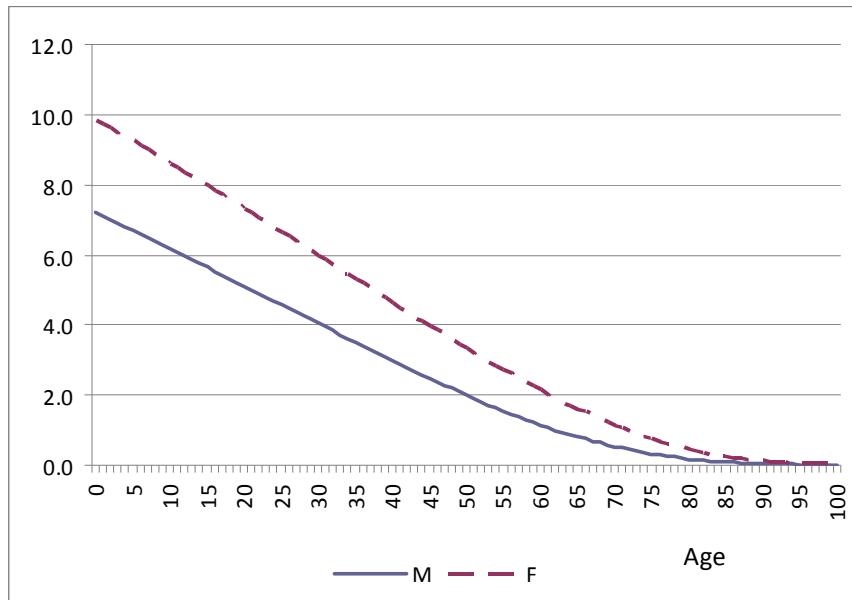


*Based on previous 25 years*



**Figure 2:** (Continued)

*Based on previous 50 years*



The differences between the period and dynamic expectancies vary with the choice of historical period but for children, young adults, and younger middle-aged adults the differences are substantial in all cases. At birth a newborn male child has a life expectancy of 76.9 years according to the period calculations, a female child an expectancy of 82.0 years. According to the dynamic calculations though, male life expectancy ranges from 84.1 to 87.3 years, female life expectancy from 87.3 to 91.9. By age 60 the differences have become relatively small – roughly a year or two for each sex.

The effects of the dynamic calculations are captured most prominently by the differences at birth ( $e_{00} - e_0$ ) based alternatively on rates of change of death probabilities over intervals of 10, 25, and 50 years. A 10-year interval (1991-2001) produces for males a difference of 9.0 years, a 25-year interval a difference of 10.4, and a 50-year interval a difference of 7.2. That the 50-year calculation is so much lower than the 25-year result is attributable to the fact that mortality declines were much slower between 1951 and 1976 than between 1976 and 2001. The published period life expectancy at birth for males was 70.2 in 1976. Had it been calculated using our

dynamic procedure based on the previous 25 years it would have been 73.3, a difference of 3.1 years. This contrasts with the difference of 10.4 years in 2001 (based again on the previous 25 years) but is consistent with the overall 7.2 year difference based on rates of change of death rates over the whole of the 50-year interval.

Female life expectancies at birth show another pattern in relation to the choice of interval. The 10-year interval produces an  $e_{00} - e_0$  difference of 5.2 years whereas the 25-year and 50-year intervals produce identical results, 9.8.

We might add at this point that the variation in differences between dynamic and period life expectancies due to the choice of historical interval may be good to have on display in presenting results. Our aim in this paper, as we emphasize, is not to suggest a method of forecasting mortality probabilities and their implications but rather to show how misleading it is to assume that a period life table provides life expectancies that can be regarded as realistic. It may be desirable in practice to show not only that dynamic expectancies are invariably greater than period expectancies but that there is uncertainty about the expectancies themselves, based on the historical record.

## 8. Life expectancy and mortality prediction

Any calculation of life expectancy for members of a cohort who are alive at the date of the calculation requires assumptions about future death probabilities. The assumptions may be viewed as forecasts – predictions of what is likely to happen (perhaps a range of possibilities) – or they may be artificial projections, for purposes of analysis. Life expectancies derived from a period life table assume, implicitly, that the age-specific probabilities in the table will remain constant. That is clearly unrealistic, given the history of mortality rates, and is not intended to be realistic by the constructor of the table, but it does serve to bring out the implications of period table probabilities. In the same way, the dynamic life expectancies that we propose serve to bring out the implications of maintaining the *rates of change* of the period table probabilities.<sup>7</sup> They too should not be interpreted as based on predictions. To view them as reflecting forecasts of actual probabilities would be in contradiction of how we wish them to be

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<sup>7</sup> The following apt analogy is from Keyfitz and Caswell (2005:63). "As Keyfitz (1972) pointed out, one of the most powerful ways to study present conditions is to examine their consequences were they to remain as they are. A speedometer works the same way. A reading of 60 miles per hour predicts that, in one hour, the car will be found 60 miles in a straight line from its present location. As a forecast, this is almost always false. But as a projection, it provides valuable information about the present situation of the automobile." The analogy applies to the standard period life table but it is applicable also to our dynamic extension. Dynamic life expectancies can also be interpreted as providing valuable information about the present situation, including now the implications of a continuation of the rates of change of death probabilities.

viewed. As we said in the introduction to this paper (and as the title implies) we regard our proposal as simply an extension of the standard period life table.<sup>8</sup>

To emphasize further the distinction, there are various approaches that we might take if forecasting death probabilities and the attendant life expectancies were our goal. The literature provides alternatives ranging from purely judgmental forecasts (single, "high, low, medium" options, or some similar variant), to forecasts based on a parameterization of the mortality age schedule, to stochastic forecasts based on probabilities inferred from historical time series. Parameterization is often based on some form of sigmoid-type representation of the age-survival pattern; an adaptation of the logistic function is a possibility, the Weibull function, or the Gompertz function, and there are others.<sup>9</sup> Stochastic forecasting was pioneered by Ronald Lee and his associates, and the Lee-Carter method is now well established (Lee and Carter 1992; Lee and Tuljapurkar 1994).<sup>10</sup> But again, we are not concerned here with forecasting, merely with showing the implications for life expectancies of the maintenance of observed rates of change of death probabilities.

## 9. Conclusion

The dynamic extension provides an informative supplement to standard life table calculations. Without purporting to represent predictions, dynamic life expectancies draw out the implications of observed historical rates of change of death probabilities. The use of period expectancies as if they represented the future (as happens implicitly in popular usage) implies no further change in the probabilities. The use of dynamic expectancies implies continuation of observed rates of change. At the least they convey

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<sup>8</sup> As Keilman (2008) observes in his study of the accuracy of European population forecasts, *projections*, strictly speaking, can never be wrong (aside from calculating errors): they are simply calculations of what the future population would be under a given set of assumptions, not what it is expected to be. Projections should thus be interpreted as different from forecasts. In that sense our dynamic life expectancies should be interpreted as lying in the domain of projections.

<sup>9</sup> The Gompertz function, for example, can be estimated from period death rates, examined, and projected, thus reducing some 100 age-specific mortality rates to a small set of parameters. (We note, parenthetically, that we have used the Gompertz function in this way to represent fertility schedules by a transformation of parameters that allows one to study and project, independently, total lifetime fertility, mean age of mothers at childbirth, and the interquartile range of ages at childbirth; see Denton and Spencer 1974.)

<sup>10</sup> Once the idea of stochastic forecasting is accepted as a way of capturing the uncertainty associated with changes in death probabilities there are various alternatives. The Lee-Carter method is one. Others include procedures involving stochastic (autoregressive models) and bootstrapping methods based on random samples drawn from historical series of changes in mortality rates, as in Denton, Feaver, and Spencer (2005).

some idea as to the possible degree of understatement that may exist in the use of period expectancies.

## **10. Acknowledgements**

We are grateful to the Social Sciences and Humanities Research Council of Canada for its support of the SEDAP (Social and Economic Dimensions of an Aging Population) Research Program through its Major Collaborative Research Initiative. We thank Christine Feaver who produced the tabular results and commented on the paper. We thank also the editor and two anonymous referees for helpful comments and suggestions.

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## Appendix

**Table A1: Comparison of period life expectancies with alternative dynamic life expectancies at single-year age intervals: Canada, 2001, males**

Age <i>x</i>	Period life expectancy $e_x$	Dynamic life expectancy ( $e_{xx}$ ) based on $q_x$ rates of change in the previous --					
		10 years		25 years		50 years	
		$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$
0	76.9	85.9	9.0	87.3	10.4	84.1	7.2
1	76.4	85.3	8.9	86.7	10.3	83.5	7.1
2	75.4	84.2	8.8	85.6	10.2	82.4	7.0
3	74.4	83.1	8.7	84.5	10.1	81.3	6.9
4	73.4	82.0	8.6	83.3	9.9	80.2	6.8
5	72.4	81.0	8.5	82.2	9.8	79.1	6.7
6	71.5	79.9	8.4	81.1	9.6	78.1	6.6
7	70.5	78.8	8.3	80.0	9.5	77.0	6.5
8	69.5	77.7	8.2	78.8	9.4	75.9	6.4
9	68.5	76.6	8.1	77.7	9.2	74.8	6.3
10	67.5	75.5	8.0	76.6	9.1	73.7	6.2
11	66.5	74.4	7.9	75.4	8.9	72.6	6.1
12	65.5	73.2	7.8	74.3	8.8	71.5	6.0
13	64.5	72.1	7.6	73.2	8.7	70.4	5.8
14	63.5	71.0	7.5	72.0	8.5	69.3	5.8
15	62.5	69.9	7.4	70.9	8.4	68.2	5.6
16	61.6	68.8	7.3	69.8	8.2	67.1	5.5
17	60.6	67.8	7.2	68.7	8.1	66.0	5.4
18	59.6	66.7	7.0	67.6	7.9	65.0	5.3
19	58.7	65.6	6.9	66.5	7.8	63.9	5.2
20	57.7	64.5	6.8	65.4	7.6	62.8	5.1
21	56.8	63.5	6.7	64.3	7.5	61.8	5.0
22	55.8	62.4	6.6	63.2	7.3	60.7	4.9
23	54.9	61.3	6.4	62.1	7.2	59.7	4.8
24	53.9	60.2	6.3	61.0	7.1	58.6	4.7
25	53.0	59.1	6.2	59.9	6.9	57.5	4.6
26	52.0	58.1	6.1	58.8	6.8	56.5	4.5
27	51.1	57.0	5.9	57.7	6.6	55.4	4.4
28	50.1	55.9	5.8	56.5	6.5	54.4	4.3
29	49.1	54.8	5.7	55.4	6.3	53.3	4.2
30	48.2	53.7	5.5	54.3	6.2	52.2	4.1
31	47.2	52.6	5.4	53.2	6.0	51.2	3.9
32	46.3	51.5	5.3	52.1	5.9	50.1	3.8

**Table A1:** (Continued)

Age <i>x</i>	Period life expectancy $e_x$	Dynamic life expectancy ( $e_{xx}$ ) based on $q_x$ rates of change in the previous --					
		10 years		25 years		50 years	
	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$	
33	45.3	50.5	5.2	51.0	5.7	49.0	3.7
34	44.3	49.4	5.0	49.9	5.6	48.0	3.6
35	43.4	48.3	4.9	48.8	5.4	46.9	3.5
36	42.4	47.2	4.8	47.7	5.2	45.9	3.4
37	41.5	46.1	4.6	46.6	5.1	44.8	3.3
38	40.5	45.0	4.5	45.5	4.9	43.7	3.2
39	39.6	44.0	4.4	44.4	4.8	42.7	3.1
40	38.6	42.9	4.2	43.3	4.6	41.6	3.0
41	37.7	41.8	4.1	42.2	4.5	40.6	2.9
42	36.8	40.7	4.0	41.1	4.3	39.6	2.8
43	35.8	39.7	3.8	40.0	4.2	38.5	2.7
44	34.9	38.6	3.7	38.9	4.0	37.5	2.6
45	34.0	37.5	3.6	37.8	3.8	36.5	2.5
46	33.1	36.5	3.4	36.7	3.7	35.4	2.4
47	32.1	35.4	3.3	35.7	3.5	34.4	2.3
48	31.2	34.4	3.2	34.6	3.4	33.4	2.2
49	30.3	33.3	3.0	33.5	3.2	32.4	2.1
50	29.4	32.3	2.9	32.5	3.1	31.4	2.0
51	28.5	31.3	2.8	31.5	2.9	30.4	1.9
52	27.6	30.2	2.6	30.4	2.8	29.4	1.8
53	26.7	29.2	2.5	29.4	2.7	28.5	1.7
54	25.9	28.2	2.3	28.4	2.5	27.5	1.6
55	25.0	27.2	2.2	27.4	2.4	26.6	1.5
56	24.2	26.2	2.1	26.4	2.2	25.6	1.5
57	23.3	25.3	2.0	25.4	2.1	24.7	1.4
58	22.5	24.3	1.8	24.5	2.0	23.8	1.3
59	21.7	23.4	1.7	23.5	1.9	22.9	1.2
60	20.8	22.4	1.6	22.6	1.7	22.0	1.2
61	20.0	21.5	1.5	21.7	1.6	21.1	1.1
62	19.3	20.6	1.4	20.8	1.5	20.3	1.0
63	18.5	19.7	1.3	19.9	1.4	19.4	0.9
64	17.7	18.9	1.2	19.0	1.3	18.6	0.9
65	17.0	18.0	1.1	18.2	1.2	17.8	0.8
66	16.3	17.2	0.9	17.3	1.1	17.0	0.8
67	15.5	16.4	0.9	16.5	1.0	16.2	0.7

**Table A1: (Continued)**

Age <i>x</i>	Period life expectancy $e_x$	Dynamic life expectancy ( $e_{xx}$ ) based on $q_x$ rates of change in the previous --					
		10 years		25 years		50 years	
		$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$
68	14.8	15.6	0.8	15.7	0.9	15.5	0.7
69	14.1	14.8	0.7	15.0	0.8	14.7	0.6
70	13.5	14.1	0.6	14.2	0.8	14.0	0.5
71	12.8	13.4	0.5	13.5	0.7	13.3	0.5
72	12.2	12.6	0.5	12.8	0.6	12.6	0.4
73	11.5	11.9	0.4	12.1	0.6	11.9	0.4
74	10.9	11.3	0.4	11.4	0.5	11.3	0.4
75	10.3	10.6	0.3	10.8	0.4	10.7	0.3
76	9.8	10.0	0.3	10.2	0.4	10.1	0.3
77	9.2	9.4	0.2	9.6	0.4	9.5	0.3
78	8.7	8.8	0.2	9.0	0.3	8.9	0.2
79	8.2	8.3	0.1	8.4	0.3	8.4	0.2
80	7.7	7.8	0.1	7.9	0.2	7.9	0.2
81	7.2	7.3	0.1	7.4	0.2	7.4	0.2
82	6.7	6.8	0.0	6.9	0.2	6.9	0.1
83	6.3	6.3	0.0	6.5	0.2	6.4	0.1
84	5.9	5.9	0.0	6.0	0.1	6.0	0.1
85	5.5	5.5	0.0	5.6	0.1	5.6	0.1
86	5.1	5.1	0.0	5.2	0.1	5.2	0.1
87	4.8	4.8	0.0	4.9	0.1	4.9	0.1
88	4.5	4.4	0.0	4.5	0.1	4.5	0.1
89	4.2	4.1	0.0	4.2	0.1	4.2	0.0
90	3.9	3.9	0.0	3.9	0.1	3.9	0.0
91	3.6	3.6	0.0	3.7	0.1	3.7	0.0
92	3.4	3.4	0.0	3.5	0.1	3.5	0.0
93	3.3	3.2	0.0	3.3	0.0	3.3	0.0
94	3.0	3.0	0.0	3.1	0.1	3.1	0.0
95	2.8	2.8	0.0	2.9	0.0	2.9	0.0
96	2.6	2.6	0.0	2.7	0.0	2.7	0.0
97	2.5	2.4	0.0	2.5	0.0	2.5	0.0
98	2.3	2.3	0.0	2.3	0.0	2.3	0.0
99	2.1	2.1	0.0	2.2	0.0	2.1	0.0
100	2.0	2.0	0.0	2.0	0.0	2.0	0.0

Note: See note to Table 1.

**Table A2: Comparison of period life expectancies with alternative dynamic life expectancies at single-year age intervals: Canada, 2001, females**

Age <i>x</i>	Period life expectancy	Dynamic life expectancy ( $e_{xx}$ ) based on $q_x$ rates of change in the previous --					
		10 years		25 years		50 years	
		$e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx} - e_x$
0	82.0	87.3	5.2	91.8	9.8	91.9	9.8
1	81.4	86.6	5.2	91.1	9.7	91.2	9.8
2	80.4	85.6	5.1	90.0	9.6	90.1	9.6
3	79.5	84.5	5.1	88.9	9.4	89.0	9.5
4	78.5	83.5	5.0	87.8	9.3	87.8	9.4
5	77.5	82.4	4.9	86.7	9.2	86.7	9.3
6	76.5	81.4	4.9	85.5	9.0	85.6	9.1
7	75.5	80.3	4.8	84.4	8.9	84.5	9.0
8	74.5	79.2	4.7	83.3	8.8	83.4	8.9
9	73.5	78.2	4.7	82.2	8.7	82.3	8.8
10	72.5	77.1	4.6	81.0	8.5	81.1	8.6
11	71.5	76.0	4.5	79.9	8.4	80.0	8.5
12	70.5	75.0	4.5	78.8	8.3	78.9	8.4
13	69.5	73.9	4.4	77.7	8.1	77.8	8.2
14	68.5	72.9	4.3	76.5	8.0	76.6	8.1
15	67.6	71.8	4.3	75.4	7.9	75.5	8.0
16	66.6	70.8	4.2	74.3	7.7	74.4	7.8
17	65.6	69.7	4.1	73.2	7.6	73.3	7.7
18	64.6	68.6	4.0	72.1	7.5	72.2	7.6
19	63.6	67.6	4.0	71.0	7.3	71.1	7.5
20	62.7	66.6	3.9	69.8	7.2	70.0	7.3
21	61.7	65.5	3.8	68.7	7.1	68.9	7.2
22	60.7	64.4	3.8	67.6	6.9	67.7	7.1
23	59.7	63.4	3.7	66.5	6.8	66.6	6.9
24	58.7	62.3	3.6	65.4	6.7	65.5	6.8
25	57.8	61.3	3.5	64.3	6.5	64.4	6.7
26	56.8	60.2	3.5	63.2	6.4	63.3	6.5
27	55.8	59.2	3.4	62.0	6.2	62.2	6.4
28	54.8	58.1	3.3	60.9	6.1	61.1	6.3
29	53.8	57.1	3.2	59.8	6.0	59.9	6.1
30	52.9	56.0	3.2	58.7	5.8	58.8	6.0
31	51.9	55.0	3.1	57.6	5.7	57.7	5.9
32	50.9	53.9	3.0	56.5	5.6	56.6	5.7

**Table A2:** (Continued)

Age	Period life expectancy	Dynamic life expectancy ( $e_{xx}$ ) based on $q_x$ , rates of change in the previous --					
		10 years		25 years		50 years	
$x$	$e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$
33	49.9	52.9	2.9	55.3	5.4	55.5	5.6
34	48.9	51.8	2.9	54.2	5.3	54.4	5.4
35	48.0	50.8	2.8	53.1	5.2	53.3	5.3
36	47.0	49.7	2.7	52.0	5.0	52.2	5.2
37	46.0	48.7	2.6	50.9	4.9	51.1	5.0
38	45.1	47.6	2.6	49.8	4.7	50.0	4.9
39	44.1	46.6	2.5	48.7	4.6	48.9	4.8
40	43.1	45.5	2.4	47.6	4.5	47.8	4.6
41	42.2	44.5	2.3	46.5	4.3	46.7	4.5
42	41.2	43.5	2.3	45.4	4.2	45.6	4.4
43	40.3	42.4	2.2	44.3	4.1	44.5	4.2
44	39.3	41.4	2.1	43.3	4.0	43.4	4.1
45	38.4	40.4	2.0	42.2	3.8	42.3	4.0
46	37.4	39.4	2.0	41.1	3.7	41.3	3.8
47	36.5	38.4	1.9	40.0	3.6	40.2	3.7
48	35.5	37.3	1.8	39.0	3.4	39.1	3.6
49	34.6	36.3	1.7	37.9	3.3	38.1	3.5
50	33.7	35.3	1.7	36.9	3.2	37.0	3.3
51	32.8	34.3	1.6	35.8	3.1	36.0	3.2
52	31.8	33.3	1.5	34.8	2.9	34.9	3.1
53	30.9	32.4	1.4	33.7	2.8	33.9	3.0
54	30.0	31.4	1.4	32.7	2.7	32.9	2.8
55	29.1	30.4	1.3	31.7	2.6	31.8	2.7
56	28.2	29.5	1.2	30.7	2.5	30.8	2.6
57	27.3	28.5	1.2	29.7	2.4	29.8	2.5
58	26.5	27.6	1.1	28.7	2.3	28.8	2.4
59	25.6	26.6	1.0	27.7	2.1	27.8	2.3
60	24.7	25.7	1.0	26.8	2.0	26.9	2.2
61	23.9	24.8	0.9	25.8	1.9	25.9	2.0
62	23.0	23.9	0.8	24.8	1.8	24.9	1.9
63	22.2	23.0	0.8	23.9	1.7	24.0	1.8
64	21.3	22.1	0.7	23.0	1.6	23.1	1.7
65	20.5	21.2	0.7	22.1	1.5	22.1	1.6
66	19.7	20.3	0.6	21.1	1.4	21.2	1.5
67	18.9	19.5	0.6	20.3	1.4	20.3	1.4

**Table A2:** (Continued)

Age <i>x</i>	Period life expectancy $e_x$	Dynamic life expectancy ( $e_{xx}$ ) based on $q_x$ rates of change in the previous --					
		10 years		25 years		50 years	
		$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$	$e_{xx}$	$e_{xx} - e_x$
68	18.1	18.7	0.5	19.4	1.3	19.4	1.3
69	17.3	17.8	0.5	18.5	1.2	18.6	1.2
70	16.6	17.0	0.4	17.7	1.1	17.7	1.2
71	15.8	16.2	0.4	16.8	1.0	16.9	1.1
72	15.1	15.4	0.4	16.0	0.9	16.1	1.0
73	14.3	14.7	0.3	15.2	0.9	15.3	0.9
74	13.6	13.9	0.3	14.4	0.8	14.5	0.8
75	12.9	13.2	0.3	13.7	0.7	13.7	0.8
76	12.2	12.5	0.2	12.9	0.7	12.9	0.7
77	11.6	11.8	0.2	12.2	0.6	12.2	0.6
78	10.9	11.1	0.2	11.5	0.6	11.5	0.6
79	10.3	10.4	0.1	10.8	0.5	10.8	0.5
80	9.7	9.8	0.1	10.1	0.5	10.1	0.5
81	9.1	9.2	0.1	9.5	0.4	9.5	0.4
82	8.5	8.6	0.1	8.9	0.4	8.9	0.4
83	8.0	8.0	0.1	8.3	0.3	8.3	0.3
84	7.4	7.5	0.1	7.8	0.3	7.7	0.3
85	7.0	7.0	0.1	7.2	0.3	7.2	0.3
86	6.5	6.6	0.1	6.7	0.2	6.7	0.2
87	6.1	6.1	0.1	6.3	0.2	6.3	0.2
88	5.7	5.7	0.1	5.9	0.2	5.8	0.2
89	5.3	5.4	0.1	5.5	0.2	5.4	0.2
90	4.9	5.0	0.0	5.1	0.2	5.1	0.1
91	4.6	4.7	0.0	4.8	0.1	4.7	0.1
92	4.3	4.4	0.0	4.4	0.1	4.4	0.1
93	4.0	4.0	0.0	4.1	0.1	4.1	0.1
94	3.7	3.8	0.0	3.8	0.1	3.8	0.1
95	3.5	3.5	0.0	3.6	0.1	3.5	0.1
96	3.2	3.3	0.0	3.3	0.1	3.3	0.1
97	3.0	3.0	0.0	3.1	0.1	3.0	0.0
98	2.8	2.8	0.1	2.9	0.1	2.8	0.0
99	2.6	2.6	0.0	2.7	0.1	2.6	0.0
100	2.4	2.5	0.1	2.5	0.1	2.4	0.0

Note: See note to Table 1.