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Research Article

Does moving for family nest-building inhibit mothers' labour force (re-)entry?

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Does moving for family nest-building inhibit mothers' labour force (re-)entry?

Stefanie Kley¹

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Abstract

BACKGROUND

Couples tend to move house around first childbirth and often into suburban or rural neighbourhoods, conforming to the normative belief that children should grow up in a 'proper family home.' Such moves are likely to increase housing costs and both partners might need to contribute to the household income. But the move might also necessitate long commutes, inhibiting mothers' labour force participation. If the family sphere is more salient for (prospective) mothers, they might accept a remote location for its family-friendly environment but also because they are not planning a rapid return to work.

OBJECTIVE

This article analyses the influence of moving around first childbirth on the timing of mothers' transitions into employment after childbirth.

METHODS

Event history methods are used on longitudinal data from the German Socio-Economic Panel 1999–2014 (N = 1334 first-time mothers).

RESULTS

Limited evidence was found for the hypothesis that moving around first childbirth accelerates mothers' labour market (re-)entry: moving for homeownership increased the entry rate into full-time employment for mothers with low earnings potential. Strong evidence was found for the hypothesis that moving around first childbirth impedes mothers' employment, particularly hampering entering part-time jobs, the domain of working mothers in Germany and other countries.

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CONCLUSION

Moving for family nest-building seems to place mothers in unfavourable structural positions for employment.

CONTRIBUTION

This article shows that social inequalities among women and within households as well as the persistence of gendered life courses can be reinforced through processes of family migration.

1. Introduction

A number of influential factors for mothers' labour force re-entry at various levels of inquiry have been identified, but it is rarely questioned in what ways the birth of a child impacts the multiple life domains of family members. There is a vast body of research demonstrating that motherhood is associated with women reducing their working hours, interrupting their employment, and dropping out of the labour force, although individual characteristics such as age at childbirth, marital status, educational level, previous labour market participation, occupational position, and earnings potential play an important role in (re-)employment decisions (e.g., Desai and Waite 1991; Drobnič 2000; Drasch 2013). In general, the findings are consistent with the human capital approach (Becker 1985), with women who invested more in education and training being more likely to return to the labour market and to return earlier after the birth of a child. On the macro level, comparative research has documented large cross-national differences in mothers' labour force participation, thus pointing to the importance of normative beliefs and structural opportunities and constraints (Sayer and Gornick 2012; Drasch 2013; Janus 2013; Ziefle and Gangl 2014). Besides the individual and societal level, it has been recognized that decisions about balancing family and work roles and the timing of re-employment after childbirth are decisively driven by household constellations and partners' characteristics (Blossfeld and Drobnič 2001; Kanji 2011; Begall and Grunow 2015). Nevertheless, research on female employment around childbirth has not recognized that family migration may contribute to mothers' employment decisions.

The aim of this article is to assess whether moving around the time of first childbirth delays mothers' labour force participation and if so, whether factors can be identified that might explain this relationship. Specifically, the present study seeks to add to the literature on both female employment and family migration by assessing the dynamic and interrelated processes of family formation, employment transitions, and residential mobility from the life course perspective.

Having children goes hand in hand with increased housing needs, not only in terms of the size of the dwelling. Like other social needs, housing is a normative concept. There is a widespread belief that children should grow up in a ‘proper family home,’ preferably a detached or semi-detached house in a suburban or rural area with its own garden (Lauster 2010). The need for additional space or the desire for a more pleasant environment in which to raise children are major determinants for families deciding to change their residence. Therefore, moves around first childbirth are common, and they normally lead families to the outskirts of cities (Rossi 1980 [1955]; Stockdale and Catney 2014), which often increases the distance between home and the workplace considerably or constrains job searching and accessibility in the case of a job change (MacDonald 1999). On the one hand, improved housing is likely associated with increased costs and is often affordable only if both partners contribute to the household income. Therefore, it can be expected that increased housing costs after moving around first childbirth speed up mothers’ (re-)entry into employment. On the other hand, long commutes make employment and particularly part-time work – the domain of mothers with young children – unattractive. Increased commuting costs and accessibility constraints may therefore counteract the push factors of increased housing costs. This argument raises the question whether (prospective) mothers and their partners might not anticipate the likely consequences of moving on their employment prospects beforehand. The outcome might be the same – constrained employment participation of mothers – but the mechanisms would be different. In the latter case, the self-selection of families would be involved, with moving expected to show a stronger effect on women with lower labour market commitment and low earnings potential.

Germany provides a particularly interesting country case. Family policy in former Federal Republic of Germany followed a long tradition of conservative policies – a joint taxation system, comparably generous child benefits, limited public support for infant childcare, and limited full-day schooling. The division of gender roles has been pronounced and children had a strong impact on women’s employment interruptions (Ziefle and Gangl 2014), with part-time work being an important form of employment for mothers (Drobnič, Blossfeld, and Rohwer 1999; Drobnič 2000). In contrast, the former German Democratic Republic was a dual-earner oriented state before reunification in 1990, but in the decades after reunification women’s labour force participation as well as gender attitudes and norms have been converging towards the West German model (Rosenfeld, Trappe, and Gornick 2004). During the 2000s, the country became markedly more dual-earner oriented, with the introduction of Nordic-style parental leave in 2007 combined with a significant expansion of public childcare (Drasch 2013; Gangl and Ziefle 2015). Still, the West German tradition of pronounced gender roles makes the country especially suited to test the assumption regarding self-selection of mothers into homemaking via moving for family nest-building, even after

the introduction of new family policies. In this study, we analyse data from the German Socio-Economic Panel Study (SOEP) in the period 1999–2014.

2. Study background

2.1 Family migration

It has long been recognized that family ties and household context have an impact on the probability of migration and on consequent changes in employment and earnings of family members (Cooke 2008). The general consensus is that family migration is most likely to occur in support of the man's career and that the woman's employment and earnings are likely to suffer as a result of such moves (Mincer 1978; Shihadeh 1991; Bielby and Bielby 1992; Boyle et al. 2001; Boyle, Feng, and Gayle 2009; Cooke et al. 2009; Abraham, Auspurg, and Hinz 2010; Shauman 2010; Lersch 2016; Vidal et al. 2017), although there is evidence that couples' migration for the sake of the female career has become more widespread (Nisic 2009, 2017). What these studies have in common is the focus on labour market position of individuals and couples. Essentially, family migration is seen as driven by career opportunities of individuals, and the research focuses on long-distance moves. Specific reasons for the move have not been considered. Exceptions are Boyle, Feng, and Gayle (2009); Abraham, Auspurg, and Hinz (2010); Lersch (2016); and Vidal, Perales, and Baxter (2016), but also these studies are predominantly concerned with employment-related reasons for family migration, distinguishing between migrants who moved for the man's job, those who moved for either the woman's or both partners' jobs, and those who moved for other reasons. The study by Lersch and Uunk (2017) identified aspirations for moving into homeownership.

Beyond family migration research, the life course perspective has addressed the interlinkage between various life domains and processes that surround family nest-building (Huinink and Kohli 2014). This robust finding proposes that changes to family composition, such as partnership formation, marriage, and childbirth, are associated with residential mobility (Rossi 1980 [1955]; Mulder and Wagner 1993; Kulu 2008; Kley 2011; Lersch 2014a; Vidal, Huinink, and Feldhaus 2017) as well as first-time homeownership (Clark, Deurloo, and Dieleman 1994; Mulder and Wagner 2001; Feijten and Mulder 2002). This literature suggests that moves are undertaken to meet (anticipated) needs for more housing space and to fulfil a desire for a child-friendly environment. Such moves typically lead families to the city outskirts as well as rural areas because spacious dwellings in inner cities can be very costly and thus not affordable for most families (Rossi 1980 [1955]; Stockdale and Catney 2014). A

‘proper family home’ is for many a self-owned house with a private garden in a residential area (Feijten and Mulder 2002), reflecting a widespread belief that children should grow up in quiet surroundings and not within the hustle and bustle of city centres (Matthews et al. 2000; van Dam, Heins, and Elbersen 2002). Housing decisions and childbearing choices are thus interrelated and best modelled jointly when the interplay between both processes is analysed (Kulu and Steele 2013).

In our study the research question is different, focusing on (re-)employment of first-time mothers. Childbirth is the starting point for our analysis, and housing transitions close to this event are the main predictor. This outline of analysis fits well into a new strand of research analysing the consequences of couples’ residential relocations with regard to shifts in women’s employment situations (Nisic 2009, 2017; Lersch 2014b; Vidal, Perales, and Baxter 2016). In the following sections, the relationship between first childbirth and housing relocation will be first analysed descriptively. It will be shown that moves close to first childbirth are widespread. Comparing area and housing characteristics before and after the move indicates that on average moving leads to a more ‘child-friendly’ environment, larger dwellings, higher housing costs for tenants, and more homeownerships. An increase in costs might primarily motivate an increase in labour supply, as was found for women anticipating relocation into homeownership (Lersch and Uunk 2017). These findings let us hypothesize that:

Hypothesis 1: Moving around first childbirth is related to mothers’ increased (re-)entry into employment after childbirth to accommodate higher costs associated with moving and better housing.

2.2 Residential mobility and the work-family nexus

However, moving around first childbirth often leads families to suburban areas (Rossi 1980 [1955]; Kulu 2008; Kulu, Boyle, and Andersson 2009; Lersch 2014a, 2014b; Stockdale and Catney 2014), and in many cases the distance between home and the workplace increases considerably. Long commuting distances as well as more job search and accessibility constraints in residential locations were found to hamper women’s labour market participation more than men’s (MacDonald 1999; Lee and McDonald 2003), which has also been referred to as the spatial entrapment hypothesis. Particularly part-time work, the domain of working mothers (Drobnič 2000; Fagan and Warren 2001; Sandor 2011), might become unattractive with long commuting. This leads us to the proposition that:

Hypothesis 2: Moving around first childbirth hampers mothers' (re-)employment chances, and this effect is particularly pronounced for entering part-time employment.

Predictions based on higher housing costs on the one hand and spatial constraints arguments on the other lead to conflicting outcomes. It thus remains an empirical question whether moving around first childbirth accelerates or delays mothers' labour force (re-)entry in contemporary societies.

Furthermore, we would like to assess the underlying mechanisms behind the one or the other empirical finding, as our knowledge in this area is limited. The constraints mothers encounter with regard to combining work and family life can be considered to fall into one of two broad categories: structural and normative (McRae 2003). Important elements of structural constraints are job availability, childcare availability and costs, and opportunity costs of forgone earnings, as well as place of residence and the opportunities for affordable 'family nest-building,' as outlined above. The normative category includes beliefs about gender roles and the spouse's/partner's role and attitudes towards family life, mothering, and childcare, including beliefs about the 'proper family home.' The different categories of constraints are likely intertwined. In the context of our study, if increased costs and spatial constraints are independent mechanisms that, respectively, accelerate and delay mothers' employment returns, the effect of moving should persist after controlling for human capital, labour market commitment, and normative views on the work-family sphere. But if the influence of moving is markedly reduced, moving is likely a mediator in a causal chain that links labour market commitment and family orientation, respectively, with re-employment behaviour of mothers. This would support the assumption that:

Hypothesis 3: Moving for family nest-building may be a path that facilitates mothers' self-selection into homemaking.

Since the data used lacks a direct measure of gender role attitudes, we use information on reasons for moving to approximate some normative and structural constraints to mothers' (re-)employment in connection with family nest-building. On the basis of the theory of planned behaviour (Fishbein and Ajzen 2010), moving house can be seen as a means to realize goals in different life spheres that become salient at specific stages of the life course, such as starting a family (De Jong and Fawcett 1981; Willekens 1987). Although people give manifold reasons for moving when questioned in a survey (Clark and Maas 2015), improving the sphere of family life was found to be the most important overarching goal behind moving house in young families (Kley 2011). Moreover, fertility intentions were found to be strong predictors for young

couples' residential relocations (Vidal, Huinink, and Feldhaus 2017). Specific motives like buying a house, improving surroundings, or getting married may all contribute to achieving this goal; they reflect perceived needs and necessities that become particularly salient when having a child. To substantiate the general findings on the effects of moving, we will test whether specific reasons for moving around first childbirth trigger or hamper mothers' (re-)employment, net of other influences. The reasons for moving taken into consideration in this study are (a) moving for homeownership, as buying a home is typically associated with high expenses; (b) moving for marriage, as for many couples' marriage is normatively connected with starting a family and therefore a sign of increased salience of the family sphere; (c) moving for occupational reasons, which is expected to directly accelerate labour market (re-)entry; and (d) moving for other reasons, which include, among others, improved housing and surroundings as well as "other family reasons." In the final step, we will examine if the findings hold for (re-)entry into both full-time and part-time work.

3. Method

3.1 Data

The present study draws on 1999–2014 German Socio-Economic Panel data (SOEP) (<http://www.diw.de/en/soep>), a nationwide study of individuals living in private households in Germany (Wagner, Frick, and Schupp 2007). Started in 1984 in former Federal Republic of Germany and extended to the former German Democratic Republic in 1990, the SOEP is a nationally representative panel survey that provides information on social and demographic backgrounds, education, fertility and marital histories, and job histories, as well as data on current employment and income for all household members aged 16 and older. The head of the household provides additional household-related information, e.g., about the size of the household and regional information. This data is well-suited for analysing the outcomes of residential mobility and internal migration as all participants in the study are followed if they relocate within the boundaries of Germany.

To analyse the duration of mothers' employment interruptions following childbirth and their (re-)entry rate into employment, we identified 1,455 women who gave birth for the first time between 2001 and 2013 and for whom data on their employment, partnership, and housing histories was available. We omitted the first decade after German reunification, which might have had an impact on unique patterns of childbearing and geographic mobility, particularly for East German respondents. Moreover, a large refreshment sample was drawn in 2000 that considerably increased

the number of respondents. We constructed a spell data set suitable for event history analysis. Each spell begins with the month of a couple's first childbirth and ends either with the first month of gainful employment or is censored at the time of the last interview. Information on the partner was collected only if he lived in the household. In our sample, 8% of the first-time mothers did not live with a partner during the whole observation window. As housing and employment issues are fundamentally different for single mothers, they were excluded from the analysis. Retained in the sample are the 4% of women who first lived alone but then moved together with a partner. If one of the partners moved out of the household or died, the observation was treated as censored (applies to 8% of the women). The final sample consists of 1,334 mothers who lived with a partner in a joint household at least some months after first childbirth. The analysed time at risk is about 35,000 person-months in total, while the women are observed for 26 months on average (minimum: 1 month; maximum: 138 months). Within the observed time span, 48% of the mothers re-entered the labour market.

3.2 Method of analysis and variables

We employed event history modelling techniques (Blossfeld, Golsch, and Rohwer 2007) that are suitable for the analysis of longitudinal data and the timing of events. To estimate the transition rate, the Cox proportional hazards model (Cox 1972) was applied to estimate the time it takes until women re-enter the labour market after first childbirth, dependent on time-constant and time-varying covariates. The complex survey design of the German Socio-Economic Panel is a challenge with regard to the question whether the data should be weighted or not. The application of design weights is generally not necessary – and it is not recommended because of a loss in efficiency in the parameter estimates – if the items of stratification are included in the correctly specified models (Winship and Radbill 1994). But we know that some groups of respondents exhibit distinct behaviour with regard to combining work and family. Mothers' labour force participation was very common in the former socialist East German state. The Eastern part of Germany is still better equipped with childcare facilities, gender attitudes are more egalitarian, and mothers' employment is more widespread compared to the Western part of Germany (Matysiak and Steinmetz 2008; Bauernschuster and Rainer 2012; Trappe, Pollmann-Schult, and Schmitt 2015). Therefore, we applied Cox regression accounting for three strata: West Germans, East Germans, and immigrants. As these groups also represent the most important sampling strata, the models were not weighted. Link tests were performed to ensure technically correct model specification (Cleves et al. 2010: 203 ff.). The Efron method of handling ties in time data is used

because it is considered more accurate than Breslow's approximation (Cleves et al. 2010: 145 f.).

We performed the analysis in two steps, first estimating the hazard rate of entering employment in general, followed by competing risk models to assess whether there are differences in determinants of entering full-time and part-time jobs. These models estimate the risk of entering full-time employment after first childbirth whereas the histories of those who started part-time work are considered in the estimation until the occurrence of the competing event (a part-time job). At this point the episode gets censored. The risk of entering part-time employment is estimated analogously. For yielding correct results, this modelling approach must meet two additional assumptions (Cleves et al. 2010: 365). First, there must not be entirely different causes behind the two competing risks. Second, censoring must not affect potential failure times. This might be debatable, since women working full-time in high-level positions might find returning to work on a part-time basis practically impossible; as a reaction, these women could extend their leave period. We tested the Kaplan–Meier estimates for staying out of the labour force for different income groups and found no evidence for such differences.

The dependent variable is the transition rate into gainful employment after first childbirth, measured on a monthly basis. Employment refers to actually working; childcare leave is treated as currently not employed, whereas having been on maternity leave at first childbirth is controlled for as time-constant information. Childcare leave in Germany consists of two distinct programs: maternity leave that stipulates a 14-week period around expected delivery, with compulsory eight weeks after the childbirth, during which mothers continue to receive their full salary; and parental leave, which is optional and can be taken by either parent. The length of paid and unpaid part of parental leave has varied during the study period. During this time parents cannot be dismissed or their job contracts renegotiated, but they do not participate in the labour market, or they participate to a low extent, as long as they receive parental leave benefits.

Thus, aside from the eight weeks of maternity leave, the duration of parental leave is at the discretion of the parents. In the sample, the mean time of childcare leave after childbirth is 12 months but there is considerable variation. Out of the mothers surveyed, 25% were on childcare leave for up to 4 months and 5% for 32 months and longer. Additional analyses reveal that the share of women who return to the labour market from childcare leave is 50%, whereas the share of women who enter the labour market without having had maternity leave around first childbirth is only 6%. The information whether employment is full-time or part-time comes from the respondent. Unfortunately, regular part-time work and marginal ('geringfügig') employment cannot be distinguished prior to 2005 when the distinction was introduced in the SOEP data.

We excluded marginal employment from the analysis because these so-called mini-jobs are not integrated into the social security system and are seldom a stable, long-term form of employment. Average volume of work is significantly lower compared to regular part-time employment.

The most important predictors with regard to the research question are those related to household moves, which are coded on a monthly basis. Out of those surveyed, 77% of women moved during the two years before and the two years after first childbirth. This large time span was chosen because couples might move to a child-friendly environment well in advance of conception (Vidal, Huinink, and Feldhaus 2017) as well as after having the first child (Kulu 2008). Moves that occurred during the 24 months before first childbirth were modelled as time-constant information and moves that occurred during the 24 months after first childbirth as time-variant. After preliminary analyses both types of moves were combined because their effects on mothers' labour market (re-)entry were similar (moved before first childbirth: hazard ratio (HR) = 0.76, $p = 0.004$; moved after: HR = 0.80, $p = 0.024$). Moves later than 24 months after childbirth were not considered (HR = 0.88, $p = 0.80$). There are 1,034 movers in the sample, of whom 28% moved before and after first childbirth. In the case of multiple moves the information on reasons for moving, housing, and area characteristics was updated with each move. Reasons for moving were asked with 15 categories, and multiple reasons were permitted. In Table 1 a sample description at $t = 1$, that is, the month of first childbirth, is depicted. A description and additional analyses of reasons for moving can be found in the Supplement.

Other covariates include the mother's education and employment characteristics: labour force experience in years, percentage of full-time experience in employment career, her previous earnings, and public vs. private sector employment. These variables were measured 12 months before first childbirth. As the age at first childbirth, ranging from 16 to 43 years, was highly correlated with labour market experience, it was not included in the models. The partner's full-time employment status and earnings are time-varying covariates. If the first child was born in 2007 or later, an indicator variable captures changes in parental leave regulations aimed at promoting mothers' earlier return to the labour market. The number of children was updated if another child was born before mother (re-)entered employment after first childbirth. Childbirth was measured on a monthly basis. If the information on the exact month of childbirth and previous earnings of either the respondent or the partner was missing, it was imputed and a dummy variable that controlled for such missing information was included in model estimates. These indicator variables were not statistically significant, and if they also had no sizeable effects, they were omitted.

Table 1: Sample description

| Variable | % (Mean) | Min. | Max. |
|--|-----------|------|--------|
| Moved around first childbirth ^a : No move | 22.5 | 0 | 1 |
| Moved for marriage | 12.3 | 0 | 1 |
| Moved for homeownership | 19.9 | 0 | 1 |
| Moved for occupational reasons | 12.6 | 0 | 1 |
| Moved for other reasons | 51.0 | 0 | 1 |
| Partnership and family characteristics | | | |
| Number of children born ^b | (1.3) | 1 | 6 |
| Child born \geq 2007 | 53.0 | 0 | 1 |
| Partner is employed full-time | 66.1 | 0 | 1 |
| Partner's monthly earnings (gross, in €) | (2,905.5) | 0 | 35,000 |
| Employment characteristics^c | | | |
| Total labour force experience (in years) | (6.7) | 0 | 24.4 |
| Percentage full-time experience | (74.1) | 0 | 100 |
| Previous monthly earnings (gross, in €) | (1,167.8) | 0 | 5,900 |
| Public sector | 18.4 | 0 | 1 |
| Education (in years) | (12.7) | 0 | 18 |
| Maternity leave | 79.5 | 0 | 1 |

Source: Own elaboration based on SOEP data, 1999–2014.

Note: N = 1,334 women and their partners. If not indicated otherwise, variables are measured at t = 1, that is, the month of first childbirth. ^a Moves are measured during the 24 months before and 24 months after first childbirth. Multiple reasons for moving were permitted. The average number of reasons per mover is 2.3. ^b In the observation window. ^c Employment characteristics are measured during the 12 months before first childbirth.

4. Findings

Moving around first childbirth is common, as 77% of women did relocate during the 24 months before and 24 months after first childbirth (movers), whereas 23% did not (stayers). For stayers, the observation window is often shorter, comprising 19 months on average (max. 80 months), whereas it is 28 months on average for movers (max. 138 months); but the (re-)entry into the labour market is observed equally often, in 48% of stayers and 49% of movers (not displayed; results available on request). Nevertheless, it is important to account for right-censoring of the data in the analysis; otherwise the effects could be biased due to group-specific lengths of the observation window. This type of censoring typically occurs in panel studies at the time of the last panel wave but could also happen due to panel attrition. It can be easily handled with event history methods (Blossfeld, Golsch, and Rohwer 2007: 41), as will be applied here.

4.1 Employment characteristics of movers and stayers

Movers and stayers also differ with regard to employment characteristics, measured one year before first childbirth (see Table 2). At first childbirth, stayers had on average

more total labour force experience and more full-time experience compared to movers, but they had much lower average earnings; instead, their partners earned more. Additionally, the share of stayers working in the public sector was small compared to movers, but stayers were more often on maternity leave when the first child arrived. These figures suggest that stayers were more often secondary earners before the first child arrived compared to movers.

Table 2: Movers' and stayers' employment characteristics at first childbirth

| | Stayer % (Mean) | Mover ^a % (Mean) | All women % (Mean) |
|--|--------------------|--------------------------------|-----------------------|
| Total labour force experience (in years) | (7.2) | (6.5) | (6.7) |
| Percentage full-time experience ^b | (78.8) | (72.2) | (74.1) |
| Previous monthly earnings (gross, in €) ^b | (547.3) | (1,405.1) | (1,167.8) |
| Public sector ^b | 6.5 | 23.0 | 18.4 |
| Education (in years) | (13.1) | (12.5) | (12.7) |
| Maternity leave | 87.3 | 77.3 | 79.5 |
| Partner is employed full-time | 70.7 | 64.3 | 66.1 |
| Partner's monthly earnings (gross, in €) | (3,249.1) | (2,774.0) | (2,905.5) |

Source: Own elaboration based on SOEP data, 1999–2014.

Note: N = 1,334 women and their partners. Variables are measured at t = 1, that is, the month of first childbirth. ^a Moves are measured during the 24 months before first childbirth. ^b Employment characteristics are measured during the 12 months before first childbirth.

4.2 Housing characteristics before and after moving

A descriptive summary of housing and area characteristics of moves around first childbirth is shown in Table 3. Families tended to move to larger homes with more rooms and higher costs. The share of homeowners increased considerably and reached 37%. The share of family homes increased; this category includes detached, semi-detached, and terraced houses for one or two families. Family homes often have their own garden, an important feature for families with young children. After moving, the share of young families living in metropolitan areas has decreased, and the share of those living in suburban or urbanized areas has increased. Since in the case of multiple moves these figures refer to the last move, the comparisons of characteristics before and after (all statistically significant) are a rather conservative measure of housing changes.

Table 3: Housing and area characteristics of young families before and after moving^a

| | Before | After | t-value |
|--|--------|-------|---------|
| Size of home (in square metre; mean) | 92.2 | 104.3 | 12.8*** |
| Number of rooms (mean) | 3.6 | 4.1 | 12.3*** |
| Rent in € (if applicable; mean) | 512.7 | 563.3 | 7.2*** |
| Homeowner (%) | 25.5 | 36.6 | 9.8*** |
| Family home ^b (%) | 36.1 | 45.2 | 7.2*** |
| Garden (%) | 52.2 | 62.0 | 7.9*** |
| Metropolitan area ^c (%) | 46.8 | 43.1 | -4.0*** |
| Suburban/urbanized area ^c (%) | 38.5 | 45.0 | 5.8*** |

Source: Own elaboration based on SOEP data, 1999–2014.

Note: N = 1,034 movers. *** $p < 0.001$, ** $p < 0.01$, * $p < 0.05$. ^a Moves are measured during the 24 months before and 24 months after first childbirth. ^b Detached or semi-detached house or terraced house for one or two families. ^c Areas are derived from the BIK classification, which is based on municipal data of numbers of inhabitants and employees' commuter flows. 'Metropolitan area' = centres of cities with at least 100,000 inhabitants, 'Rural area' = settlements and towns with up to 20,000 inhabitants, 'Suburban/urbanized areas' = city peripheries and small to medium sized municipalities (>20,000 – <100,000 inhabitants).

4.3 Timing of the labour market entry after first childbirth

To assess the impact of housing relocation on mothers' employment while controlling for potential confounding factors, we turned to event history analysis, using stratified Cox regression, as described in the Methods section. Model 1 in Table 4 shows that moving around first childbirth was associated with a delayed (re-)entry into the labour market. When specific reasons for moving were included in Model 2, a more complex picture emerged. Moving for marriage deterred mothers' (re-)entry into employment. The hazard of entering employment after first childbirth was reduced by 38% ($1 - 0.62$) compared to recent mothers who did not move house, whereas opposite effects of the same size were observed for mothers who moved for homeownership and occupational reasons. The remaining category that subsumed other reasons for moving showed a strong and highly significant negative effect on mothers' (re-)employment.

Partner and family characteristics were added in Model 3 but the coefficients for moving remained largely unchanged. In Model 4, employment characteristics and indicators of women's human capital were included. As expected, higher education, full-time work experience, higher earnings in jobs before first childbirth, employment in the public sector, and having been on childcare leave all accelerated the timing of mothers' (re-)entry into the labour market. The partner's occupational resources had an opposite but barely significant effect. If their partner was employed full-time, women entered the labour market later compared to when their partner was employed part-time, enrolled in education, or in some other non-employment status. If first birth occurred after the introduction of Nordic-style parental leave in 2007, the (re-)entry rate into employment accelerated significantly, whereas giving birth to another child strongly

reduced it. The coefficients for moving for homeownership and occupational reasons were markedly diminished and were no longer statistically significant when employment resources were controlled, such as previous labour market experience, earnings, and educational level. The employment-detering effect of moving for marriage and other reasons around first childbirth remained strong and significantly negative.

Table 4: Hazard ratio of mothers' labour market (re-)entry after first childbirth

| | Model 1 | | Model 2 | | Model 3 | | Model 4 | | Model 5 | |
|--|---------|---------|---------|---------|---------|----------|---------|----------|---------|----------|
| | HR | SE B | HR | SE B | HR | SE B | HR | SE B | HR | SE B |
| Moved around first childbirth^a | | | | | | | | | | |
| Moved (ref. = no move) | | | | | | | | | | |
| For all/other reasons | 0.72 ** | 0.10 | 0.68*** | 0.11 | 0.79* | 0.12 | 0.67** | 0.13 | 0.67*** | 0.12 |
| For marriage | | | 0.62*** | 0.14 | 0.63** | 0.14 | 0.69* | 0.14 | | |
| For homeownership | | | 1.38** | 0.10 | 1.41** | 0.11 | 1.00 | 0.11 | | |
| For occupational reasons | | | 1.35** | 0.12 | 1.32* | 0.12 | 1.10 | 0.12 | | |
| Partnership and family characteristics | | | | | | | | | | |
| Number of children | | | | | 0.48*** | 0.12 | 0.47*** | 0.12 | 0.47*** | 0.12 |
| Child born ≥ 2007 | | | | | 1.36** | 0.09 | 1.30** | 0.10 | 1.31** | 0.10 |
| Partner is employed full-time | | | | | 0.75* | 0.13 | 0.77* | 0.13 | 0.76* | 0.13 |
| Partner's earnings (gross, in €100) | | | | | 1.00 | 0.00 | 0.99* | 0.00 | 0.99* | 0.00 |
| Employment characteristics^b | | | | | | | | | | |
| Total labour force experience (years) | | | | | | | 1.01 | 0.01 | 1.01 | 0.01 |
| Percentage full-time experience | | | | | | | 1.73*** | 0.14 | 1.71*** | 0.14 |
| Previous earnings (gross, in €100) | | | | | | | 1.02*** | 0.00 | 1.02*** | 0.00 |
| Public sector | | | | | | | 1.44*** | 0.10 | 1.44*** | 0.10 |
| Education (in years) | | | | | | | 1.10*** | 0.02 | 1.11*** | 0.02 |
| Maternity leave at first childbirth | | | | | | | 2.04*** | 0.19 | 2.06*** | 0.19 |
| Person-months | 35,259 | | 35,259 | | 35,259 | | 35,259 | | 35,259 | |
| No. of persons | 1,334 | | 1,334 | | 1,334 | | 1,334 | | 1,334 | |
| No. of events | 638 | | 638 | | 638 | | 638 | | 638 | |
| LR χ^2 | | 10.6*** | | 37.7*** | | 101.6*** | | 278.0*** | | 270.4*** |
| df | | 1 | | 4 | | 9 | | 15 | | 12 |

Source: Own elaboration based on SOEP data, 1999–2014.

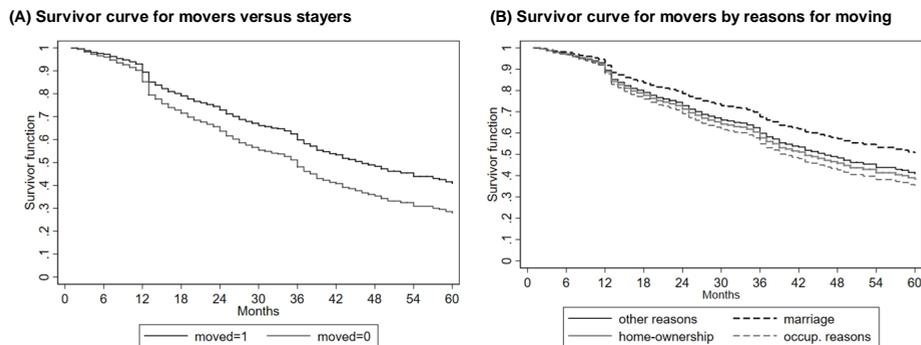
Note: Cox regressions, stratified by East Germans, West Germans, and immigrants. Controls are: partner information missing (omitted from the table). HR = e^B = Hazard Ratio. ***p < 0.001, **p < 0.01, *p < 0.05. ^a Moves are measured during the 24 months before and 24 months after first childbirth. ^b Employment characteristics are measured during the 12 months before first childbirth.

These findings speak against the assumption that buying a home, which typically implies high costs adding up for the new owner, significantly accelerates (re-)entry into the labour force independent of other factors. Rather, the findings suggest that families who buy their home are those who anticipate being able to afford the costs associated with homeownership. Additional analyses reveal that having a high educational level, high previous income, and a job in the public sector are the factors that explain the

association between moving for homeownership and employment (not displayed). These factors are also important when applying for the mortgage. However, the negative impact on employment of moving for marriage, a sign of adherence to the normative image of the family and an increasing salience of the family sphere, supports the assumption of self-selection into homemaking, largely independent of labour force commitment and human capital. Furthermore, the strong negative effect of other reasons for moving, which include factors such as having more space and a child-friendly environment, support the view that spatial constraints may play an important role. As Model 5 shows, the negative effect of moving around first childbirth on mothers' (re-)employment – reasons not specified – persisted, compared to Model 1, when all covariates were included in the estimation.

Figure 1 illustrates that the survivor function of staying out of the labour force for the first five years after childbirth was affected by moving around first childbirth generally (A, left-hand side) and by moving for particular reasons (B, right-hand side). The graphs are based on the estimates of the full Models 4 and 5 in Table 4. In the first 12 months after childbirth the curves of both graphs are rather similar, after 12 months the difference between movers and stayers as well as between those who moved for particular reasons becomes clearly visible. Graph A reveals that mothers who moved around first childbirth (re-)entered the labour market later than their counterparts, who did not move. Graph B shows that mothers who moved for occupational reasons (re-)entered employment the fastest, followed by those who moved for homeownership and other reasons. Women who moved for marriage had the longest employment interruptions.

Figure 1: Estimated survivor curve for duration of absence from the labour market after first childbirth



Source: Own elaboration based on SOEP data, 1999–2014. Estimated on the basis of Models 4 and 5 in Table 4.

4.4 Full-time or part-time employment after childbirth

The results in Table 4 suggest that moving around first childbirth delayed mothers' labour market (re-)entry in general, and there is some evidence for self-selection of women with a stronger family orientation, little work experience, and low earnings potential into homemaking. As these women are particularly responsive to the convenience of work settings (Desai and Waite 1991), we estimated Cox regressions for two competing outcomes: full-time or part-time jobs. Self-selection into homemaking via moving around first childbirth can be expected to delay entering part-time work more strongly than entering full-time work because working part-time might be an expression of a mother's relatively low commitment to the world of work (Hakim 2000), and living in the city outskirts normally comes at the cost of relatively high commuting distance, which makes part-time work unattractive (MacDonald 1999; Lee and McDonald 2003).

Models 6 to 8 in Table 5 present hazard ratio estimates of entering full-time employment after first childbirth whereas the histories of those who started part-time work were considered in the estimation until the occurrence of the competing event (part-time job). At this point the episodes were censored. The hazard of entering the labour market on a part-time basis was estimated analogously and is presented in Models 9 and 10. Again, Cox regression was applied, stratified for East Germans, West Germans, and immigrants to account for stratified samples.

In the case of full-time employment, the hazard rate was strongly affected by human capital measures, partners, and family characteristics (Model 6). Better-educated women who predominantly worked full-time in the past and had high earnings entered full-time employment significantly faster. Having additional children strongly reduced the likelihood of starting a full-time job. Women with a full-time working partner and those with a high-earning partner were less likely to start working full-time compared to other women. Additionally, mothers' returns to the labour market on a full-time basis were triggered by the Nordic-style parental leave introduced in Germany in 2007, when the previous low flat-rate benefit was substituted by a generous earnings replacement but at the same time the paid leave for the primary caregiver was shortened to 12 months. Moving house around first childbirth delayed working full-time considerably, but the coefficient was not statistically significant.

Table 5: Hazard ratio of mothers' (re-)entry into full-time and part-time work, respectively, after first childbirth

| | Full-time | | | | Part-time | | | | | |
|--|-----------|------|----------|------|-----------|------|----------|------|----------|------|
| | Model 6 | | Model 7 | | Model 8 | | Model 9 | | Model 10 | |
| | HR | SE B | HR | SE B | HR | SE B | HR | SE B | HR | SE B |
| Moved around first childbirth^a | | | | | | | | | | |
| Moved (ref. = no move) | | | | | | | | | | |
| For all/other reasons | 0.74 | 0.22 | 0.66† | 0.23 | 0.52* | 0.29 | 0.67** | 0.15 | 0.68* | 0.15 |
| For marriage | | | 0.60† | 0.30 | 1.04 | 0.42 | | | 0.75† | 0.17 |
| For homeownership | | | 1.21 | 0.20 | 2.71** | 0.34 | | | 0.93 | 0.13 |
| For occupational reasons | | | 1.52* | 0.20 | 1.58 | 0.32 | | | 0.94 | 0.16 |
| Partnership and family characteristics | | | | | | | | | | |
| Number of children | 0.37*** | 0.25 | 0.36*** | 0.26 | 0.36*** | 0.26 | 0.51*** | 0.14 | 0.52*** | 0.14 |
| Child born ≥ 2007 | 1.79*** | 0.17 | 1.78*** | 0.17 | 1.79*** | 0.17 | 1.15 | 0.12 | 1.15 | 0.12 |
| Partner is employed full-time | 0.73* | 0.15 | 0.72* | 0.16 | 0.70* | 0.16 | 0.95 | 0.11 | 0.98 | 0.11 |
| Partner's earnings (gross, in €100) | 0.98** | 0.01 | 0.98** | 0.01 | 0.98** | 0.01 | 1.00 | 0.00 | 1.00 | 0.00 |
| Employment characteristics^b | | | | | | | | | | |
| Total labour force experience (years) | 0.96† | 0.02 | 0.96† | 0.02 | 0.96† | 0.02 | 1.02* | 0.01 | 1.02* | 0.01 |
| Percentage full-time experience | 2.32*** | 0.25 | 2.37*** | 0.25 | 2.26** | 0.26 | 1.52* | 0.17 | 1.53* | 0.17 |
| Previous earnings (gross, in €100) | 1.02** | 0.01 | 1.02** | 0.01 | 1.02† | 0.01 | 1.02*** | 0.00 | 1.02*** | 0.00 |
| Public sector | 1.22 | 0.19 | 1.25 | 0.19 | 1.23 | 0.19 | 1.52*** | 0.12 | 1.51*** | 0.12 |
| Education (in years) | 1.13*** | 0.03 | 1.11*** | 0.03 | 1.12*** | 0.03 | 1.10*** | 0.02 | 1.10*** | 0.02 |
| Maternity leave at first childbirth | 1.89† | 0.35 | 1.85† | 0.35 | 1.83† | 0.36 | 2.16*** | 0.23 | 2.16*** | 0.23 |
| Prev. earnings x moved for other reasons | | | | | 1.02 | 0.02 | | | | |
| Prev. earnings x moved for marriage | | | | | 0.96† | 0.02 | | | | |
| Prev. earnings x moved for homeownership | | | | | 0.95** | 0.02 | | | | |
| Prev. earnings x moved for occup. reasons | | | | | 0.99 | 0.02 | | | | |
| Person-months | 35,259 | | 35,259 | | 35,259 | | 35,259 | | 35,259 | |
| No. of persons | 1,334 | | 1,334 | | 1,334 | | 1,334 | | 1,334 | |
| No. of events | 195 | | 195 | | 195 | | 443 | | 443 | |
| LR χ^2 | 106.2*** | | 114.8*** | | 124.0*** | | 180.5*** | | 183.4*** | |
| df | 11 | | 14 | | 18 | | 11 | | 14 | |

Source: Own elaboration based on SOEP data, 1999–2014. Note: Cox regression, stratified by East Germans, West Germans, and immigrants. *** $p < 0.001$, ** $p < 0.01$, * $p < 0.05$, † $p < 0.1$. HR = e^B = Hazard Ratio. ^a Moves are measured during the 24 months before and 24 months after first childbirth. ^b Employment characteristics are measured during the 12 months before first childbirth.

In further model specifications it became evident that mothers who moved for occupational reasons or homeownership were more likely to (re-)enter the labour market on a full-time basis than stayers (Model 7). The influence of moving for homeownership interacted with women's labour market potential (Model 8). The accelerating effect of moving for homeownership was especially pronounced for mothers with low previous earnings, as the effect size decreased with rising earnings potential. We can conclude that for a share of mothers who purchased a new home

around the birth of the first child, the likely high costs associated with moving and housing had an accelerating effect on entering full-time employment. Moving for other reasons, such as a small dwelling, bad facilities, location, and surroundings, deterred mothers from entering full-time employment.

The determinants of the hazard rate were rather different for part-time employment (Models 9 and 10). There was a significant negative impact of moving around first childbirth on women's part-time employment, as expected. The relative hazard for movers was 67% of the hazard of those who did not move around childbirth. Mothers' human capital indicators were again influential, but partners' employment characteristics were of minor importance when it came to part-time employment. Having additional children again hindered employment, while extensive rights to part-time work for public-sector employees were reflected in their significantly increased hazard of entering part-time employment. Having the first child under the new parental leave regulations after 2007 was not significant.

When reasons for moving were differentiated, it became obvious that the negative impact of moving around first childbirth on mothers' part-time (re-)employment was related to moving for marriage or other reasons, whereas moving for homeownership or occupational reasons had no effect. There were no significant interactions between moving for particular reasons and a mother's earnings potential on part-time (re-)entry (see Table A-3 in the Supplement).

5. Discussion

We have explored the links between childbearing, housing relocation, and mothers' employment, using longitudinal data on family formation, residential mobility, and employment patterns of first-time mothers and their partners in Germany. Moves around first childbirth are very common and are usually made to accommodate (anticipated) changing housing needs at the family formation stage. If families move at this time, they tend to move to significantly larger homes with a garden, which are more often located in suburban or urbanized areas compared to their previous dwellings, and which are more often self-owned. Thus, housing relocation comes at a price, particularly in terms of higher housing costs and normally an increased commuting distance for working family members. On the one hand, the financial contribution of women to the household income and therefore their employment becomes more crucial; on the other hand, constraints on paid work are likely to increase in the case of a move into suburban areas. However, normative forces are also activated in the course of moving around childbirth. The perception of housing needs to accommodate

childrearing and the way such needs are met reflect normative ideals on family nest-building involving a 'proper family home' and a child-friendly environment.

Empirical results in the present longitudinal study contribute to the body of literature about interrelated processes in different life domains over the family life course. Decisions on family formation, housing relocation, and women's employment conditions are closely intertwined; these processes constrain or reinforce each other, also depending on the household's and women's resources. First, housing relocation around first childbirth was found to be related to an increased employment rate after childbirth for those women who undertook their moves for homeownership, but this effect was largely explained by mothers' human capital and labour market commitment. Therefore, only limited evidence was found for the hypothesis that moving around first childbirth accelerates mothers' labour market (re-)entry due to high costs for housing and moving. In general, it seems that couples who buy a family home around first childbirth do so if they anticipate that they can afford it without significant consequences for their labour supply. An exception were mothers with low earnings potential. Moving for homeownership significantly increased their (re-)entry rate into full-time employment. This result is consistent with the findings that some women expand their labour force participation to meet increased financial demands after childbirth (Estes and Glass 1996), and that women with low work commitment are especially responsive to financial pressures (Desai and Waite 1991). These women might perceive a necessity to participate in the labour force although this might not necessarily reflect the preferences of the women and/or their partners (Gash 2008).

Second, the analysis lends stronger support for the hypothesis that moving around first childbirth impedes mothers' employment. An overall deterring effect of residential mobility on mothers' (re-)entry in paid work was found, over and above that of human capital, previous employment, and family characteristics. The finding that moving around first childbirth particularly hampers entering part-time jobs is consistent with the view that working part-time becomes unattractive with long commutes. Although we were not able to test whether moving in fact increased the distance to the (former) workplace, this assumption is plausible in the light of our descriptive analysis of housing and area characteristics before and after moving, and it is plausible in the light of previous findings and the arguments on spatial constraints for mothers (MacDonald 1999; Lee and McDonald 2003). Future research might want to test our argumentation.

Finally, our results indicate that first-time mothers are a heterogeneous group of women. They behave in accordance with the human capital predictions and tend to respond to the opportunity costs of staying out of the labour market. However, net of their occupational resources, women who moved for marriage and other reasons that are often related to an improved family sphere (size of dwelling, quality of facilities, location and surroundings, and other family-related reasons) have a significantly lower

risk of entering employment than stayers. It seems that family orientation and normative views on family nest-building may play a role in their moving decisions, which as a result place them in an unfavourable structural position for (re-)employment. Overall, the picture that emerges is consistent with the view that moving for family nest-building often leads to situations that make early labour market (re-)entry after first childbirth less likely. Staying out of the labour force for longer periods of time diminishes women's chances of (re-)entry due to depreciation of human capital and again reinforces their life orientation towards the family sphere. Thus, social inequalities among women and within households as well as the persistence of gendered life courses can be reinforced through effects of family migration on mothers (re-)employment.

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Supplement

Description and additional analysis of reasons for moving

Reasons for moving were reported by the household head. These may not reflect the reasons of other household members, which may be conflicting. Respondents could choose multiple reasons from a list of 15. The distribution of reasons for moving and their effect on mothers' re-entry to the labour market are described in Table A-1. The most frequent reasons for the (last) move were "dwelling too small" (49%), "family, other" (32%), and "buying a dwelling" (referred to as "homeownership," 25%); "childbirth" was not included in the list of reasons for moving. Most movers reported one or two reasons for moving, 2% reported no reason, and 1% reported seven to eight reasons (Table A-2). The reasons "dwelling too small," "bad facilities," "bad surroundings," and "bad location" were found to be significantly correlated. These reasons for moving reflect changing needs with regard to the dwelling and its location that are interpreted as being associated with childbirth.

For the interpretation of the effects of reasons for moving on mothers' (re-)employment reported in the Tables 4, 5, A-1, and A-2, it is important to keep in mind that they are partial effects. The coefficient of "moving for marriage," for example, represents the effect of moves undertaken for this reason net of moves undertaken for all other reasons compared to having not moved at all. This strategy enables us to extract the effect of "moving for marriage" compared to not moving at all, under control of moving for other reasons, and is therefore suitable when people state multiple reasons for moving.

Table A-1: Distribution of reasons for moving among movers¹ and partial effects of moves for particular reasons before or after first childbirth regressed on mothers' (re-)employment²

| Reasons for moving | % | Moved before or after HR | SE B | Moved before childbirth HR | SE B | Moved after childbirth HR | SE B |
|---------------------------|--------------|--------------------------|------|----------------------------|------|---------------------------|------|
| Reference: no move | | | | | | | |
| Notice of termination | 3.0 | 0.72 | 0.24 | 0.77 | 0.26 | 0.75 | 0.35 |
| Homeownership | 25.3 | 1.35** | 0.14 | 1.38** | 0.15 | 1.31 | 0.25 |
| Inheritance | 1.5 | 1.77 † | 0.52 | 1.82 * | 0.54 | · | · |
| Occupational reasons | 16.1 | 1.33 * | 0.16 | 1.31 * | 0.16 | 1.45 † | 0.28 |
| Marriage | 15.9 | 0.63** | 0.09 | 0.65** | 0.10 | 0.77 | 0.19 |
| Separation | 4.9 | 1.16 | 0.29 | 1.23 | 0.31 | 0.94 | 0.38 |
| Leaving parental home | 15.7 | 0.92 | 0.12 | 0.93 | 0.12 | 1.07 | 0.22 |
| Family, other | 32.3 | 1.01 | 0.10 | 1.02 | 0.10 | 1.16 | 0.20 |
| Dwelling too small | 49.2 | 1.03 | 0.10 | 1.04 | 0.10 | 1.15 | 0.20 |
| Dwelling too large | 1.5 | 0.79 | 0.34 | 0.78 | 0.33 | 0.61 | 0.46 |
| Dwelling cost too high | 9.4 | 0.85 | 0.15 | 0.89 | 0.15 | 0.79 | 0.23 |
| Bad facilities | 10.1 | 0.99 | 0.16 | 0.98 | 0.16 | 0.91 | 0.23 |
| Bad location | 7.8 | 1.03 | 0.18 | 1.04 | 0.20 | 0.81 | 0.23 |
| Bad surrounding | 14.4 | 1.07 | 0.16 | 1.05 | 0.16 | 1.42 † | 0.30 |
| Other | 18.2 | 1.02 | 0.12 | 1.07 | 0.13 | 1.07 | 0.21 |
| Total | 225.3 | | | | | | |

Source: Own elaboration based on SOEP data, 1999–2014.

Note: *** $p < 0.001$, ** $p < 0.01$, * $p < 0.05$, † $p < 0.1$. ¹ Moves are measured during the 24 months before and 24 months after first childbirth. Multiple reasons for moving were permitted. In the case of multiple moves, the reasons for the last move are displayed. N = 1,034 movers. ² Cox regressions, stratified by East Germans, West Germans, and immigrants. N = 35,259 person-months. The number of persons is 1,334. Controls are: reason for moving missing (omitted from the table). HR = e^B = Hazard Ratio. · = not enough cases.

Table A-2: Number of statements about reasons for moving

| Count | Freq. | Percent | Cum. percent |
|--------------|--------------|--------------|--------------|
| 0 | 23 | 2.2 | 2.2 |
| 1 | 325 | 31.4 | 33.7 |
| 2 | 299 | 28.9 | 62.6 |
| 3 | 250 | 24.2 | 86.7 |
| 4 | 74 | 7.2 | 93.9 |
| 5 | 30 | 2.9 | 96.8 |
| 6 | 19 | 1.8 | 98.6 |
| 7 | 13 | 1.3 | 99.9 |
| 8 | 1 | 0.1 | 100.0 |
| Total | 1,034 | 100.0 | |

Source: Own elaboration based on SOEP data, 1999–2014.

Note: Moves are measured during the 24 months before and 24 months after first childbirth. Multiple reasons for moving were allowed. In the case of multiple moves, the reasons for the last move are displayed. N = 1,034 movers.

Table A-3: Hazard ratio of mothers' (re-)entry into part-time job after first childbirth, and interactions with previous income

| | HR | SE B |
|--|-----------|------|
| Moved around first childbirth^a | | |
| Moved (ref. = no move) | | |
| For other reasons | 0.66 * | 0.19 |
| For marriage | 0.78 | 0.26 |
| For homeownership | 1.00 | 0.25 |
| For occupational reasons | 0.69 | 0.26 |
| Partnership and family characteristics | | |
| Number of children | 0.52 *** | 0.14 |
| Child born ≥ 2007 | 1.21 | 0.12 |
| Partner is employed full-time | 0.98 | 0.11 |
| Partner's earnings (gross, in €100) | 1.00 | 0.00 |
| Employment characteristics^b | | |
| Total labour force experience (years) | 1.02 * | 0.01 |
| Percentage full-time experience | 1.50 * | 0.18 |
| Previous earnings (gross, in €100) | 1.01 | 0.01 |
| Public sector | 1.49 ** | 0.12 |
| Education (in years) | 1.10 *** | 0.02 |
| Maternity leave at first childbirth | 2.20 *** | 0.23 |
| Prev. earnings x moved for other reasons | 1.01 | 0.01 |
| Prev. earnings x moved for marriage | 1.00 | 0.01 |
| Prev. earnings x moved for homeownership | 1.00 | 0.01 |
| Prev. earnings x moved for occup. reasons | 1.02 | 0.01 |
| Person-months | 35,259 | |
| No. of persons | 1,334 | |
| No. of events | 443 | |
| LR χ^2 | 187.1 *** | |
| df | 18 | |

Source: Own elaboration based on SOEP data, 1999–2014.

Note: Cox regression, stratified by East Germans, West Germans, and immigrants. HR = e^B = Hazard Ratio, *** $p < 0.001$, ** $p < 0.01$, * $p < 0.05$, † $p < 0.1$. ^a Moves are measured during the 24 months before and 24 months after first childbirth. ^b Employment characteristics are measured during the 12 months before first childbirth.

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