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Research Article

Anticipatory child fostering and household economic security in Malawi

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Lauren K. Bachan 1

Abstract

BACKGROUND

While there is a rich literature on the practice of child fostering in sub-Saharan Africa, little is known about how fostering impacts receiving households, as few studies consider household conditions both before and after fostering. Despite the fact that circumstances surrounding fostering vary, the literature's key distinction of fostering is often drawn along the simple line of whether or not a household is fostering a child. This paper argues that anticipation of fostering responsibilities, in particular, is a useful dimension to distinguish fostering experiences for receiving households.

OBJECTIVE

This paper examines the relationship between receiving a foster child and subsequent changes in household wealth. Particular emphasis is placed on how these changes are conditioned by differing levels of anticipation of the fostering event.

METHODS

This study uses data from Tsogolo la Thanzi (TLT), a longitudinal survey in Balaka, Malawi. Using data from 1754 TLT respondents, fixed effects pooled time-series models are estimated to assess whether and how receiving a foster child changes household wealth.

RESULTS

This paper demonstrates the heterogeneity of fostering experiences for receiving households. The results show that households that anticipate fostering responsibilities experience a greater increase in household wealth than both households that do not foster and those that are surprised by fostering.

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CONCLUSION

Households that anticipate fostering responsibilities exhibit the greatest increase in household wealth. While fostering households that do not anticipate fostering responsibilities may not experience these gains, there is no evidence to indicate that such households are negatively impacted relative to households that do not foster. This finding suggests that additional childcare responsibilities may not be as detrimental to African households as some researchers have feared.

1. Introduction

Child fostering—the custom of children living outside of the natal home—is practiced throughout much of sub-Saharan Africa (Bledsoe 1990; Bledsoe and Isiugo-Abanihe 1989; Goody 1982; Madhavan 2004; Monasch and Boerma 2004; Urassa et al. 1997). Although the practice varies from region to region, fostering is commonly thought to be a mutually beneficial arrangement between sending and receiving households (Bledsoe 1990; Caldwell 1997; Drah 2012; Isiugo-Abanihe 1985). Empirical work recognizes child fostering as a strategy African households use to distribute the cost of childrearing and offset economic insecurities by sending children to be fostered by other households (Akresh 2005; Eloundou-Enyegue and Shapiro 2004).

Three facts about child fostering have been well established in the qualitative and ethnographic literature but are underdeveloped in survey research. First, child fostering is part of a larger negotiation process that takes place among extended family networks (Akresh 2005; Caldwell 1997). Second, receiving a foster child happens under a variety of circumstances (Bledsoe and Isiugo-Abanihe 1989; Goody 1982). Third, ethnographies further suggest that households that receive foster children benefit from the added labor of the foster child and from the social insurance of investing in other people's children (Bledsoe 1990; Caldwell 1997). However, the tangible benefits for receiving families remain largely unexplored and are less established in quantitative research.

This paper explores the economic consequences of child fostering for receiving house-holds in southern Malawi. Using intensive longitudinal data, I track the incidence of child fostering over an approximately two and a half year time period in order to test competing hypotheses about how fostering a child triggers changes in household wealth. The results demonstrate that receiving a foster child as part of an anticipated arrangement is associated with improved household wealth over time. Although further research is needed to examine the mechanisms underlying these findings, this study provides initial insights

into how receiving households are affected by child fostering and how the heterogeneity of the fostering experience differentially contributes to household economic outcomes.

2. Households and child fostering in sub-Saharan Africa

Households are units that consist of individuals who make joint decisions regarding consumption and production, labor force participation, savings, and capital acquisition (Becker 1991). The description of a household as an economic unit is particularly relevant in much of sub-Saharan Africa, where subsistence agriculture is a primary livelihood for many people. Household structure in sub-Saharan Africa tends to be complex, with members extending beyond the nuclear family. While the specifics vary from society to society, households containing multiple generations (vertical complexity) and extended family members (horizontal complexity) are common throughout the region (Bongaarts 2001; McDaniel and Zulu 1996; Van de Walle 2006).

Child fostering adds to the compositional complexity of households in sub-Saharan Africa. Unlike Western societies, where fostering is regulated by state governments, the vast majority of child fostering in Africa is akin to informal parenting, namely when a child is reared by another family (or multiple families) for extended periods of time in the absence of government intervention (Goody 1982). Much of the classic fostering literature focuses on West Africa, but the practice of child fostering is widespread across other regions of Africa as well (Bicego, Rutstein, and Johnson 2003; Madhavan 2004; McDaniel and Zulu 1996; Monasch and Boerma 2004; Urassa et al. 1997). Furthermore, fostering in Africa takes place under a variety of circumstances. Children are fostered when their natal parents can no longer fulfill their parental roles because of illness, divorce, or death (Goody 1982). Foster children are also frequently sent to live with nonnatal family members as part of a deliberate, and often mutually-beneficial, arrangement (Madhavan 2004; McDaniel and Zulu 1996).

A rich literature on the functions of child fostering in sub-Saharan Africa demonstrates that, in addition to managing crises (e.g., death of a parent), fostering is used to strengthen kinship ties; share the cost of childrearing (Goody 1982); and smooth demographic inequalities, such as childlessness or unfavorable gender composition of existing children (Bledsoe and Isiugo-Abanihe 1989; Isiugo-Abanihe 1985; Lloyd and Desai 1992). Additionally, fostering acts as a safety net for families and often expands children's opportunities for upward social mobility, particularly in the form of access to education and occupational training (Ainsworth and Filmer 2006; Akresh 2009; Bledsoe 1990; Isiugo-Abanihe 1985).

3. Household economics and fostering

Studies of the various purposes of child fostering have emphasized that the practice of sending and receiving children is a strategy households employ to distribute the costs and benefits of children. In other words, child fostering is one way in which households can offset economic insecurities (Akresh 2005; Caldwell 1997; Eloundou-Enyegue and Shapiro 2004; Isiugo-Abanihe 1985). Such frameworks imply that the consequences of fostering for receiving households should be measured in economic terms. However, despite the theoretically established economic links between receiving and households, the questions of whether and to what extent fostering triggers changes in wealth for receiving households remain unanswered.

This gap stems from two fundamental characteristics of the literature. First, since the onset of the AIDS epidemic in the region, research on fostering in Africa has almost entirely focused on the educational, health, and developmental disadvantages of children—particularly of orphans (Beegle et al. 2010a; Beegle et al. 2010b; Case et al. 2004; Cluver et al. 2007; Cluver and Orkin 2009; Deininger et al. 2003; Monasch and Boerma 2004; Subbarao et al. 2001; for a notable exception, see Ainsworth and Filmer 2006). Evidence that foster children—especially those who are not biologically related to the household head—face disadvantages has given the impression that households have been struggling to meet the demands of fostering. However, there are reasons to believe that fostering is not overwhelming families. In the face of the AIDS epidemic, households have shown a remarkable ability to care for both orphaned and non-orphaned foster children (Caldwell 1997; Grant and Yeatman 2012; Hosegood et al. 2007a; Monasch and Boerma 2004).

Second, the small literature that *does* consider fostering and household economics focuses primarily on the absolute differences between fostering and non-fostering households.² From this body of research we know that foster children tend to live in wealthier households (Beegle et al. 2010a; Bicego, Rutstein, and Johnson 2003; Weinreb, Gerland, and Fleming 2008), suggesting that selection effects are key in understanding foster child placement. However, without explicitly measuring *change* in wealth in response to fostering a child, we cannot discern with any certainty whether foster children are initially placed in wealthier households or whether households become wealthier as a result of receiving.

Further complicating matters is the fact that empirical research treats fostering conceptually and methodologically as an "unanticipated shock" (literally, such as in Deininger et al. 2003). Much of the broader literature, however, indicates that fostering happens un-

²One notable exception comes from a case study in Uganda, for which researchers used two waves of panel data to show that receiving a foster child was associated with a significant reduction in household investment (as measured by agricultural, structural, and transport equipment) over an eight year period (Deininger, Garcia, and Subbarao 2003). This study, however, conceptualizes fostering solely as a binary shock to the household.

der a variety of circumstances (Goody 1982; Madhavan 2004; McDaniel and Zulu 1996), and research on decision-making and networks in Africa suggests that fostering a child is *not* normally an unanticipated shock. Rather, fostering is part of a process of constant negotiation and risk-assessment, in which social networks and extended family play integral roles (Akresh 2005; Caldwell 1997). Foster parents often volunteer to take children (Bledsoe and Isiugo-Abanihe 1989) or, at a minimum, anticipate that they will do so in the future (Johnson-Hanks 2006; Ntozi 1995; Trinitapoli and Weinreb 2012). Such evidence suggests that anticipation is an important dimension by which to examine the fostering experiences of receiving households.

3.1 Epidemiological context

Over the past 30 years the AIDS epidemic is thought to have disrupted the long-standing fostering process by increasing the number of orphans.³ Evidence that less-preferred types of caretakers (e.g., poor or non-traditional family members) are beginning to play a larger role in fostering children (Howard et al. 2006; Nyambedha, Wandibba, and Aagaard-Hansen 2003) indicates that AIDS has altered the normative patterns of fostering. For instance, in her in-depth study of three communities across Botswana (a country with high HIV-prevalence), Dahl (2009) finds that many community members believe that traditional kin-based fostering is failing to provide sufficient care for children.

Despite any disruptions to fostering in the context of the AIDS epidemic, the nature of the disease may have elongated the fostering negotiation/expectation process, making anticipation of fostering responsibilities a pertinent dimension by which to evaluate if and how receiving a foster child impacts household wealth. AIDS has a long latent period, and its symptoms are almost certainly diagnosed by family members and friends as the disease progresses. Sick parents make contingency plans for their children in preparation for their own impending mortality (Klaits 2010). For example, when Malawian parents know they are HIV positive, they are more likely to invest in their children's future by enrolling them in school (Grant 2008). Across sub-Saharan Africa, children are at the forefront of AIDS-related decision making and conversations (Dahl 2009). In rural Malawi, people recognize AIDS as a "profound danger" to children (Watkins 2004:694) and often evoke child wellbeing into the discourse on prevention strategies (Smith and Watkins 2005; Watkins 2004).

³Between 1990 and 2009, the number of children in sub-Saharan Africa who were orphaned due to AIDS increased from less than one million to 14.8 million (UNAIDS 2007, 2010).

3.2 Competing hypotheses on economic outcomes for receiving households

In the simplest terms there are two ways in which fostering a child could impact the wealth of a receiving household: Wealth could either increase or decrease in response to fostering. Existing empirical and theoretical evidence lends support for both competing possibilities.

Drawing from Caldwell's wealth flow theory (1976; 1983), which acknowledges the unique value of children in developing settings, fostering a child may increase household wealth through added domestic labor. Foster children (especially girls) can be sources of cheap domestic labor within fostering households (Ainsworth and Filmer 2006; Bledsoe 1990; Goody 1982), freeing up time for other household members to participate in economically productive endeavors.⁴ Benefits of fostering can also extend beyond what a household receives from the foster child. Mende grandmothers in Sierra Leone and Igbo women in Nigeria often receive food and money from natal parents in return for fostering their grandchildren (Bledsoe and Isiugo-Abanihe 1989; Isiugo-Abanihe 1985). In other contexts, people who take in children from outside their kin group do so with the explicit expectation of receiving material compensation (Dahl 2009). Taken together, this evidence suggests the following:

Hypothesis 1: Household wealth will increase in response to receiving a foster child.

Adding a young dependent to a household could, on the other hand, drain family resources leading to a decrease in household wealth. This may be especially true when households are already impoverished and/or when fostering takes place in response to a family crisis. Households that foster children in response to parental illness or death often suffer financially, as other ill effects of a family crisis (e.g., funeral and/or medical costs) are exacerbated by increased dependency ratios (Hosegood et al. 2007b). Additionally, the current AIDS epidemic and resulting increase in the number of orphans in the region may be intensifying the challenges foster families face (Grant and Yeatman 2012; Hosegood et al. 2007b). Recent studies show that in high HIV-prevalent areas, poorer households are playing a larger role in child fostering today than they have in the past (Bicego, Rutstein, and Johnson 2003). While we have thus far been unable to establish if these households were poor when they began fostering or became poor because of foster-

⁴Wealth flow perspectives also highlight *disincentives* for natal families to out-foster their children. Because children are significant sources of social insurance, sending families may be reluctant to voluntarily foster out their children. However, fostering is not a zero-sum game with only one family claiming the benefits of a child. Out-fostering does not typically entail the relinquishment of parental rights (Goody 1982), thus natal families can still expect to receive future benefits from their out-fostered children. Seen this way, fostering is a strategic decision on part of the sending family such that the benefits of sending the child to be fostered outweigh those of keeping the child (Akresh 2005).

ing, the following could be true:

Hypothesis 2: Household wealth will decrease in response to receiving a foster child.

4. Data, measures, and methods

4.1 Data

To test these competing hypotheses, this study focuses on Malawi, a country with a strong tradition of child fostering (Bandawe and Louw 1997; Munthali 2002). In 2010, approximately 28% of households were fostering a child under the age of 18 (National Statistical Office and ICF Macro 2011). Malawi is experiencing one of the most severe AIDS epidemics in sub-Saharan Africa, with an adult prevalence rate estimated at 11%. AIDS alone has orphaned more than half a million children in the country (UNAIDS 2010). Because institutionalized foster care is considered a last resort in Malawi, the majority of additional child care responsibilities resulting from the AIDS epidemic takes place within households (Munthali 2002).

The data used come from Tsogolo la Thanzi (TLT), a panel study fielded in the Balaka district of southern Malawi. TLT collected eight waves of data at four-month intervals between June 2009 and January 2012 from an initial random sample of 1500 female and 600 male respondents (initial response rate of 95.6%). Respondents were randomly selected from a sampling frame of 15- to 25-year-olds living in census enumeration areas within seven kilometers of the district capital—an area including Balaka township and the surrounding rural villages. I use data from Waves 1-8, making adjustments to the data structure to reflect the sequencing of events and the changes being measured (discussed in greater detail below).

Three features make the TLT data uniquely suitable for assessing changes within households. First, TLT gathers both individual-level characteristics and household-level information, including ownership of household goods and household rosters, which each respondent updates at every wave. Second, the longitudinal design is optimal for measuring change over time. Finally, the short intervals between survey waves allow economic fluctuations to be examined over brief periods of time, something researchers are often unable to examine in panel studies in which years pass between data collection waves.

One limitation of the TLT data is that it is a sample of young adults, many of whom are not the head of their household and some of whom share the same household. I handle the latter limitation by including only one respondent from each household sampled. While I do not assume that respondents are the primary caretaker of any fostered child, I do assess each respondent's socio-demographic characteristics—in addition to household characteristics—as proxies for household dynamics. In other words, I assume that

the respondent's characteristics (e.g., income, education, etc.) reasonably reflect their household.

I restrict the TLT sample in four ways. First, to accommodate my analytic approach (discussed below), I omit data from Wave 1—with the exception of the anticipation measure—from the analysis. This omission results in the exclusion of 143 individuals who only contributed data to Wave 1 or who provided fewer than two waves of data in Waves 2-8 (the average number of waves per respondents is 6.3). Second, because I am interested in households, I remove 41 respondents who are living primarily in educational institutions (e.g., boarding schools or universities) rather than a traditional household. Third, 119 households in the TLT sample contain more than one respondent. In these cases I randomly select one member to represent his or her household. This results in removing 129 respondents.⁵ Finally, I further restrict the sample to respondents for whom complete information was available on the measures used in this analysis. Missing data are rare and employing listwise deletion removed an additional 13 respondents, resulting in a final study sample of 1754 individuals who contribute 10,987 person-waves to the pooled data. As shown in Appendix A (Table A1), the initial and analytic samples are quite comparable, suggesting that sample selection bias is not introduced by these restrictions.

4.2 Measures

4.2.1 Household wealth

The outcome measure is the household's wealth. Wealth is measured on a linear index comprising nine durable goods and one household asset. Unlike wealth indices that include several structural assets (e.g., flooring material, water supply, sanitation, etc.), this index relies heavily on durable goods within the household in order to better capture economic fluctuation over the relatively short timespan of the study. The goods used in the index include a bed with a mattress, a television, a radio, a landline or mobile phone, a refrigerator, a bicycle, a motorcycle, an animal-drawn cart, and an automobile. The wealth measure also considers whether or not the household has electricity.

Weights are calculated for each asset using principal-components analysis following the same procedure used to construct the Demographic Health Survey wealth index. The resulting index places households on a continuous scale relative to the sample (Filmer and Pritchett 2001; Howe, Hargreaves, and Huttly 2008; Rutstein and Johnson 2004). This approach to measuring wealth–something difficult to accurately estimate in devel-

⁵In some cases, there were more than two respondents per household, which is why more than 119 respondents were removed at random. Supplementary analyses including *all* respondents showed that the results are not significantly different from the results presented in this paper.

oping countries—has been validated by previous researchers (Filmer and Pritchett 2001; Howe, Hargreaves, and Huttly 2008). To ensure factorial invariance, the weights are kept consistent across waves. No item received a negative weight. In this study, television and electricity are assigned the largest weights and the lowest weights are given to an animal-drawn cart and bicycle.

4.2.2 Child fostering

To explore the effect of fostering on household wealth, I measure the incidence of fostering within the study period. Because it is impossible to know whether a respondent's household was fostering a child before the TLT study began, I assume that no households were fostering at the baseline wave. While it is likely that some respondents' households were fostering a child at the outset of the study, this assumption is reasonable for the purpose of making inferences about the consequences of *new* fostering experiences. The incidence of child fostering is measured in two distinct ways. First, based on each respondent's self-report of having a non-biological child join his or her household since his or her last interview, I use a binary indicator for whether each respondent's household fostered a child between waves (1=yes, 0=no).

The second measure is a categorical variable that takes into account respondents' anticipation of the fostering event. I created this variable using the binary indicator in combination with information about anticipation from the preceding wave. At each wave, respondents were asked, "In the next year, how likely is it that you will foster a new child in your household?". Responses were measured through an interactive technique wherein the respondent was given a set of 10 beans and asked to shift the number of beans to a small plate to represent the likelihood of a future event, with a greater number of beans indicating a higher likelihood.⁶ Combining answers to this question with the report of whether the respondent's household fostered a child in the subsequent wave. I categorize respondents into four groups: 1) respondents whose household fostered a child but indicated in the previous wave that there was no likelihood of their household fostering in the coming year (zero beans at the previous wave), 2) those whose household fostered and anticipated those responsibilities (10 beans at the previous wave), 3) respondents whose household fostered a child but were uncertain about whether or not they would foster in the future (one to nine beans at the previous wave), and 4) respondents whose household did not foster between waves (reference group). These cut-points for the categories were

⁶The interviewer introduces this method of questioning by asking respondents simple questions about frequent events. The interviewer gradually moves to more complex questions to ensure the respondents understand the interactive nature of the questioning. This approach has previously been used and validated by Delavande and Kohler (2009) and is described in detail in Trinitapoli and Yeatman (2011).

not determined empirically to ensure any specific distribution or results, but were defined conceptually.⁷

To clarify the complexity of this measure, Figure 1 illustrates how the longitudinal data were leveraged to create the categorical fostering variable. Figure 1 depicts hypothetical Respondents A, B, and C, all of whom fostered a child at least once during the seven-wave interval, while hypothetical Respondent D did not foster during this time period. Respondent A reported fostering a child at Wave 2 after indicating at Wave 1 that there was no likelihood (zero beans) of doing so. Respondent A is a fosterer in the binary sense and, specifically, a "surprised fosterer" in categorical terms from Wave 2 onward. At Wave 4, Respondent B reported fostering a child and is treated as a fosterer over the subsequent waves. Because Respondent B did not contribute data to Wave 3, her level of anticipation at Wave 2 (10 beans) was used, categorizing her as an anticipated fosterer. Respondent C reported fostering twice: both at Wave 3 and Wave 5.9 This respondent remained a fosterer, in binary terms, from Wave 3 onward, but moved from an "uncertain" fosterer to an "anticipated" fosterer in the categorical measure based on the different level of anticipation preceding the second fostering event.

Figure 1: Construction of the anticipatory fostering variable

Connects the fostering event to the wave at which level of anticipation is used to categorize fostering by type.
 Number of beans indicates level of anticipation of fostering within the next year.

	Wave 1	Wave 2	Wave 3	Wave 4	Wave 5	Wave 6	Wave 7	Wave 8
Respondent A	0 beans —Non-Fos	terer —		(Surpri	ised) Fosterer –			
Respondent B	5 beans	10 beans	missed wave	*	-(Anticipated) I	Fosterer ———		
Respondent C	10 beans	6 beans	*	10 beans incertain) Foster	*	ipated) Fosterer_		
Respondent D	0 beans	0 beans	5 beans	7 beans — Non-Fosterer	5 beans	3 beans		

⁷Ancillary models were estimated to test different categories that define surprised, anticipated, and uncertain fostering. Expanding the surprised and anticipated categories by one bean each and restricting the uncertain category by two beans did not change the results of the multivariate analysis. A discussion and presentation of these models can be found in Appendix B.

⁸There were only eight respondents who fostered a child and did not contribute data to the preceding wave. In these cases, the anticipation used to categorize them was always obtained from within one year prior to fostering.

⁹For respondents who reported fostering more than one time over the course of the study, the categorical variable of fostering was updated to reflect the most recent fostering experience, but the binary fostering variable remained constant.

4.2.3 Control variables

I adjust for time-varying socio-demographic variables that the literature establishes as important predictors of wealth. All control variables are treated as intervening variables that is, they happen after fostering has occurred. Household size is a continuous measure of the number of people who regularly sleep at the respondent's house and was derived from the household roster, which was updated at every wave. To measure the effect of adding additional household members other than a foster child and to ensure the effect of fostering is not captured in changes in household size, one household member was subtracted for households that fostered. Marital status is a binary measure that distinguishes respondents who are currently single from those who are married or cohabiting. Educational attainment is a continuous measure of the highest level of education completed by the respondent. In addition to educational attainment I include a dummy variable that signifies whether the respondent was enrolled in school at each interview wave. Income is measured in Malawian Kwacha, and represents the amount of money the respondent reported making in the month preceding his or her interview (referred to as "monthly income" in subsequent tables). An indicator of survey wave is included as a proxy for time and adjusts for any maturation effect as individuals move through the life course. 10 Finally, a dummy variable for whether the respondent moved houses between interviews is included to account for discontinuity in the household rosters.

4.3 Analytic approach

I use fixed-effects pooled time series models to estimate the effect of fostering on subsequent changes in household wealth. Fixed effects procedures use within-individual differences to estimate the difference between person-level variance across groups that did and did not experience a given event within the observed time period. Fixed effects models have several advantages for estimating the effects of events on continuous outcomes in panel data, two of which are particularly salient for the present study (see Allison 1994 and Johnson 1995 for a detailed description of the costs and benefits of this modeling strategy). First, fixed effects models allow for the inclusion of respondents who have not contributed information to every wave. This is especially appropriate for intensive longitudinal studies, such as TLT, in which respondents may miss waves due to temporary migration or illness without attriting from the sample entirely. Second, fixed-effects procedures account for both observed and unobserved time-invariant characteristics. This

¹⁰ After respondents who skipped an interview or interviewed at an unscheduled time have been accounted for, the time between interviews ranges from one to seven months for the analytic sample. However, the average length of time between interviews was four months, with 80% of respondents interviewing consistently at that interval. Supplementary analyses, which use the exact interview date, produce highly similar results.

ensures the model estimates are independent of selection bias and all other time-invariant differences between respondents (Allison 1994).

Fixed effects models rely on a mean-deviation algorithm that computes the means over time for all time-varying variables (both dependent and independent) for each individual. These person-specific means are then subtracted from the observed values of each variable and the difference in the dependent measure is regressed on the difference in the independent variable(s) (Allison 2009). Like other family transitions, the full effects of fostering on households may not be instantaneous. Therefore, comparing household wealth pre- and post-fostering is more appropriate than estimating the immediate effect of receiving a foster child in a one-wave interval. Thus for fixed effects estimation, once a respondent's household has fostered a child, I considered it a fostering household in all subsequent waves (as indicated in Figure 1).

To ensure that exposure to fostering precedes changes in household wealth, I structure the data so that anticipation and reported fostering are combined (discussed above) and are used to predict changes in *future* household wealth. In other words, fostering a child between Waves 1 and 2 is used to assess changes in household wealth that take place from Wave 3 onward. Similarly, fostering that occurred between Waves 5 and 6 is used to assess changes in household wealth in Waves 7 and 8. This technique ensures that a contemporaneous change in household wealth can not be confused with a real effect of fostering or that a change in household wealth prompted a household to foster a child.

5. Results

Table 1 provides a descriptive overview of the pooled time-series data. The characteristics of the pooled data are highly consistent with the cross sectional data taken from Wave 2 (see Appendix A, Table A2), with some expected differences (e.g., increasing educational attainment and a higher percentage of marriage as respondents move through the life course, which is reflected in the pooled data).

Table 1: Fostering characteristics and descriptive statistics for pooled time-series data

Variable	Mean	S.D.	Range
Household Wealth	-0.09	1.70	-1.85 - 9.36
% Fostered a child	19.4	0.40	0 - 1
% Surprised	2.9	0.17	0 - 1
% Anticipated	2.7	0.16	0 - 1
% Uncertain	13.8	0.35	0 - 1
Household Size	4.8	2.18	1 - 18
Years of Education	7.3	2.73	0 - 13
% Married	44.1	0.50	0 - 1
% Student	34.1	0.47	0 - 1
Monthly Income (thousands of Kwacha)	2.3	6.99	0 - 250
% Moved household	12.6	0.33	0 - 1
N=10,987 person-waves representing 1754	individuals		

Over the course of the study, most households experienced a net gain in wealth, but this gain was not universal across households. An examination of the household wealth measure (not shown) reveals that 42% of households experience a net increase in wealth between Waves 2 and 8. Approximately one third (31%) of households experienced a net decline in household wealth, while about a quarter showed no net change over the same time period.

Within each wave analyzed, the incidence of new fostering among households ranged between 4 and 11%. Exposure to fostering within the pooled data represents 19% of all person-waves analyzed. Table 1 shows the percent of fostering experiences disaggregated into anticipation categories. Surprised and anticipated fostering make up roughly equal proportions of fostering experiences (15% and 13%, respectively), while, as expected, uncertain fostering represents the majority of fostering circumstances.

Table 2 shows the bivariate association between the fostering categories and key independent variables. This table is split into two panels: the top panel, which measures variables at Wave 1 of the study before the measured fostering events took place, ¹¹ and a bottom panel, which shows the average across all observations in the pooled dataset. The top panel reveals that at the beginning of the observed interval, respondents whose household eventually fostered a child are, on average, more educated and wealthier than

¹¹ The most recent fostering experience was used to categorize respondents whose household fostered multiple times in the interval observed.

respondents whose household did not foster a child. Interestingly, households that eventually fostered and anticipated doing so began the study with lower household wealth than households that fostered under other circumstances. This suggests that wealth is not necessarily a pre-requisite for the ability to anticipate fostering responsibilities.

The patterns in the top panel of Table 2 are largely reflected in the pooled data (bottom panel of Table 2). The notable exception is among households who fostered a child and anticipated doing so. When observations are averaged across the pooled data, households that fostered out of anticipation are not less notably less wealthy than other types of households that fostered.

Table 2: Mean values for key variables by fostering categories

	Fostering Category				
Variable	Non-Fosterers	Surprised	Anticipated	Uncertain	
Household Wealth	-0.31	0.13	-0.12	0.20	
Household size	4.98	5.57	4.88	5.01	
Years of Education	7.01	7.73	7.24	7.55	
% Married	35.72	24.71	50.00	38.14	
Monthly Income (thousands of Kwacha)	1.95	1.74	2.58	2.32	
N=1750 individuals for whom information	was available at W	ave 1			
Poole	ed time-series data	a			
Household Wealth	-0.17	0.31	0.23	0.25	
Household Size	4.95	4.78	4.52	4.29	
Years of Education	7.19	8.23	7.55	7.79	
% Married	42.90	31.33	58.02	50.95	
Monthly Income (thousands of Kwacha)	2.21	1.78	4.16	2.55	

5.1 Multivariate analysis

I begin by assessing the general pattern of wealth over time for households. I then estimate the effects of fostering as a binary status by comparing those whose household fostered a child to those who did not. Before examining the role of anticipation, I establish the validity of TLT's measure of anticipation of fostering. Finally, to understand whether and how anticipation of fostering may produce differential outcomes for foster families, I estimate models that treat fostering as an event conditioned by anticipation to assess its consequences for household wealth.

5.1.1 Fostering as a binary status

Model 1 in Table 3 contains the results of the first fixed-effects time series model, which simply evaluates how household wealth changed over the observed interval. In the absence of any controls, household wealth in this sample increased marginally over time. Model 2 in Table 3 estimates the effect of fostering, as a binary status, on household wealth, controlling for time. This model indicates that fostering a child is not associated with statistically discernible changes in household wealth. This suggests that child fostering has no impact on the wealth of receiving households. However, evidence from previous studies indicates that receiving a foster child is a heterogeneous experience. Because fostering may matter for households under certain circumstances, I turn my attention to evaluating whether the effect of fostering is conditioned by anticipation.

Table 3: Fixed effects regression models of the effect of time and fostering on household wealth

	Model 1	Model 2
Time	0.027*** (0.004)	0.026*** (0.004)
Fostering		0.022 (0.032)
Constant	-0.223*** (0.019)	-0.221^{***} (0.019)
Person-waves	10,987	10,987
Individuals	1754	1754
Within-individual \mathbb{R}^2	0.006	0.006

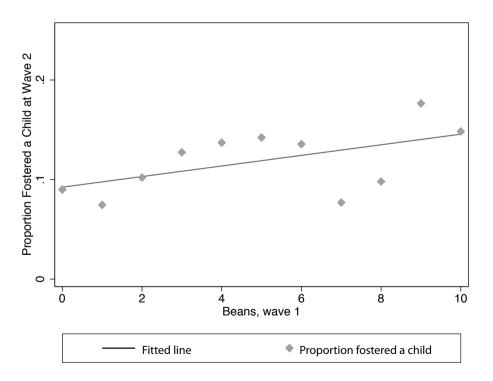
Notes: Standard errors in parentheses. *** p<0.001, ** p<0.01, * p<0.05

5.1.2 Fostering conditioned by anticipation

Before considering the role of anticipation in outcomes for foster families, I establish the validity of the anticipation measure of one selected household representative by examining its relationship with subsequent fostering. Figure 2 shows the proportion of respondents whose household fostered a child in Wave 2 by their level of anticipation (as measured in beans) in Wave 1, fitted with a regression line. This figure shows a clear relationship between the respondent's anticipation of fostering and actual future fostering.

The positive relationship between anticipation and future fostering is robust to a number of controls. Using logistic regression, I regress fostering at Wave 2 (1=yes, 0=no) on Wave 1 levels of anticipation, controlling for basic sociodemographic characteristics such as gender, age, household wealth, and marital status (also measured at Wave 1). The results indicate that the measure of anticipation is a meaningful predictor of future fostering, with a person's odds of fostering a child at Wave 2 increasing by 5% for every additional bean he or she shifted to the "positive" side of the plate in Wave 1 (p<0.05) (see Table A3 in Appendix A for full model estimates). Further investigation of the anticipation measure shows that anticipation of future fostering operates similarly across different types of respondents within the TLT sample. Supplementary analyses testing for interaction effects with anticipation provided no evidence to suggest that the predictive ability of anticipation of fostering differs by respondent's age, marital status, gender, education level, or socioeconomic status (models not shown but available upon request).

Figure 2: Relationship between anticipation of fostering and subsequent fostering



Despite the strong relationship between anticipation of fostering responsibilities and subsequent fostering, surprised fostering happened within the TLT sample. Nearly 10% of people who reported fostering at Wave 2 indicated at Wave 1 that there was no chance of their household fostering a child within the next year. Likewise, slightly less that 15% of those reporting fostering at Wave 2 had completely anticipated doing so during their Wave 1 interview.

Model 3 in Table 4 estimates the impact of fostering categorized by prior anticipation on change in household wealth. This model reveals that fostering had a statistically significant positive effect on household wealth when it was anticipated. Fostering under these circumstances was associated with a 0.15 gain in the household wealth score compared to respondents whose household did not foster within the observation period. Restructuring the data to ensure fostering preceded changes in household wealth allows us to reasonably interpret this as an actual gain in wealth rather than a selection effect.

For households that fostered under surprised or uncertain circumstances, we observe no such advantage. These households experienced neither a significant increase nor decrease in wealth in comparison to households that did not experience fostering within the interval observed. However, rotating the reference categories to examine pairwise relationships shows that fostering had different consequences for those households that anticipated the event and those that were surprised by it (model not shown). Specifically, households that anticipated fostering experience a 0.24 increase in wealth over households that fostered by surprise (p<0.05). Uncertain fostering households, however, remain statistically indistinguishable from anticipated fosterers in terms of changes to household wealth.¹²

One explanation for these findings could be that the anticipation of fostering—irrespective of *actually* fostering a child in the future—is itself associated with an increase in wealth. Perhaps people expect that their household will come into wealth in the near future, prompting them to also anticipate helping extended family by taking on additional childcare responsibilities. To rule out this possibility, I estimate the effect of the change in the anticipation of future fostering on future changes in household wealth. This analysis shows that anticipation in and of itself has no relation to household wealth (see Table A4 in Appendix A). Additionally, there are only weak correlations (r<0.3) between respondents' anticipation of fostering and their anticipation of other wealth-related future events, such as the ability to save money, open a bank account, and/or open a business. Taken together, this suggests that people who anticipate their household will foster are not simultaneously anticipating an increase in wealth.

¹²These results remain robust to addition of control variables.

Table 4: Fixed effects regression models of the effect on household wealth by anticipatory fostering status

	Model 3	Model 4
Time	0.026*** (0.004)	0.023*** (0.004)
<i>Prior Anticipation (ref. = non-fosterers)</i>		
Surprised Fostering	-0.085 (0.067)	-0.032 (0.067)
Anticipated Fostering	0.152^* (0.072)	$0.192^{**} (0.071)$
Uncertain Fostering	0.023 (0.036)	0.072^* (0.036)
Sociodemographic controls		
Household size		0.064^{***} (0.007)
Years of Education		0.047^* (0.020)
Student		0.098** (0.033)
Married		0.142*** (0.036)
Monthly Income (thousands of Kwacha)		0.006*** (0.002)
Household move		0.023 (0.025)
Constant	-0.221^{***}	-0.990***
	(0.019)	(0.150)
Person waves	10,987	10,987
Individuals	1754	1754
Within-individual R^2	0.007	0.020

Notes: Standard errors in parentheses. *** p<0.001, ** p<0.01, * p<0.05

Turning to Model 4 in Table 4, we see that the intervening sociodemographic controls do not explain how anticipation of fostering acts further increase household wealth. After controlling for a host of other changes, the significant effect of anticipated fostering on household wealth remains and even increases in effect size and significance. We also see that uncertain fostering is associated with a small increase in household wealth in this model. However, supplementary analyses suggest this latter effect is weak and more sensitive to the techniques used to categorize fostering (see Table B1 in Appendix B). Other covariates operate in ways consistent with what we would expect. The addition of a household member above and beyond the foster child increases household wealth, likely because the addition of an adult is accompanied by that individual's capital. In addition, being enrolled in school and an increase in the respondent's previous month's income act to increase the wealth of a respondent's household.

6. Summary and discussion

Child fostering in sub-Saharan Africa has long been regarded as a beneficial practice for households and children. However, little empirical work has explored the impact of fostering a child for receiving families. This study provides further insight into how fostering households cope with extra childcare responsibilities over time by testing two competing hypotheses about how fostering a child effects household wealth. Importantly, this study takes the perspective that fostering experiences are heterogeneous. Treating child fostering as conditioned by prior anticipation acknowledges the reciprocal nature of decision-making among African kin networks.

By illustrating the role of anticipation in outcomes for foster families, this research demonstrates the importance of accounting for circumstantial factors in research on child fostering. In this research, the empirical results show that when fostering is treated simply as an event that either happened or did not happen, fostering does not seem to impact a household's wealth within the approximately two-and-a-half years of the study period. However, conditioning the fostering experience on prior anticipation reveals that household wealth does change in response to fostering a child under certain circumstances. Households that anticipated fostering responsibilities experienced significant gains in wealth. This finding is robust to a host of controls and lends support for Hypothesis 1, which predicted that household wealth increases in response to fostering a child.

The findings presented in this paper yield new results that are empirically robust and theoretically coherent. Nevertheless, this study has limitations. First, the TLT sample is young, and respondents are not necessarily the heads of their households. In using a relatively young sample, the analysis does not capture "skipped generation" households containing only foster children and grandparents. Thus, caution should be taken in gen-

eralizing the findings for these types of households. It is possible that such households may not experience the benefits associated with anticipated fostering due to unfavorable dependency ratios.

A related concern is that this study uses the anticipation reported by one household member as a proxy for the level of anticipation of the household. However, given the collective nature of fostering decisions, it is likely that all household members have some awareness of future fostering—though not all equally. Because of the significant predictive ability of the anticipation measure, and because anticipation does not operate differently for different types of people, the anticipation of the respondent as a representative household member is validated.

Second, there are limits to relying on a household wealth index as the dependent variable. Although the use of wealth indices has numerous benefits in contexts for which data on income and consumption are limited (Filmer and Pritchett 2001; Rutstein and Johnson 2004), such a measure may not accurately capture the entire picture of wealth in all developing contexts (Bingenheimer 2007). Like many standard wealth indices, the wealth index used in this study does not include measures of livestock and land ownership. While these are more traditional forms of wealth in many sub-Saharan African countries, TLT only measured these assets at the first and last wave of the study—an interval that is too wide for the present analysis. Therefore, the study's findings should be interpreted in relation to more modern, cash-oriented forms of wealth.

Finally, this study does not consider every piece of the fostering puzzle, and there is more for future research in this area to explain. Specifically, I do not explore whether characteristics of the foster child differentially influence household outcomes. It will be a productive avenue for future research to account for the parental survival status of the foster child in order to draw distinctions between households which foster orphans and households which foster non-orphans. It was also beyond the scope of this research to look at the movement of foster children in and out of households. While this is important for future research to consider, it does not compromise the ability to examine household wealth before and after a fostering event. Given that the time period covered in the study is relatively short and that the economic effects of fostering a child likely play out over an extended period of time, this is unlikely to substantially impact the results presented here.

While this study is unable to explain the process by which anticipation acts to increase wealth among fostering households empirically, evidence from broader literature on child fostering allows us to speculate about mechanisms that this study does not consider. These may include child labor or domestic work (Ainsworth 1996; Andvig, Canagarajah, and Kielland 1999; Caldwell 1997; Goody 1982) and incentives received from natal families (Bledsoe and Isiugo-Abanihe 1989; Dahl 2009; Isiugo-Abanihe 1985). It is possible that anticipation is related to the degree to which households are able to choose the child they foster. Households that anticipate fostering may purposely select a child who would

contribute to the household or may negotiate a favorable arrangement with the sending family.

An alternative explanation may be that anticipating fostering represents a number of other qualities that make fostering households more likely to experience an increase in wealth. For example, anticipation could represent better communication within the household and between extended kin networks outside the household. If such superior communication extends into the community at large, it may be indicative of the status of a household within a community or, at a minimum, serve as a proxy of that household's ample access to information and resources.

Prior knowledge of an event reduces uncertainty and increases the ability to plan accordingly. Thus, anticipation of fostering could allow households to mobilize the necessary resources in advance of receiving a foster child, making the household better able to productively leverage additional household members. It could be argued that patterns of superior communication and planning at the household level protect households against instability more generally. Thus, to the extent that anticipation captures these two attributes, it may represent greater stability within the household.

Fostering households that do not fully anticipate fostering responsibilities may not experience the same gains in wealth that are associated with anticipated fostering. However, the fact that these fostering households are not significantly worse-off than their non-fostering counterparts is encouraging, especially considering that these households—particularly the "surprised" households—may have experienced fostering as a true "shock." In other words, despite recent concerns about the toll that orphanhood and child care responsibilities are taking on African families (Barnett and Blaikie 1992; Gregson, Mushati, and Nyamukapa 2007), fostering under perhaps less-than-optimal circumstances may not be hurting households economically. Still, more research is needed to tease out how the variety of circumstances under which fostering takes place differentially effects fostering households.

I have attempted to offer explanations for why anticipatory fostering, in particular, contributes positively to changes in household wealth. However, these explanations are speculative, and more work is needed to understand the meaning of anticipation as it relates to the process of child fostering. While TLT provides a rich dataset for initiating such an inquiry, without data on household decision-making, communication, and indicators of community status, the true role of anticipation must be left for future studies on this topic. Likewise, anticipation is but one aspect that characterizes a fostering situation. To ascertain a fuller picture of the experience for fostering households, future research should consider how the effects of other surrounding circumstances, such as the characteristics of the foster child (e.g., age, relationship to the household head, gender) shape household outcomes. Moreover, wealth is only one measure of household wellbeing. Studies focus-

ing on fostering households should also explore changes in household food security and members' health to understand other ways receiving a foster child impacts households.

While the orphan crisis is certainly impacting sub-Saharan Africa, additional caretaking responsibilities may not be as detrimental to the African household as conventional wisdom suggests or as other scholars have speculated. In the current study context, child fostering does not seem to damage the wealth of host households. In fact, under some circumstances, households that absorb a foster child may experience net economic benefits over time. This supports the idea that child fostering continues to be a highly functional system in certain sub-Saharan African contexts. The findings also indicate that the system of child fostering deserves to be examined in more complexity.

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Appendix A: Supplementary tables

Table A1: Comparison of full and analytic sub-sample for key variables

	Full Sample		Analytic Sample		
Variable	Mean	S.D.	Mean	S.D.	
Household wealth	-0.13	1.8	-0.09	1.7	
% Fostered a child	19.3	0.4	19.4	0.4	
Household Size	5.1	2.2	4.8	2.2	
Years of Education	7.4	2.7	7.3	2.7	
% Married	41.4	0.5	44.1	0.5	
% Student	36.9	0.5	34.1	0.5	
Monthly Income (thousands of Kwacha)	2.3	7.2	2.3	7.0	
Moved house	12.8	0.3	12.6	0.3	
Person-waves	14,555		10,987		
Individuals	208	30	17	754	

Table A2: Cross-sectional description of sample at the beginning of the observed interval (Wave 2)

Variable	Mean	S.D.	Range
Household wealth	-0.11	1.73	-1.85 - 8.15
Household Size	5.10	2.18	1 - 14
Years of Education	7.11	2.72	0 - 13
% Married	39.00	0.49	0 - 1
% Student	43.1	0.50	0 - 1
Monthly Income (thousands of Kwacha)	2.35	8.37	0 - 250
(N=1684 individuals)			

Table A3: Wave 1 predictors of fostering in Wave 2

	Odds Ratio
Anticipation level	1.053* (0.026)
Household wealth	1.180*** (0.052)
Married	1.820** (0.383)
Gender (ref.=male)	1.195 (0.239)
Years of Education	1.061 (0.036)
Age	0.970 (0.031)
Monthly Income (thousands of Kwacha)	1.00 (0.014)
Constant	0.084*** (0.050)
Individuals	1668

Notes: Standard errors in parentheses; all predictor variables measured at wave 1. *** p<0.001, ** p<0.01, * p<0.05

Table A4: Fixed effects regression model of the independent effect of anticipation on household wealth

	Coefficient
Fostering (ref. = non-fosterers)	0.022
Anticipation of Fostering	(0.032) -0.001 (0.003)
Time	0.026*** (0.004)
Constant	-0.219*** (0.021)
Person Waves	10,987
Individuals	1754
Within individual R^2	0.006

Notes: Standard errors in parentheses. *** p<0.001, ** p<0.01, * p<0.05

Appendix B: Expanding the definition of anticipated and surprised fostering

Supplementary analyses explored the effect of anticipation using different cut-points to define surprised and anticipated fostering. Expanding the definition of surprised fostering to include those who selected either a 0- or 1-bean chance of fostering and anticipated fostering to include respondents who selected either a 9- or 10-bean chance of fostering produced similar results. Table B1 shows the results of the multivariate analyses using these new categorical definitions. Anticipated fosterers under the expanded definition continue to experience a significantly larger increase in wealth than their non-fostering counterparts. But unlike Model 4 in Table 4, uncertain fosterers do not experience a statistically distinguishable change in wealth under this new definition, suggesting that the effect of uncertain fostering is less robust than that of anticipated fostering. Furthermore, households that fostered out of anticipation under this expanded definition continue to show a significantly greater increase in wealth than their surprised counterparts (p<0.01, model not shown).

Table B1: Fixed effects regression models of the effect on household wealth by anticipatory fostering status (expanded anticipation cut-points)

	Model 1	Model 2
Time	0.026*** (0.004)	0.024*** (0.004)
Prior Anticipation (ref. = non-fosterers)		
Surprised Fostering	-0.049 (0.057)	0.007 (0.057)
Anticipated Fostering	0.171** (0.066)	$0.202^{**} (0.065)$
Uncertain Fostering	-0.004 (0.040)	0.044 (0.038)
Sociodemographic controls		
Household Size		0.064^{***} (0.007)
Years of Education		0.047^* (0.020)
Student		0.099** (0.033)
Married		0.141*** (0.036)
Monthly Income (thousands of Kwacha)		0.006*** (0.002)
Household Move		0.023 (0.025)
Constant	-0.222^{***} (0.019)	-0.981^{***} (0.150)
Person-waves	10,987	10,987
Individuals	1754	1754
Within individual R^2	0.007	0.020

Notes: Standard errors in parentheses. *** p<0.001, ** p<0.01, * p<0.05

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